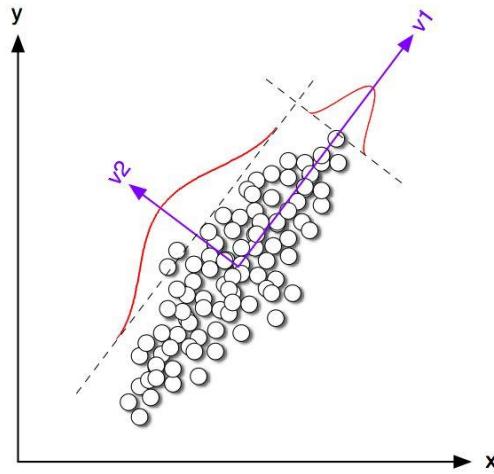


生物统计学： 生物信息中的概率统计模型

2019年秋



有关信息

- 授课教师: 宁康
 - Email: ningkang@hust.edu.cn
 - Office: 华中科技大学东十一楼504室
 - Phone: 87793041, 18627968927
- 课程网页
 - <http://www.microbioinformatics.org/teach/#>
 - QQ群: 882140516



2019生物统计学
扫一扫二维码，加入该群。

课程安排

- 生物背景和课程简介
- 传统生物统计学及其应用
- 生物统计学和生物大数据挖掘
 - Hidden Markov Model (HMM)及其应用
 - Markov Chain
 - HMM理论
 - HMM和基因识别 (Topic I)
 - HMM和序列比对 (Topic II)
 - 进化树的概率模型 (Topic III)
 - Motif finding中的概率模型 (Topic IV)
 - EM algorithm
 - Markov Chain Monte Carlo (MCMC)
 - 基因表达数据分析 (Topic V)
 - 聚类分析-Mixture model
 - Classification-Lasso Based variable selection
 - 基因网络推断 (Topic VI)
 - Bayesian网络
 - Gaussian Graphical Model
 - 基因网络分析 (Topic VII)
 - Network clustering
 - Network Motif
 - Markov random field (MRF)
 - Dimension reduction及其应用 (Topic VIII)
- 面向生物大数据挖掘的深度学习

研究对象：
生物序列，
进化树，
生物网络，
基因表达
...

方法：
生物计算与生物统计

第9章: Dimension Reduction

- Feature selection
- PCA, LDA, CCA, SVD
- MDS, tSNE
- LASSO
- Others

Dimension Reduction

- Inputs: high dimensional

$$X_1, X_2, \dots, X_n \in R^D$$

- Outputs: low dimensional

$$Y_1, Y_2, \dots, Y_n \in R^d, d \ll D$$

- Constraints: Nearby points remain nearby, distant points remain distant.

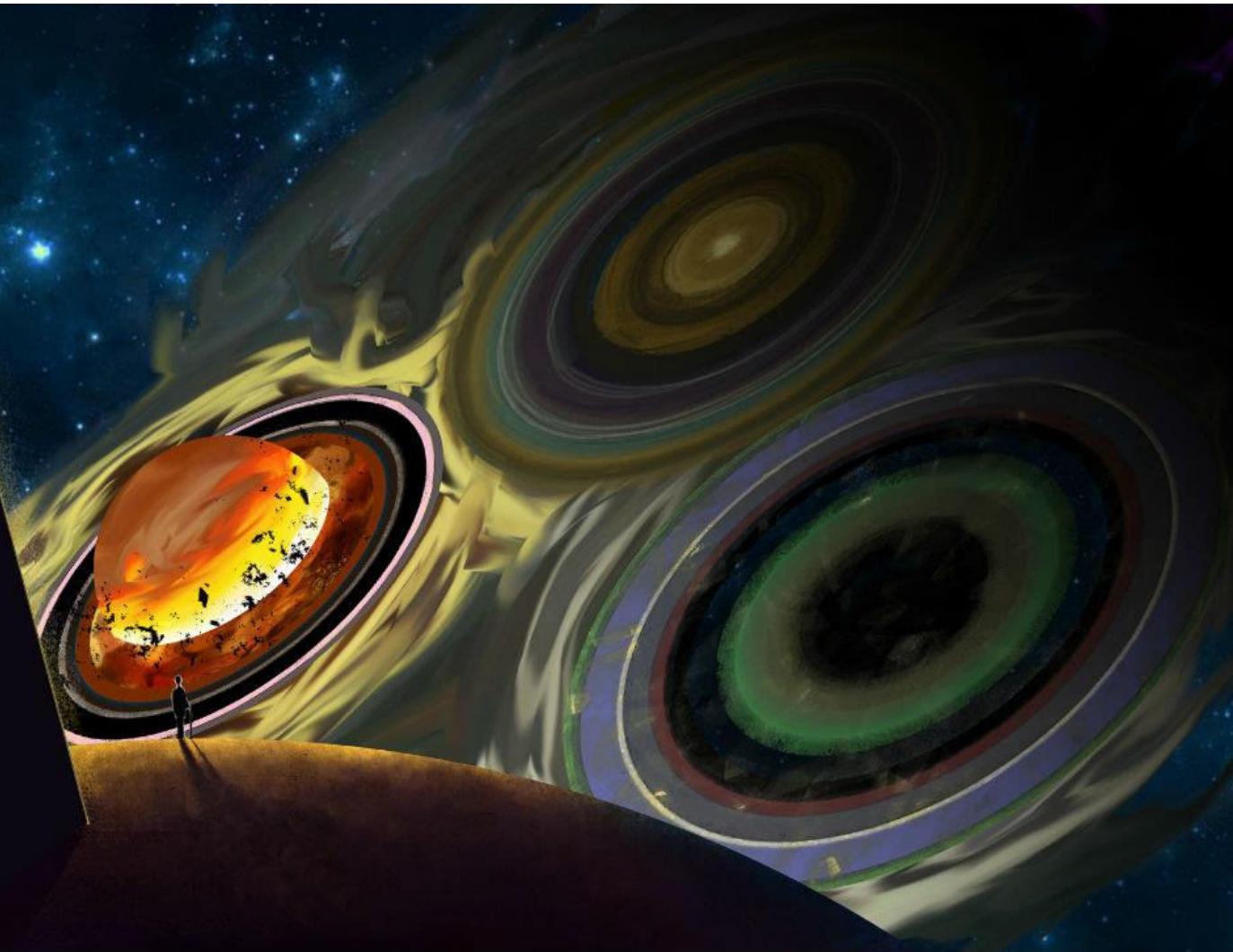
Curse of Dimensionality

- A major problem is *the curse of dimensionality*.
- If the data x lies in high dimensional space, then an enormous amount of data is required to learn distributions or decision rules.
- Example: 50 dimensions. Each dimension has 20 levels. This gives a total of 20^{50} cells. But the no. of data samples will be far less. There will not be enough data samples to learn.

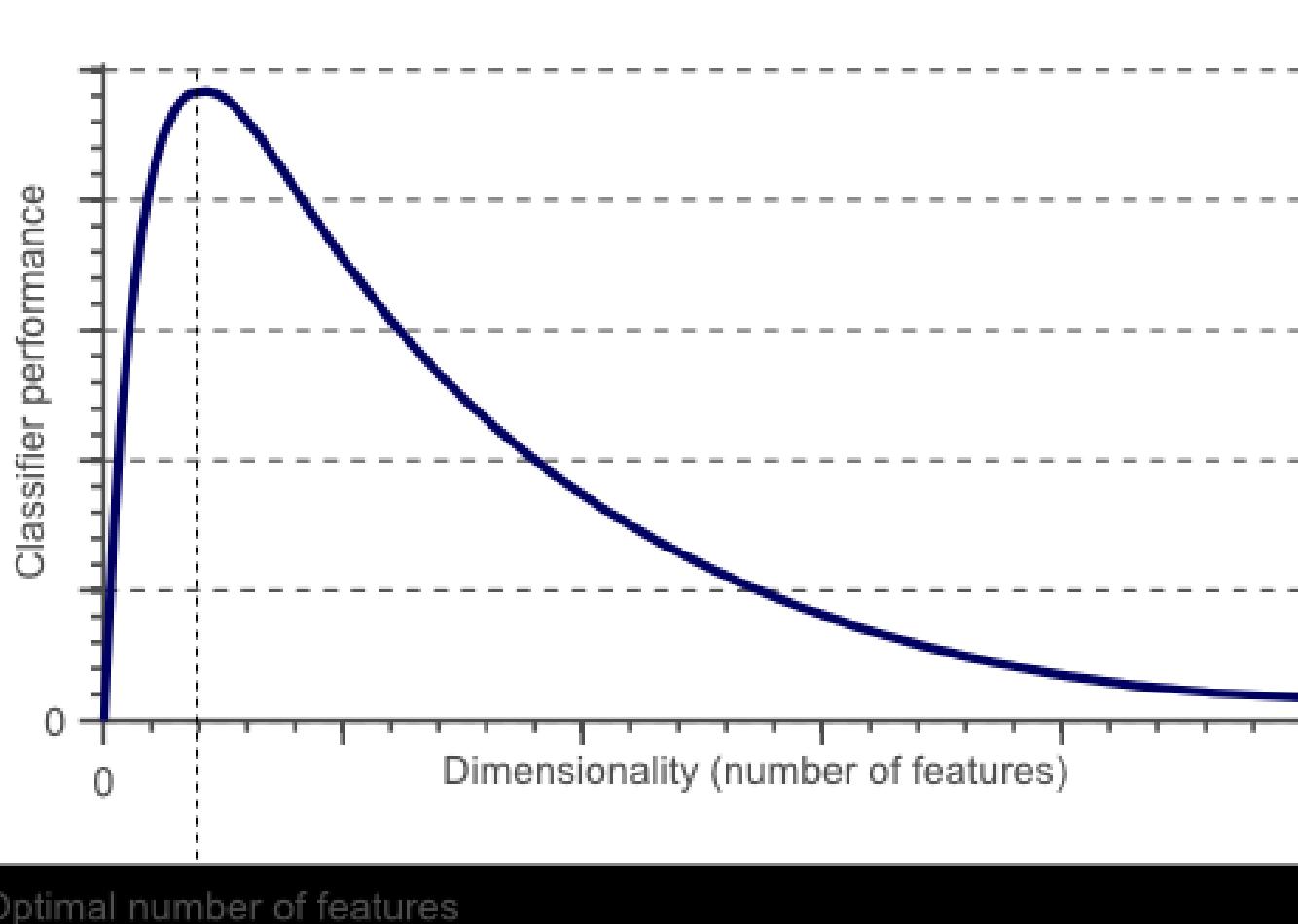
Curse of Dimensionality

- One way to deal with dimensionality is to assume that we know the form of the probability distribution.
- For example, a Gaussian model in N dimensions has $N + N(N-1)/2$ parameters to estimate.
- Requires $O(N^2)$ data to learn reliably. This may be practical.

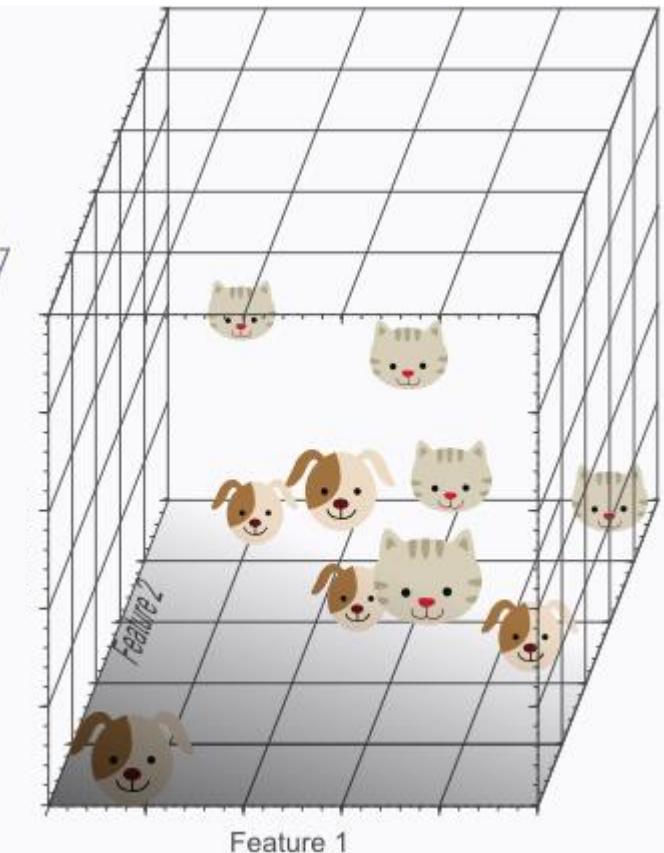
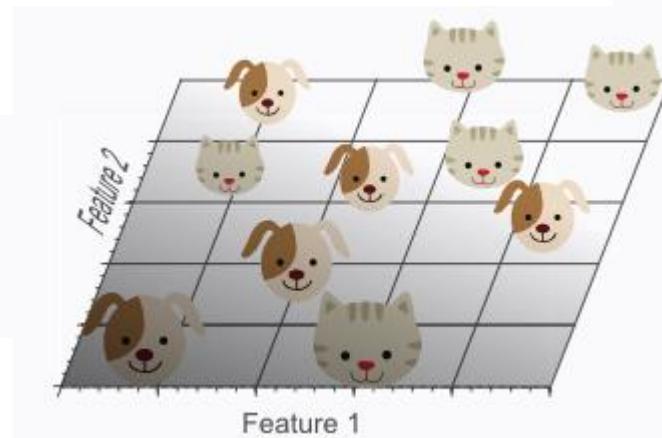
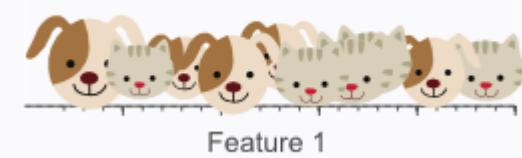
Curse of Dimensionality



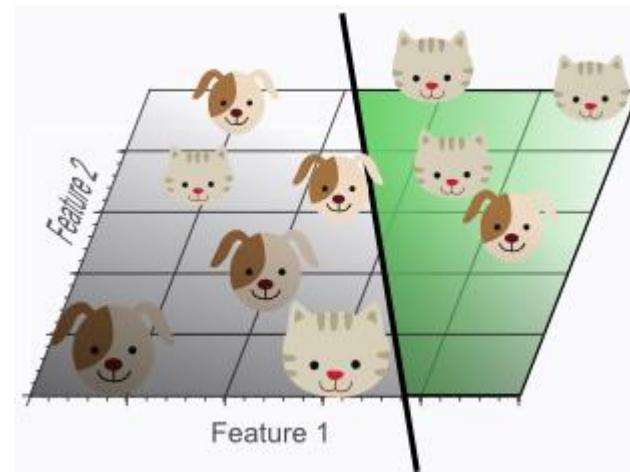
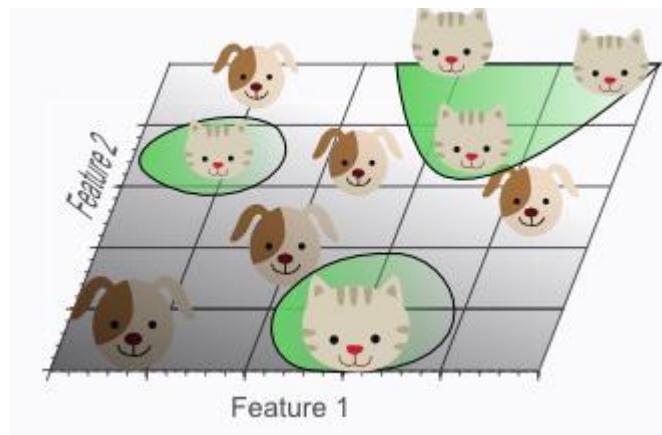
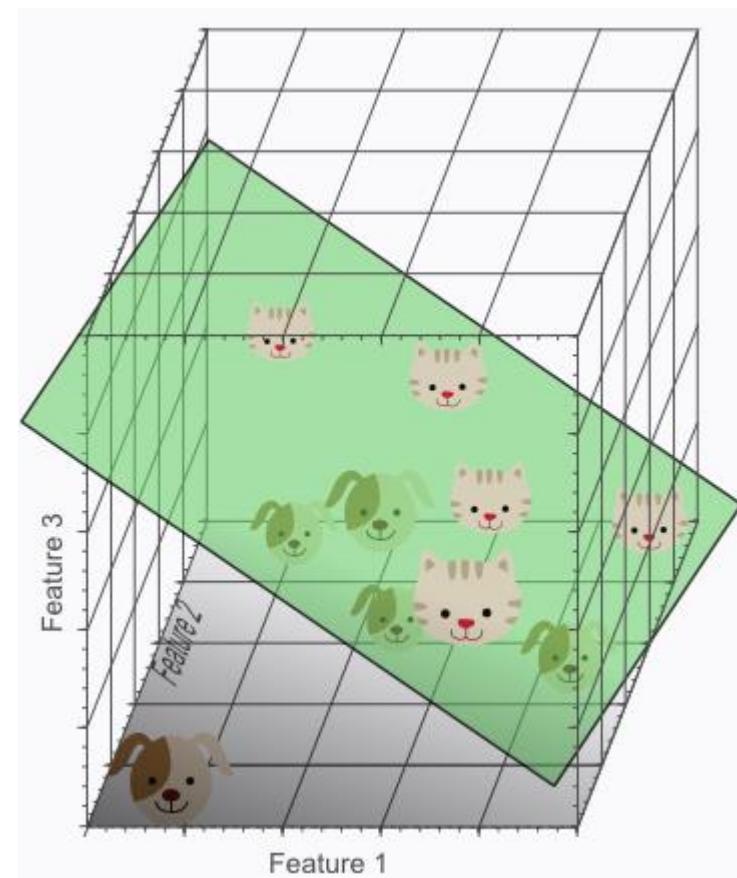
Curse of Dimensionality



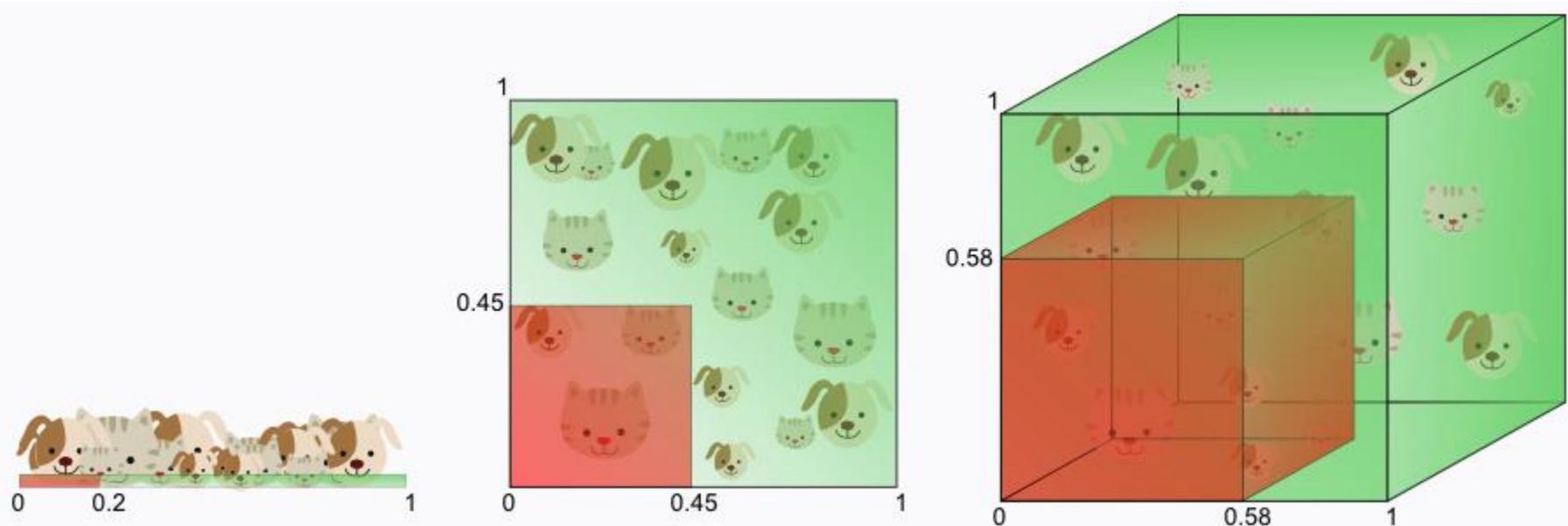
Curse of Dimensionality



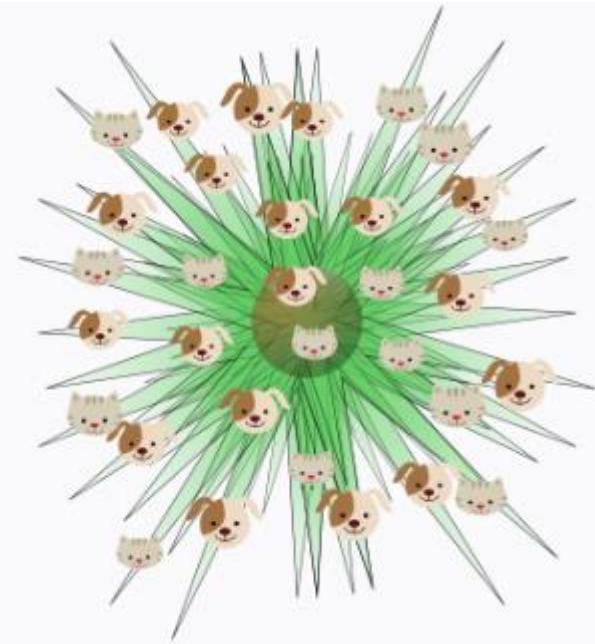
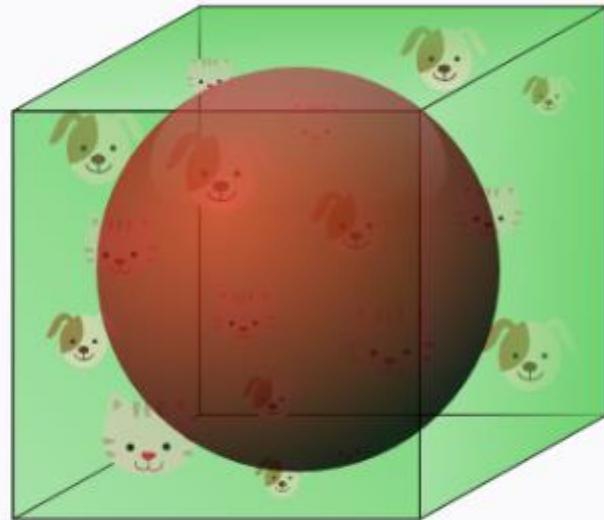
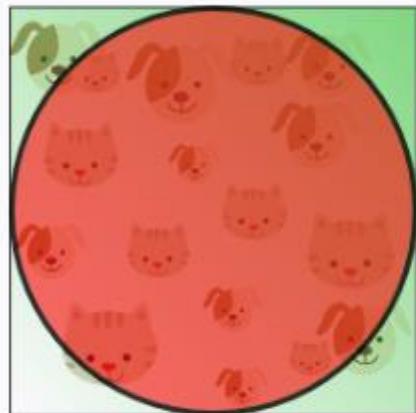
Curse of Dimensionality



Curse of Dimensionality



Curse of Dimensionality

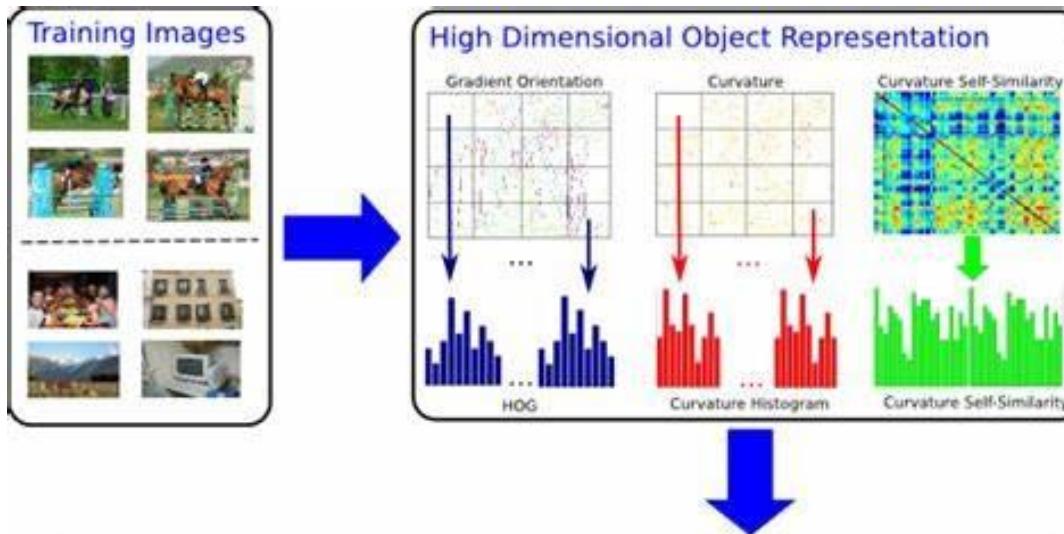


Curse of Dimensionality

Solutions:

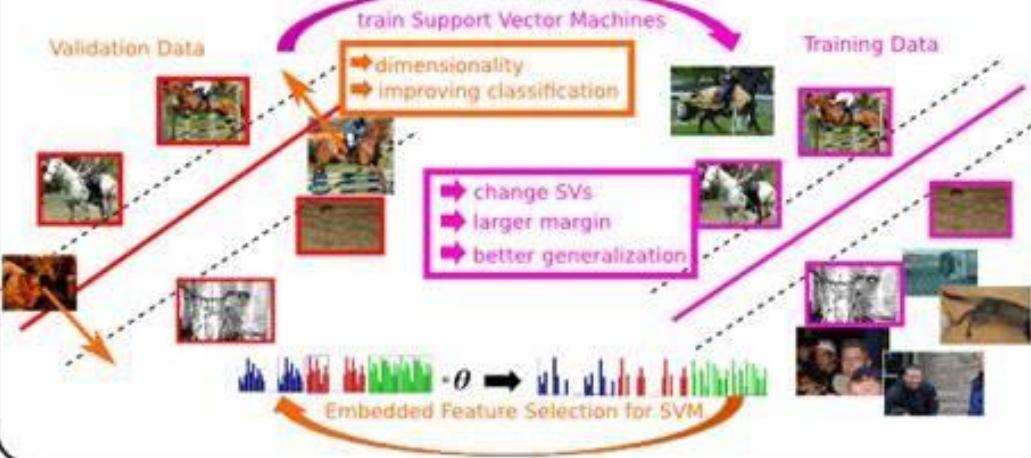
- Feature selection
- Projection of data onto a lower-dimensional space

Feature selection

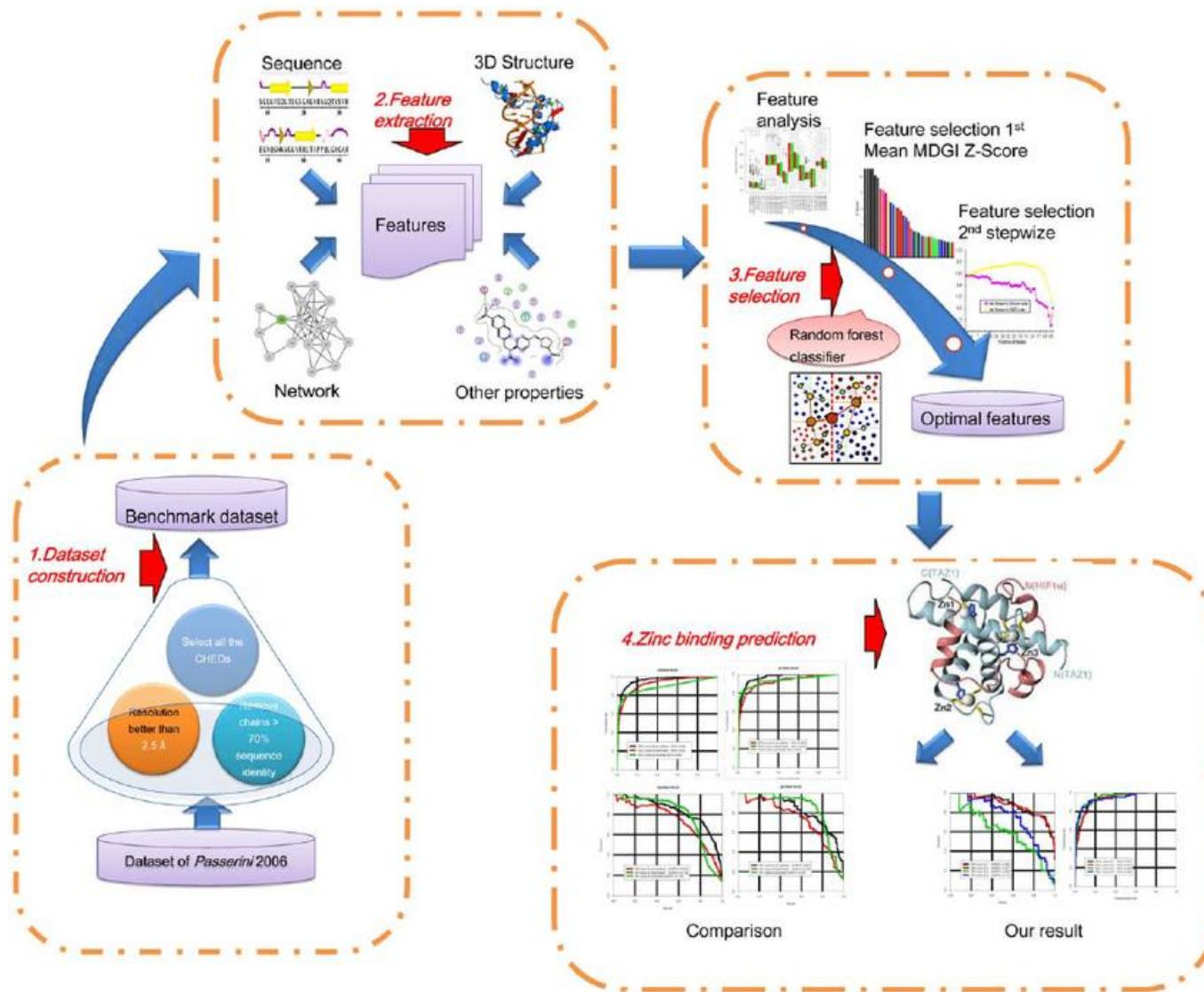


Embedded Feature Selection for Support Vector Machines

$$\min_{\theta} \min_{w, b, \xi} \frac{1}{2} \|w\|_2^2 + C \sum_{i=1}^N \xi_i \quad \text{subject to: } y_i (w^T \psi(\theta \cdot x_i) + b) \geq 1 - \xi_i \quad \wedge \quad \xi_i \geq 0 \quad \wedge \quad \|\theta\|_1 \leq \theta_1$$

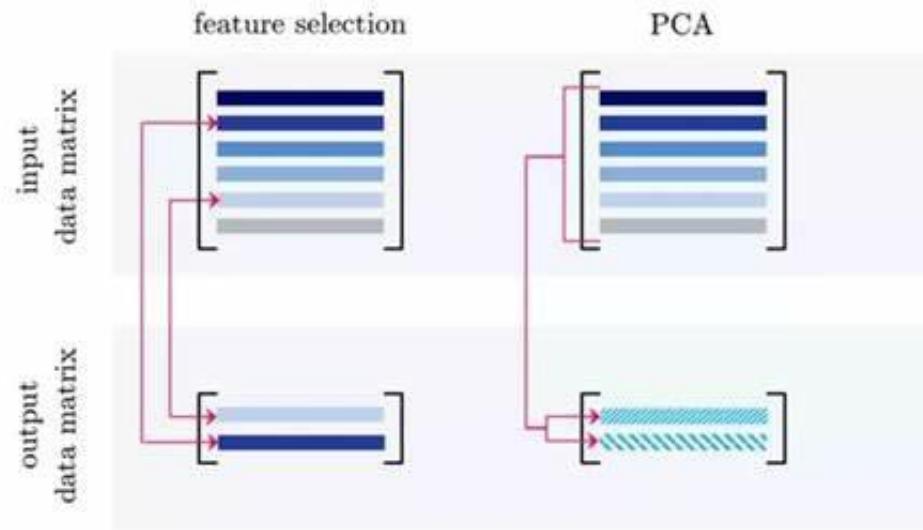


Feature selection



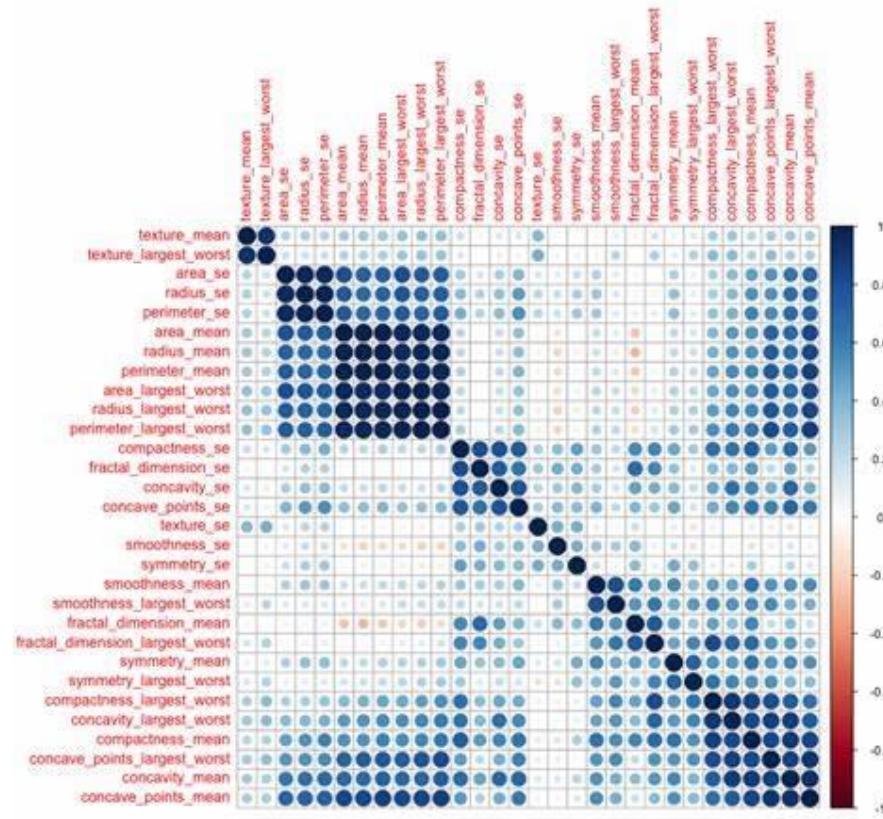
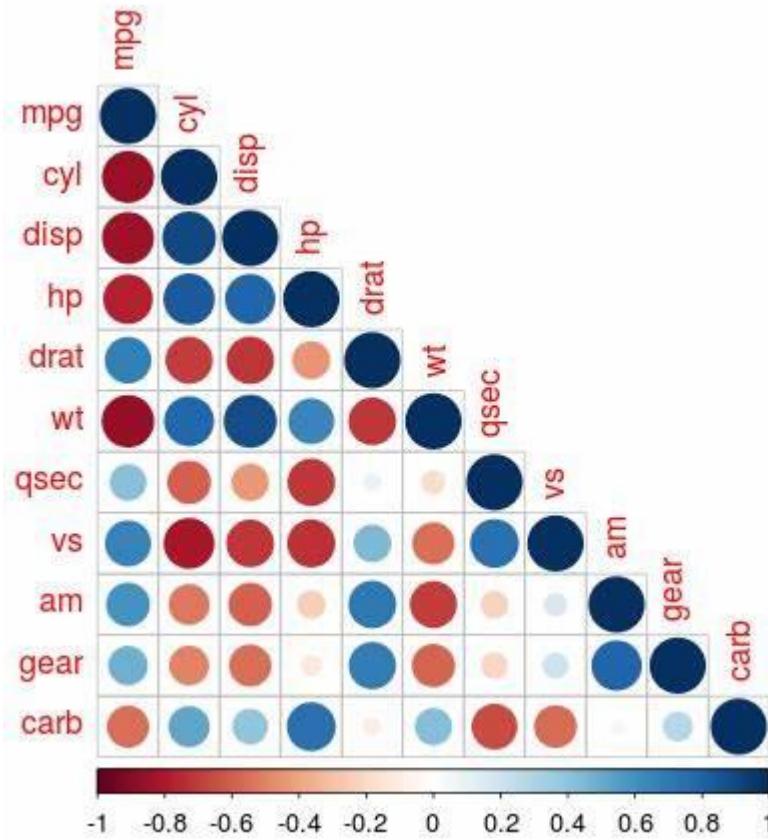
Feature selection

Reason 1: representativeness of features



Feature selection

Reason 2: reduction of redundant features



Feature selection

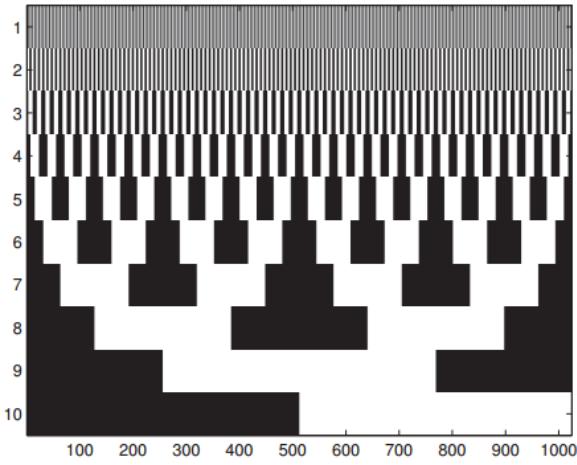
下面是特征选择/稀疏性非常有用的一些例子：

- 在许多问题中，我们的特征的数目 D 远大于我们的训练样本数 N 。这个就叫做 N 小 D 大问题，这样的问题很容易导致过拟合。所以我们就要降低特征也就是 D 的数目。而且在这个时代这样的现象越来越多，因为我们的传感器越来越丰富，所能够获得的信息的维度越来越高。
- 在信号处理中，常用小波基函数表示信号(图像、语音等)。为了节省时间和空间，找到信号的稀疏表示形式是有用的，这种稀疏表示形式是用少量的这种基函数表示的。这使我们能够从少量的测量中估计信号，以及压缩信号。

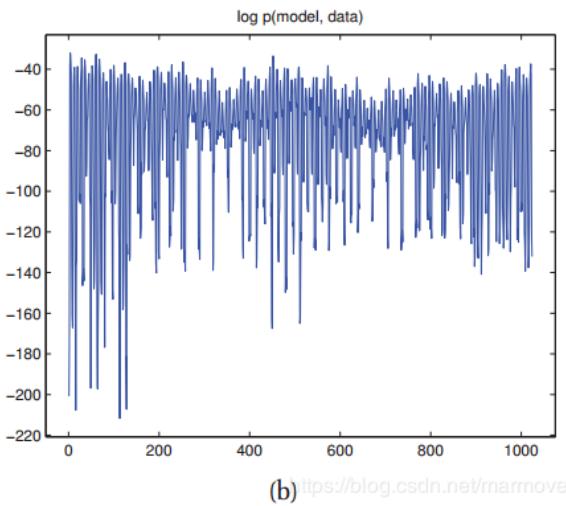
因此我们可以看到特征选择和稀疏性是当前机器学习/统计中最活跃的领域之一。

Feature selection

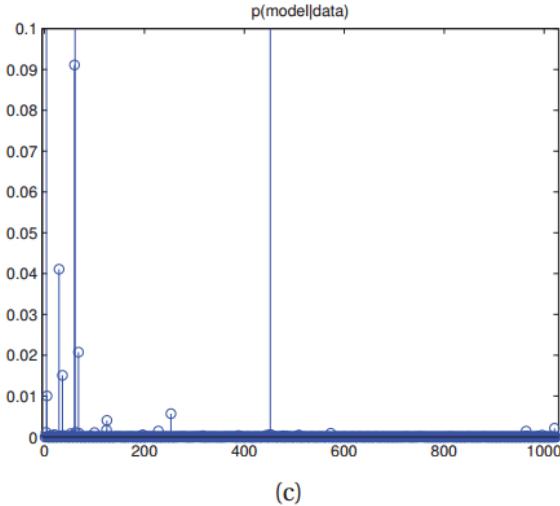
Subset selection



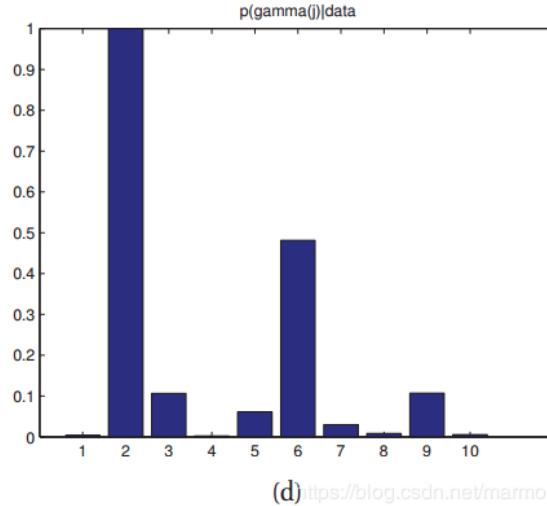
(a)



(b) <https://blog.csdn.net/marmmove>



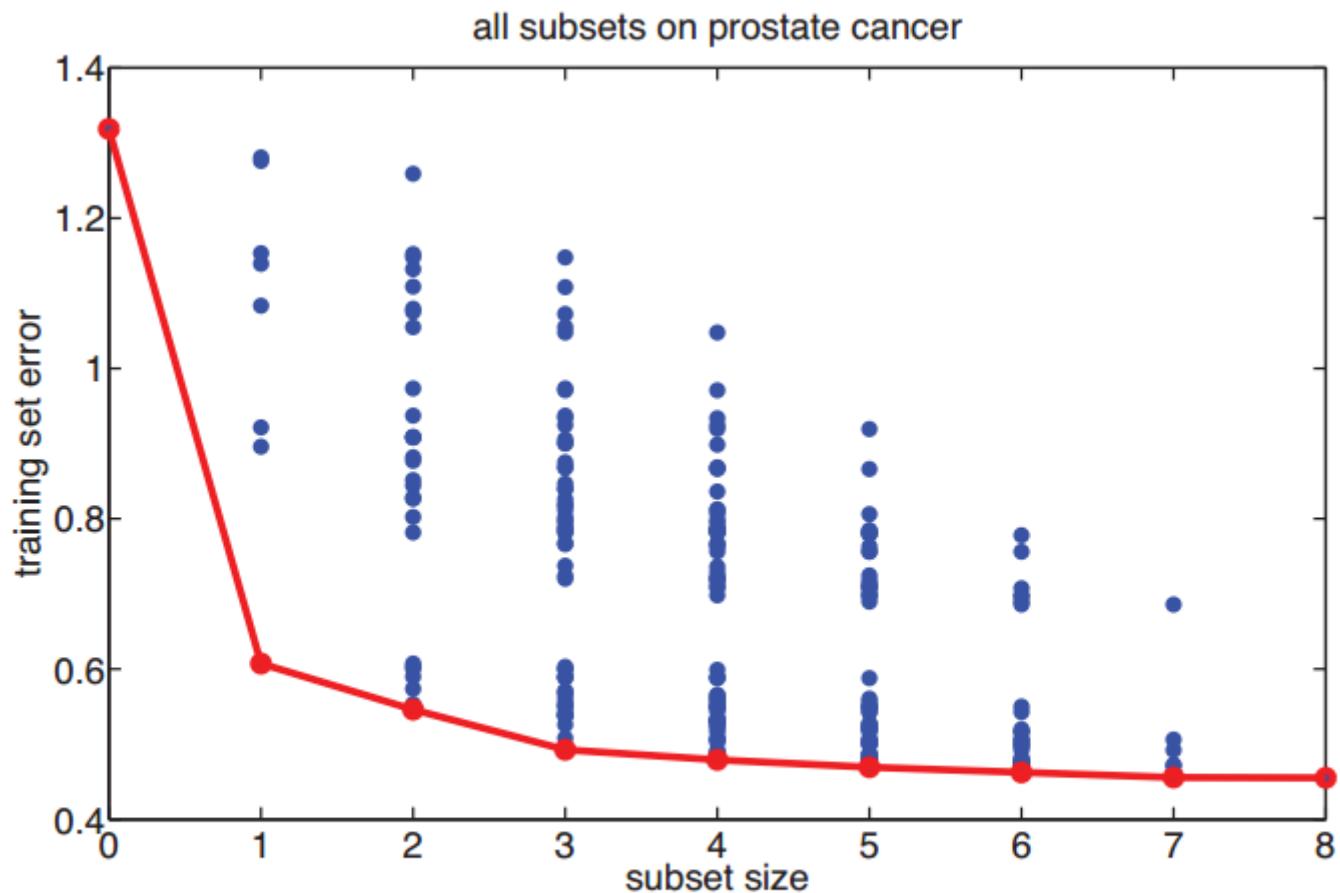
(c)



(d) <https://blog.csdn.net/marmmove>

model	prob	members
4	0.447	2,
61	0.241	2, 6,
452	0.103	2, 6, 9,
60	0.091	2, 3, 6,
29	0.041	2, 5,
68	0.021	2, 6, 7,
36	0.015	2, 5, 6,
5	0.010	2, 3,

Feature selection



(b)

<https://blog.csdn.net/marmove>

Feature selection

- mRMR (minimum Redundancy Maximum Relevance Feature Selection)
- Random forest

Feature selection

mRMR (minimum Redundancy Maximum Relevance Feature Selection)

[[Frequently Asked Questions](#) | [Online Version](#) | [C/C++ Version with Source Codes](#) | [Matlab Version](#) |
[Sample Data Sets](#) | [Major Publications](#) | [BIBM'07 Tutorial Slides](#)]

-
- **Online Version:** You can run the program using your own data through the following form, -- you can also download the program and run on your own machines (see links below).

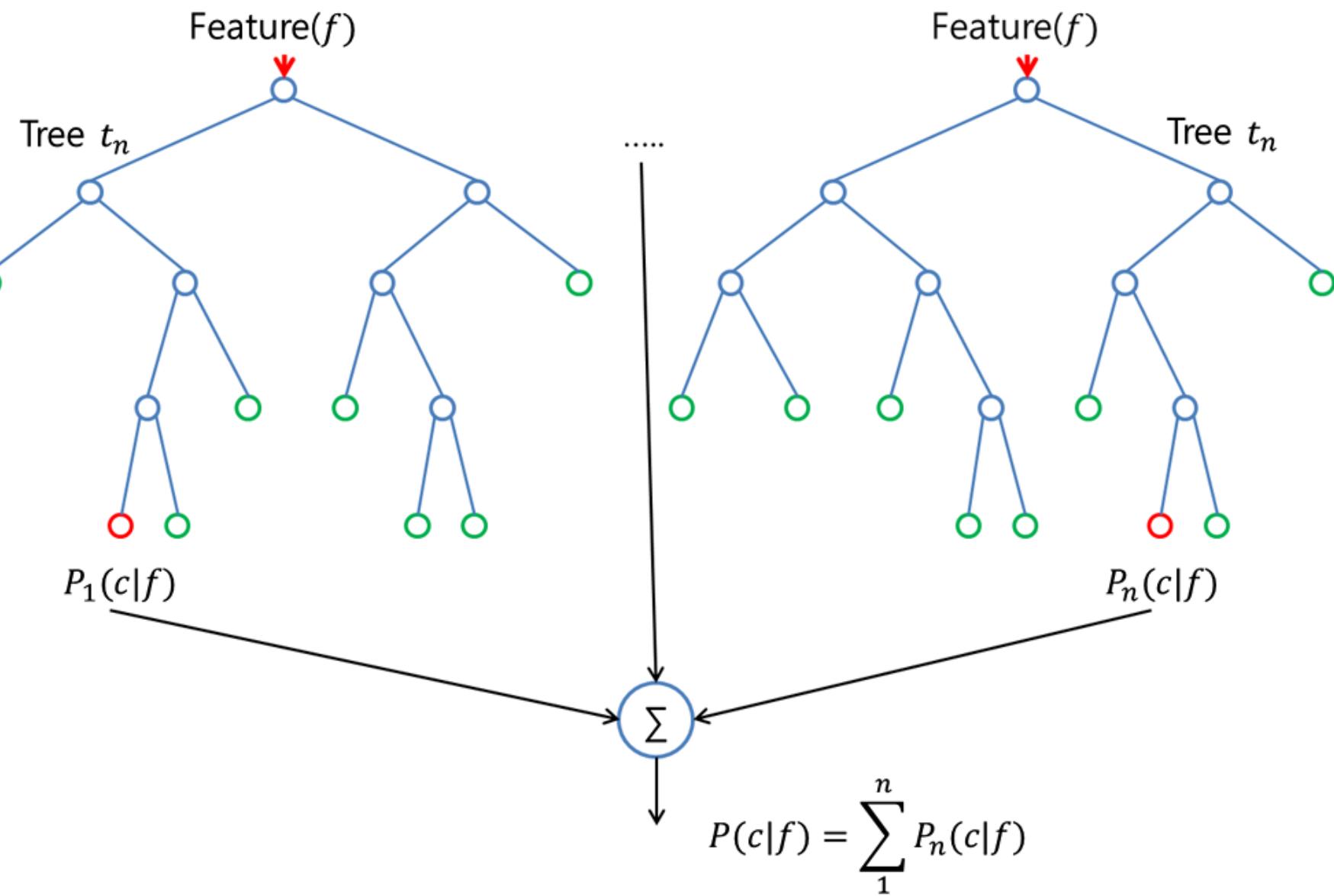
* Data file (Standard CSV file format, where each row is a sample and each column is a variable/attribute/feature. **MAKE SURE YOUR DATA IS SEPARATED BY COMMA, BUT NOT BLANK SPACE OR OTHER CHARACTERS!!** The first row must be the feature names, and the first column must be the classes for samples. You may download a testing example data set [here](#), which is micrarray data of lung cancer (7 classes). The data has been discretized as 3-states. Note that the web-based program can only accept a data file with the maximum size 2M bytes, and maximum number of variables = 10000 -- if you have a larger data set, you should download the program and run on your own machine (see download links below).

未选择任何文件

* What is the feature selection scheme you want to use:

* How many features you want to select:

Random forest



Curse of Dimensionality

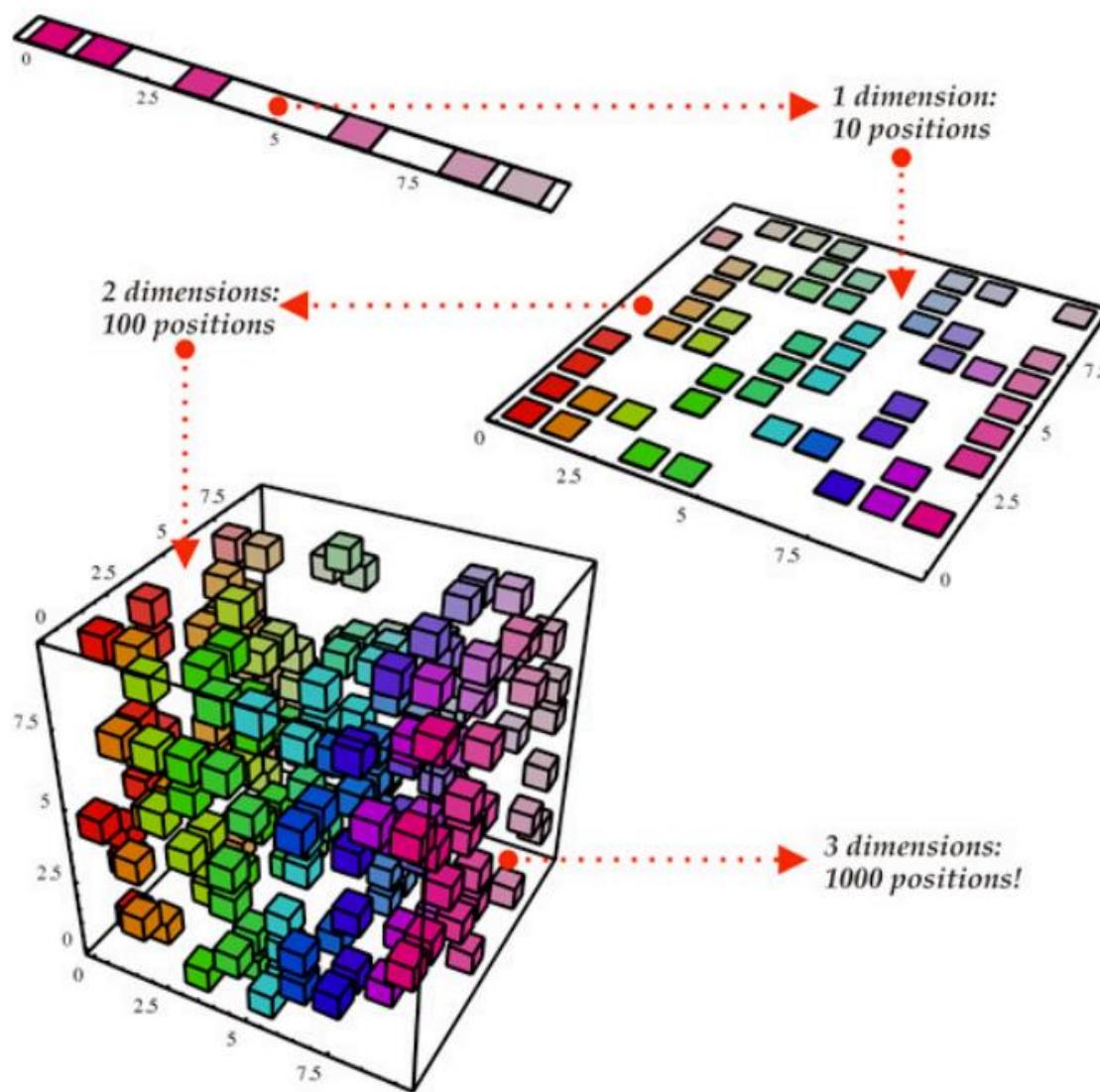
Solutions:

- Feature selection
- Projection of data onto a lower-dimensional space

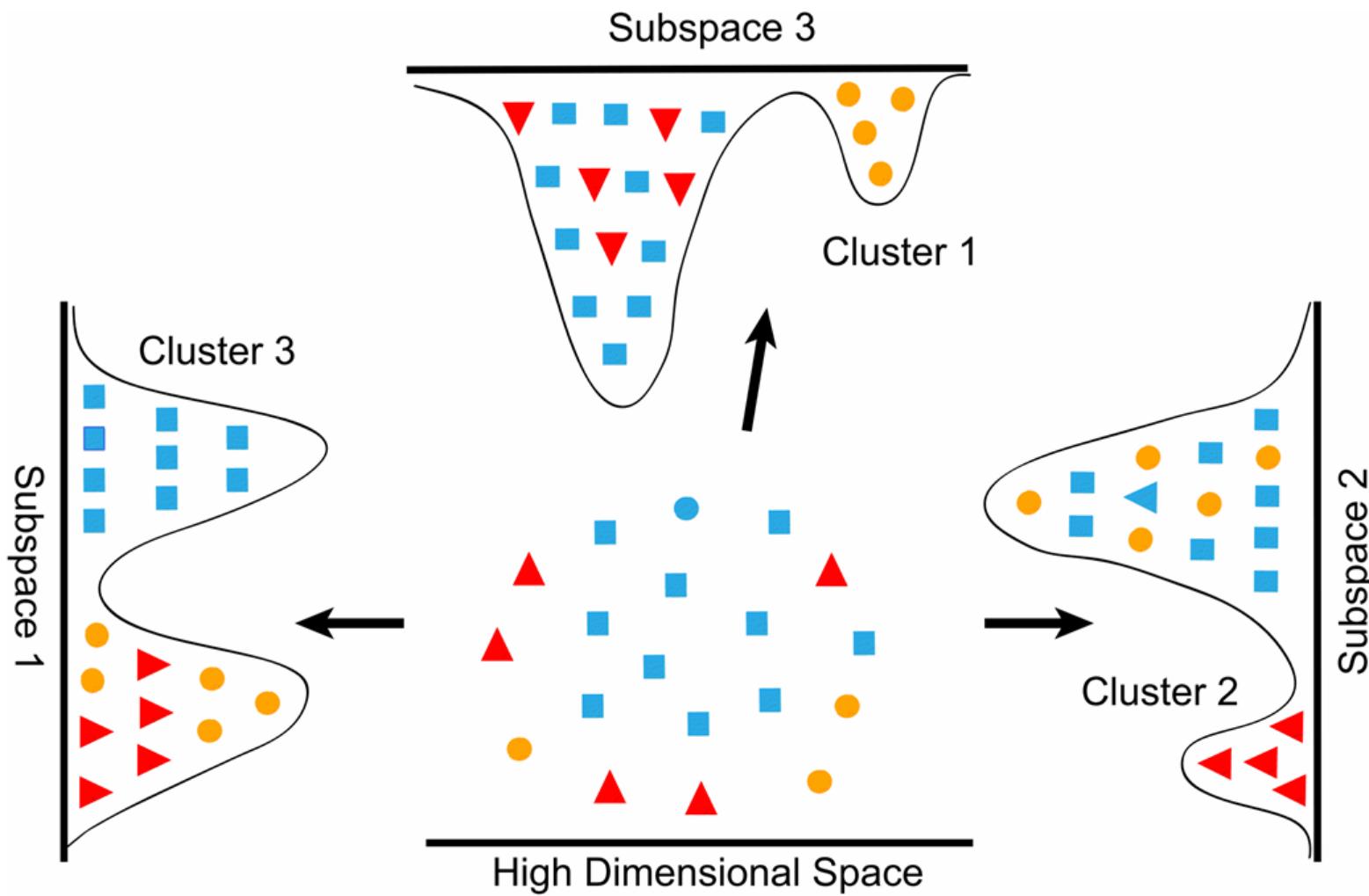
Dimension reduction illusions



Dimension Reduction



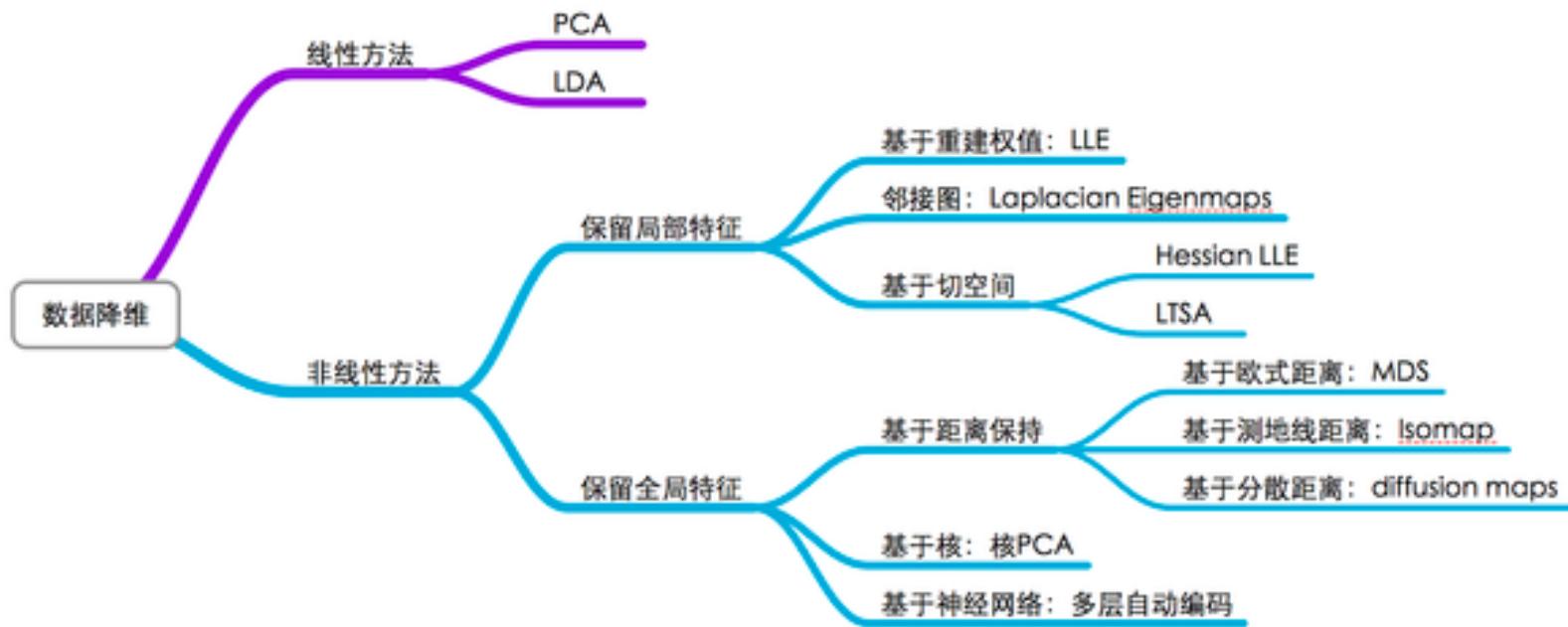
Dimension Reduction



Dimension Reduction

- One way to avoid the curse of dimensionality is by projecting the data onto a lower-dimensional space.
- Techniques for dimension reduction:
 - Principal Component Analysis (PCA)
 - Singular value decomposition (SVD)
 - Multi-dimensional Scaling (MDS).

Dimension Reduction



(1) 线性降维: PCA、ICA、LDA、LFA、LPP

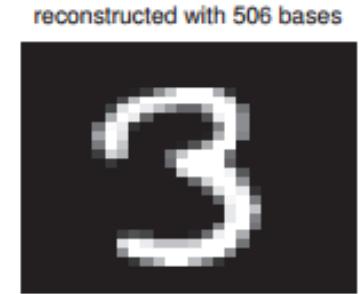
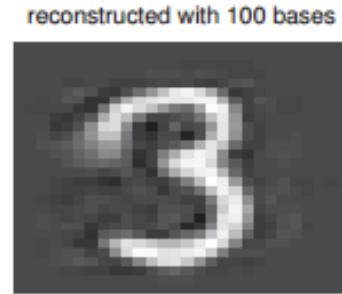
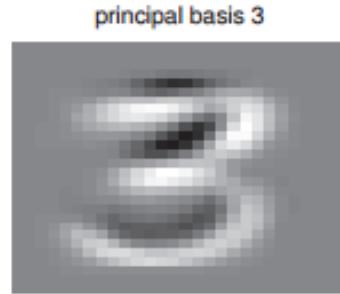
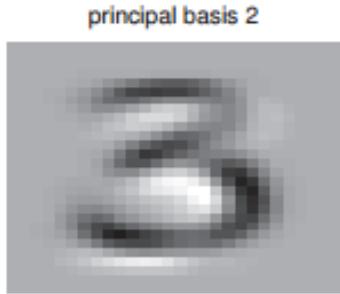
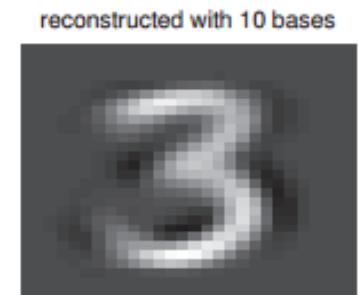
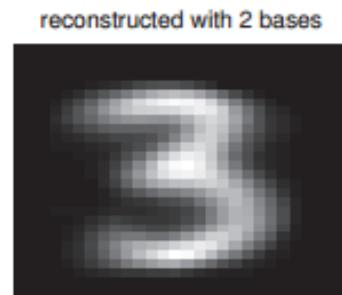
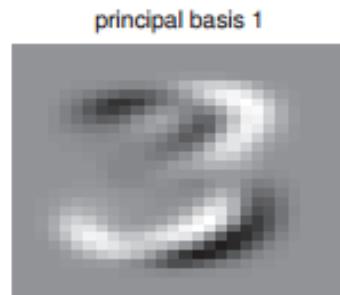
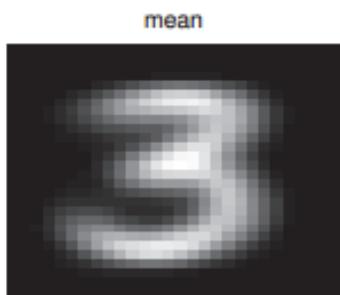
(2) 非线性降维方法:

① 基于核函数的方法: KPCA、KICA、KDA

② 基于特征值的方法: ISOMAP、LLE、LE、LPP、LTSA、MVU

Principle Component Analysis (PCA)

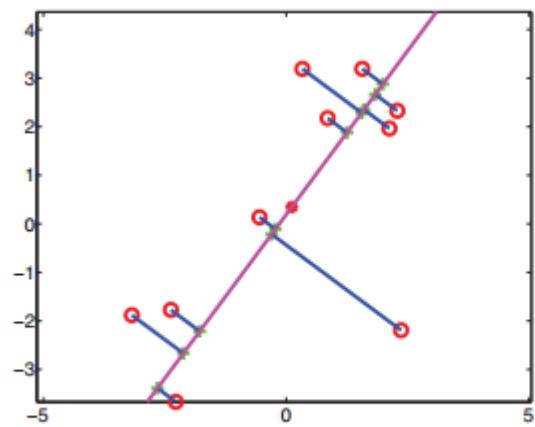
Dimension Reduction & PCA



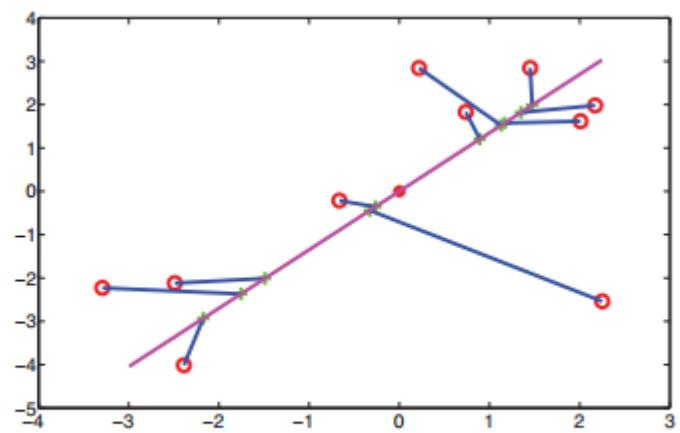
(a)

<https://log.csdn.net/marmmove>

Principal Component



(a)



http://blog.csdn.net/marnove

Dimension Reduction & PCA

- PCA reduce dimensions of data without much loss of information.
- Used in machine learning and in signal processing and image compression (among other things).

Variance and Covariance

- Let X, Y are random variable

$$Var(X) = E(X^2) - E(X)^2$$

$$\begin{aligned}Cov(X, Y) &= E[(X - EX)(Y - EY)] \\&= E(XY) - E(X)E(Y)\end{aligned}$$

Multivariate Random Variable

- For multidimensional random vector

$$\vec{X} = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_p \end{pmatrix}$$

Variance and Covariance

- Variance of the multivariate variable
(commonly referred to as covariance)

$$Var(\vec{X}) = E[(\vec{X} - E(\vec{X}))(\vec{X} - E(\vec{X}))^T]$$

- For simplicity, we simply note as

$$Var(X) = E[(X - E(X))(X - E(X))^T]$$

Projection to 1D

- Let $Z = \alpha^T X$ be the projection, which is a random variable (1D)

$$E(Z) = \alpha^T E(X)$$

$$\begin{aligned}Var(Z) &= E(Z^2) - E(Z)^2 \\&= E[(\alpha^T X)(\alpha^T X)^T] - [(\alpha^T E(X))(\alpha^T E(X))^T] \\&= \alpha^T [E(XX^T) - E(X)E(X)^T]\alpha \\&= \alpha^T Var(X)\alpha\end{aligned}$$

Sample Variance and Covariance

- Given the i.i.d samples, $X^{(1)}, X^{(2)}, \dots, X^{(n)}$

$$\sigma_{ij} = \frac{1}{n-1} \sum_{k=1}^n (X_i^{(k)} - \bar{X}_i)(X_j^{(k)} - \bar{X}_j)$$

$$\bar{X}_i = \frac{1}{n} \sum_{k=1}^n X_i^{(k)}, \quad \bar{X}_j = \frac{1}{n} \sum_{k=1}^n X_j^{(k)},$$

$$\hat{\Sigma} = (\sigma_{ij})_{p \times p}$$

Principle Components

- The first principle component is the projections that catches the maximum variance, we can obtain a optimization problem

$$\begin{aligned} & \max_{\alpha} \alpha^T \hat{\Sigma} \alpha \\ & s.t. \alpha^T \alpha = 1 \end{aligned}$$

- The solution is the eigenvector corresponding to the largest eigenvalue

$$\hat{\Sigma} \alpha = \lambda \alpha$$

Principle Components

- Similarly, we can get the second, third principle components with extra orthogonal constraints,

$$\hat{\Sigma}\alpha_1 = \lambda_1\alpha_1$$

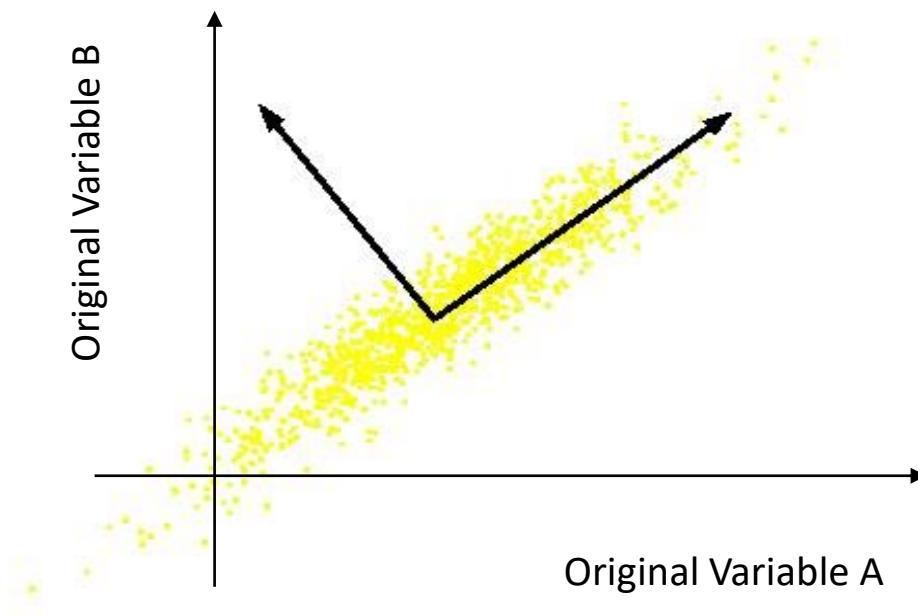
$$\hat{\Sigma}\alpha_2 = \lambda_2\alpha_2$$

.....

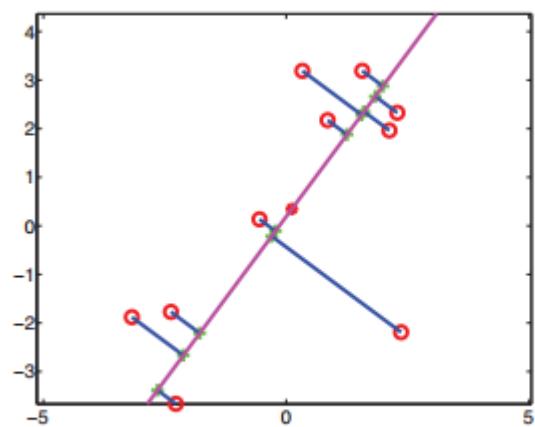
$$\hat{\Sigma}\alpha_p = \lambda_p\alpha_p$$

$$\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p$$

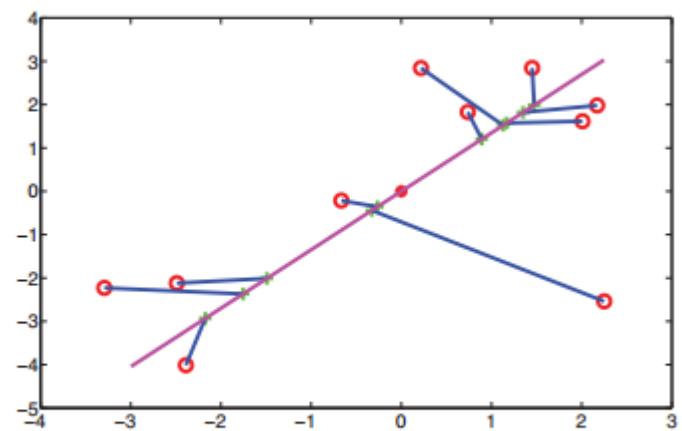
Principal Component



Principal Component



(a)



http://blog.csdn.net/marnove

Matrix Diagonalization

- For a covariance matrix $\hat{\Sigma}$, there exist a orthogonal matrix Q, such that

$$Q^T \hat{\Sigma} Q = \begin{pmatrix} \lambda_1 & 0 & \cdots & 0 \\ 0 & \lambda_2 & \cdots & 0 \\ \cdots & \cdots & \cdots & \cdots \\ 0 & 0 & \cdots & \lambda_p \end{pmatrix}$$

where $\lambda_1 \geq \lambda_2 \geq \cdots \lambda_p \geq 0$

Using Eigenvectors for Diagonalization

- Let $\alpha_i, i = 1, 2, \dots, p$ be a set of orthogonal unit eigenvectors with eigenvalue λ_i

$$\hat{\Sigma} \alpha_1 = \lambda_1 \alpha_1$$

$$\hat{\Sigma} \alpha_2 = \lambda_2 \alpha_2$$

.....

$$\hat{\Sigma} \alpha_p = \lambda_p \alpha_p$$

$$\begin{aligned}
& \left(\begin{array}{c} \alpha_1^T \\ \alpha_2^T \\ \dots \\ \alpha_p^T \end{array} \right) \hat{\Sigma} (\alpha_1, \alpha_2, \dots, \alpha_p) \\
&= \left(\begin{array}{c} \alpha_1^T \\ \alpha_2^T \\ \dots \\ \alpha_p^T \end{array} \right) (\lambda_1 \alpha_1, \lambda_2 \alpha_2, \dots, \lambda_p \alpha_p) \\
&= \left(\begin{array}{cccc} \lambda_1 \alpha_1^T \alpha_1 & \lambda_2 \alpha_1^T \alpha_2 & \dots & \lambda_p \alpha_1^T \alpha_p \\ \lambda_1 \alpha_2^T \alpha_1 & \lambda_2 \alpha_2^T \alpha_2 & \dots & \lambda_p \alpha_2^T \alpha_p \\ \dots & \dots & \dots & \dots \\ \lambda_1 \alpha_p^T \alpha_1 & \lambda_2 \alpha_p^T \alpha_2 & \dots & \lambda_p \alpha_p^T \alpha_p \end{array} \right) \\
&= \text{diag}\{\lambda_1, \lambda_2, \dots, \lambda_p\}
\end{aligned}$$

Principle Components

- Given p orthogonal unit eigenvectors α_i

$$\hat{\Sigma}\alpha_i = \lambda_i\alpha_i, i = 1, 2, \dots, p$$

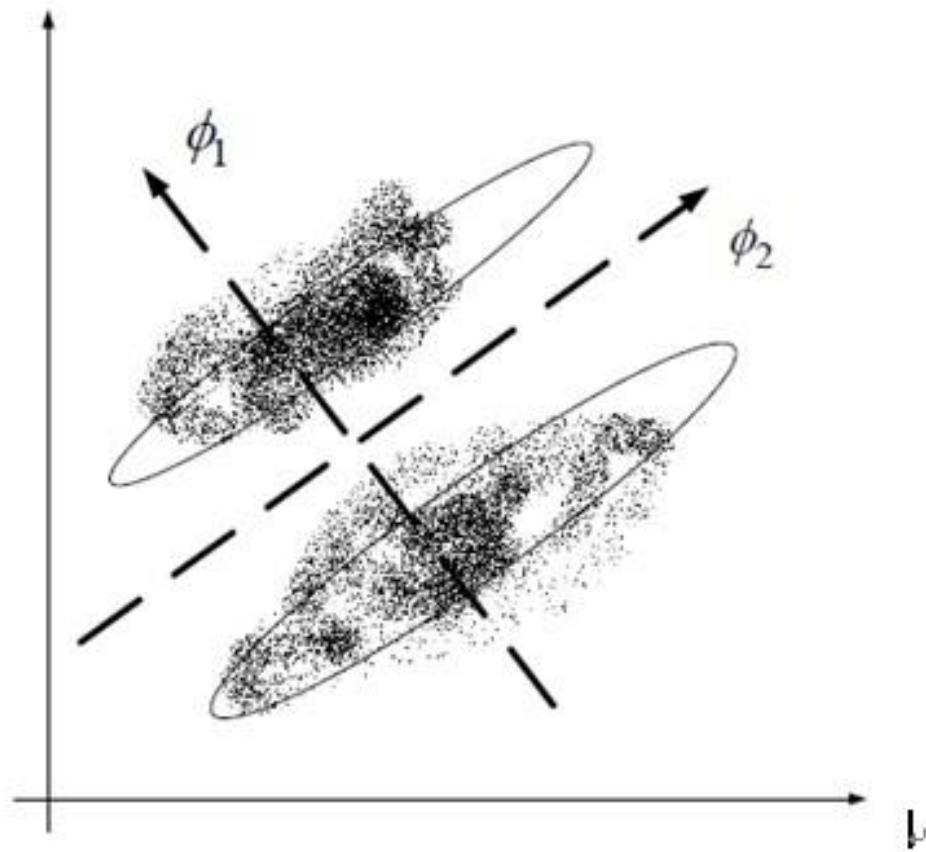
- Let $Z_i = \alpha_i^T X$, i.e $Z = Q^T X$, where

$$Z = \begin{pmatrix} Z_1 \\ Z_2 \\ \dots \\ Z_p \end{pmatrix} \quad Q = (\alpha_1, \alpha_2, \dots, \alpha_p)_{p \times p}$$

PCA in R

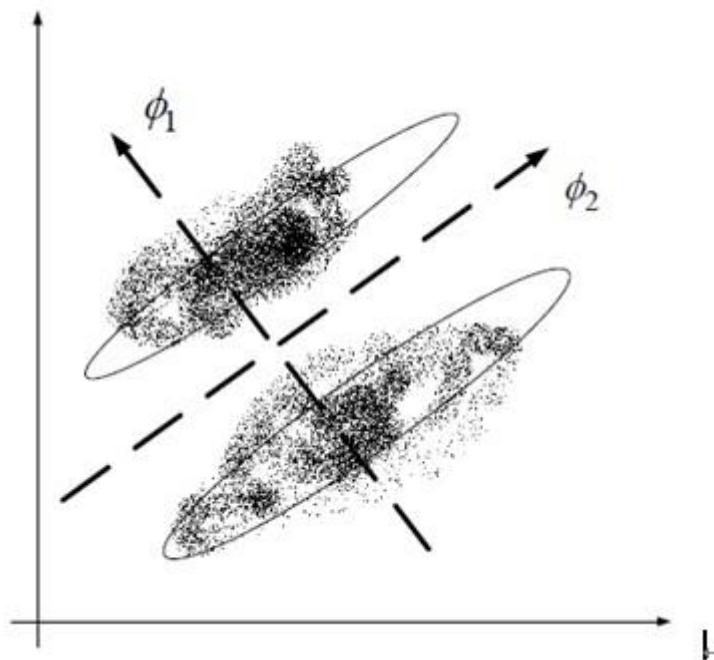
- `prcomp(stats)`
 - Principal Components Analysis (preferred)
- `princomp(stats)`
 - Principal Components Analysis
- `screeplot(stats)`
 - Screeplot of PCA Results
- `summary(obj)`
- `loadings(obj)`

PCA problems



PCA追求的是在降维之后能够最大化保持数据的内在信息，并通过衡量在投影方向上的数据方差的大小来衡量该方向的重要性。但是这样投影以后对数据的区分作用并不大，反而可能使得数据点揉杂在一起无法区分。

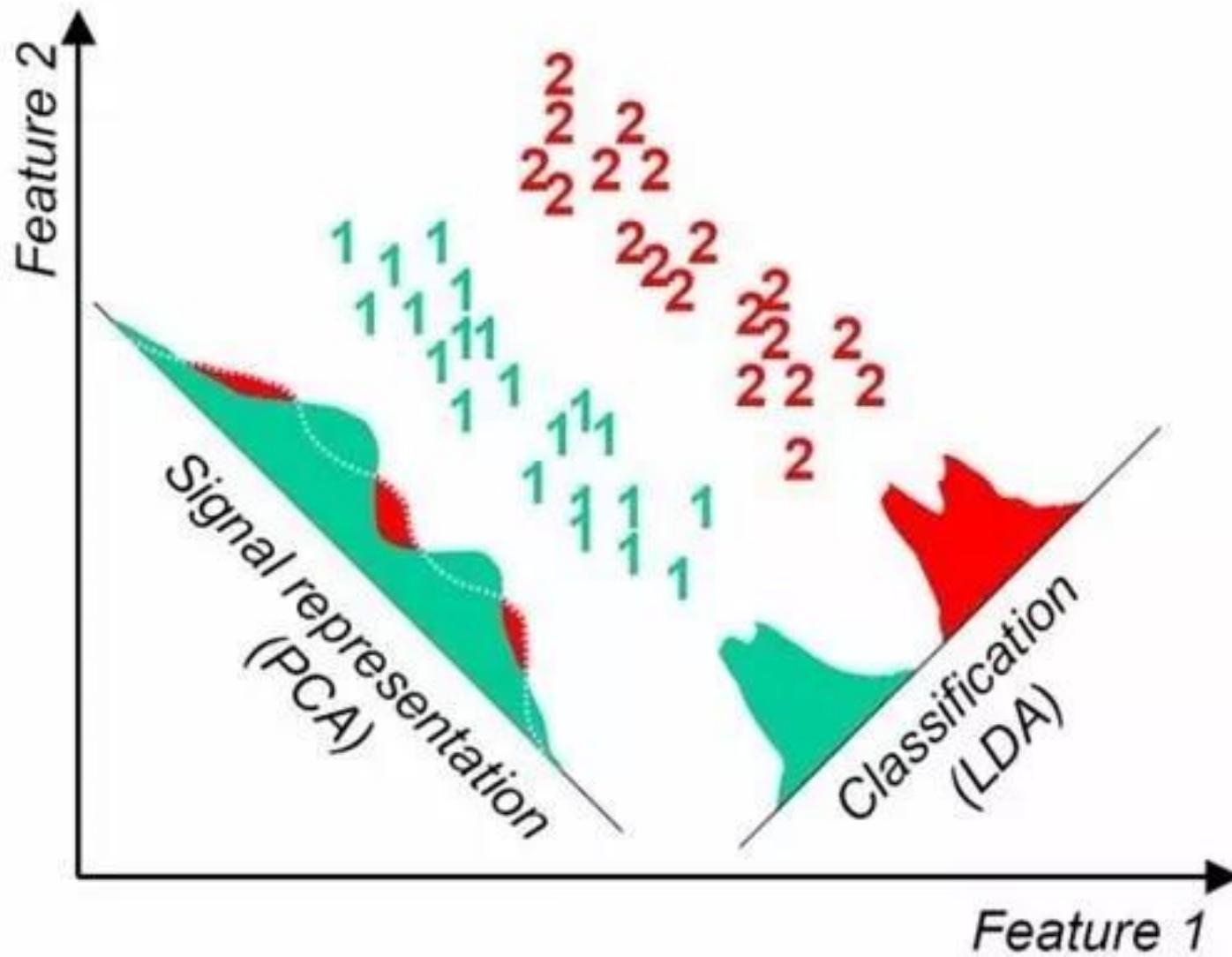
PCA & LDA



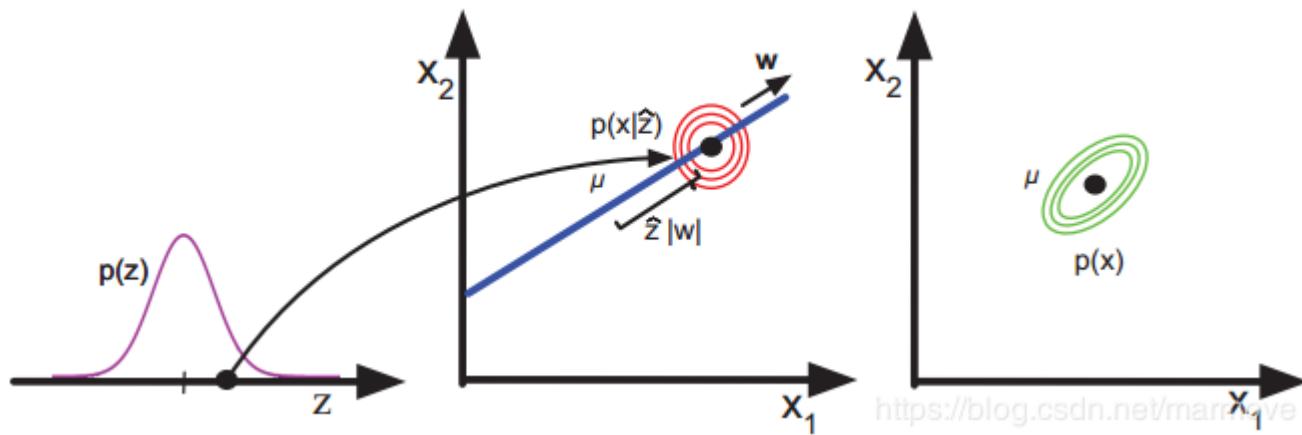
LDA:

- 1、同类的数据点尽可能的接近（**within class**）
- 2、不同类的数据点尽可能的分开（**between class**）

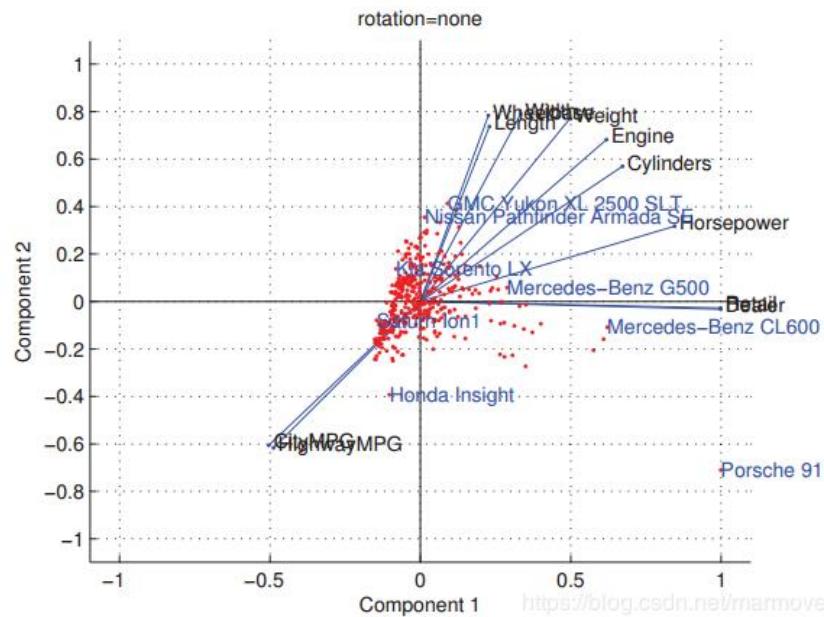
PCA & LDA



factor analysis (要素分析)



<https://blog.csdn.net/marmove>



PCoA

CCA

<https://blog.csdn.net/marmove>

factor analysis (要素分析)

CCA: canonical correlation analysis

典型相关分析

$$\rho(X, Y) = \frac{\text{cov}(X, Y)}{\sqrt{D(X)} \sqrt{D(Y)}}$$

- 例子：我们拿到两组数据，第一组是人身高和体重的数据，第二组是对应的跑步能力和跳远能力的数据。那么我们能不能说这两组数据是相关的呢？
- 如果X是包括人身高和体重两个维度的数据，而Y是包括跑步能力和跳远能力两个维度的数据，就不能直接使用相关系数的方法。那我们能不能变通一下呢？
- CCA使用的方法是将多维的X和Y都用线性变换为1维的X'和Y'，然后再使用相关系数来看X'和Y'的相关性。
- 对于我们的CCA，它选择的投影标准是降维到1维后，两组数据的相关系数最大。

factor analysis (要素分析)

CCA: canonical correlation analysis
典型相关分析

CCA的SVD解法:

输入: 各为m个的样本X和Y, X和Y的维度都大于1

输出: X,Y的相关系数 ρ ,X和Y的线性系数向量a和b

1. 计算X的方差 S_{XX} , Y的方差 S_{YY} , X和Y的协方差 S_{XY} , Y和X的协方差 $S_{YX} = S_{XY}^T$
2. 计算矩阵 $M = S_{XX}^{-1/2} S_{XY} S_{YY}^{-1/2}$
3. 对矩阵M进行奇异值分解, 得到最大的奇异值 ρ , 和最大奇异值对应的左右奇异向量u,v
4. 算X和Y的线性系数向量a和b, $a = S_{XX}^{-1/2} u$, $b = S_{YY}^{-1/2} v$

factor analysis (要素分析)

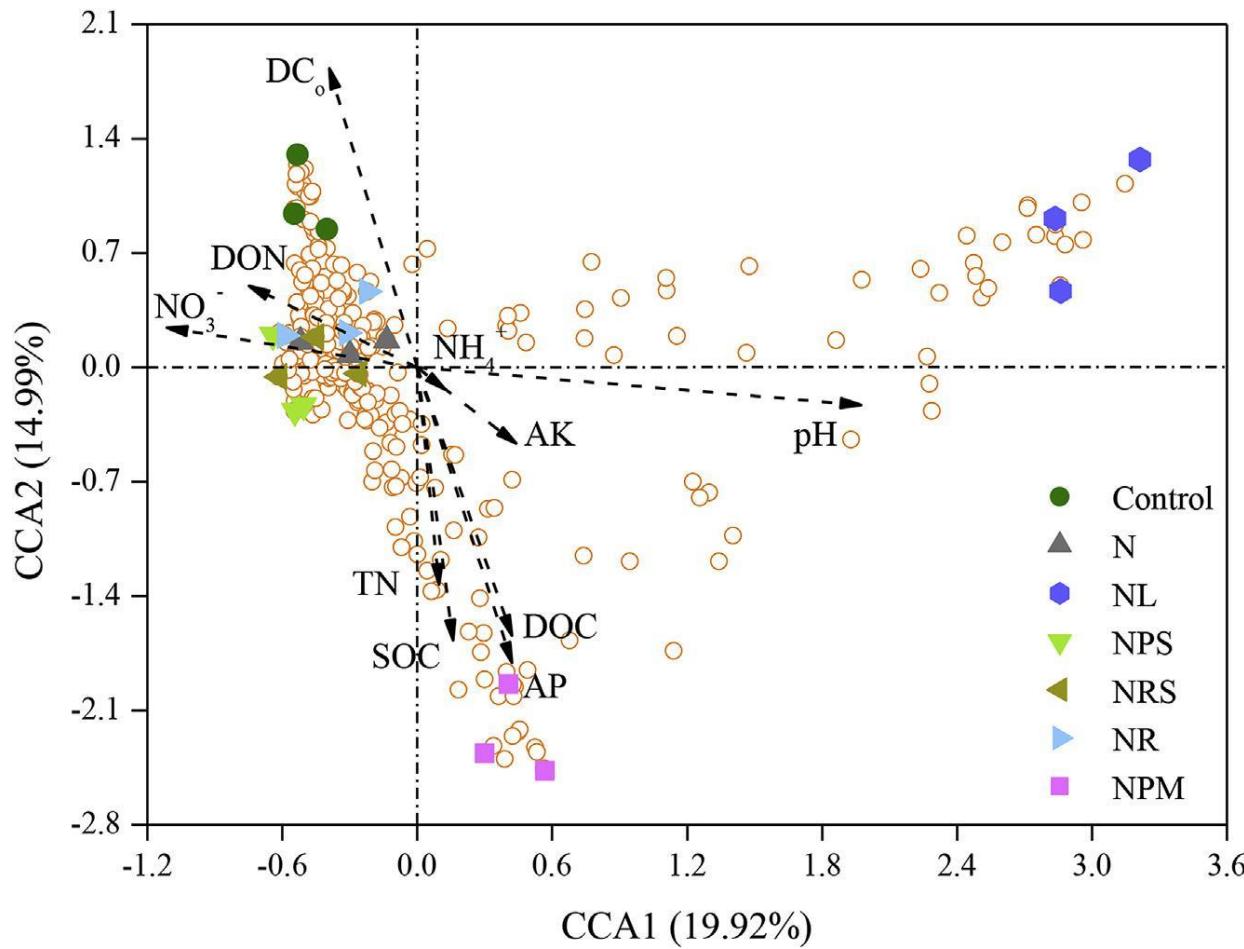
CCA: canonical correlation analysis

典型相关分析

- 我们在算法里只找了相关度最大的奇异值或者特征值
- 作为数据的相关系数，实际上我们也可以像PCA一样找出第二大奇异值，第三大奇异值，。。。得到第二相关系数和第三相关系数

factor analysis (要素分析)

CCA: canonical correlation analysis
典型相关分析



PCA, LDA & CCA

- **主成分分析PCA:** 降维的原则是投影方差最大
- **线性判别分析LDA:** 降维的原则是同类的投影方差小，异类间的投影方差大
- **典型相关分析CCA:** 降维的原则是降维到1维后，两组数据的相关系数最大

$$\underbrace{\arg \max_{a,b}}_{\text{arg max}_{a,b}} \frac{\text{cov}(X', Y')}{\sqrt{D(X')}} \sqrt{D(Y')}$$

奇异值分解(SVD)及其应用

- 奇异值分解(SVD, Singular Value Decomposition)
- Application I: Information retrieval
- Application II: Gene expression data analysis

$$\begin{matrix} & D \\ \begin{matrix} N \\ \textcolor{brown}{X} \end{matrix} & = \end{matrix} \begin{matrix} D & N-D \\ \begin{matrix} \textcolor{brown}{U} \end{matrix} & \end{matrix} \begin{matrix} D \\ \begin{matrix} \sigma_1 & \dots & \sigma_D \\ 0 \end{matrix} \end{matrix} \begin{matrix} D \\ \begin{matrix} \textcolor{brown}{V}^T \end{matrix} \end{matrix}$$

奇异值分解(SVD)

- 定理1：设 A 是秩为 r 的实值矩阵，则存在 m 阶正交矩阵 U ， n 阶正交矩阵 V ，使得 $A=UDV^T$. 其中 D 为对角矩阵 $D_{m \times n} = (\text{diag}\{\sigma_1, \dots, \sigma_r, 0, \dots, 0\}, 0)$ ，满足

$$\sigma_1 \geq \sigma_2 \geq \dots \geq \sigma_r > 0, \text{ 则}$$

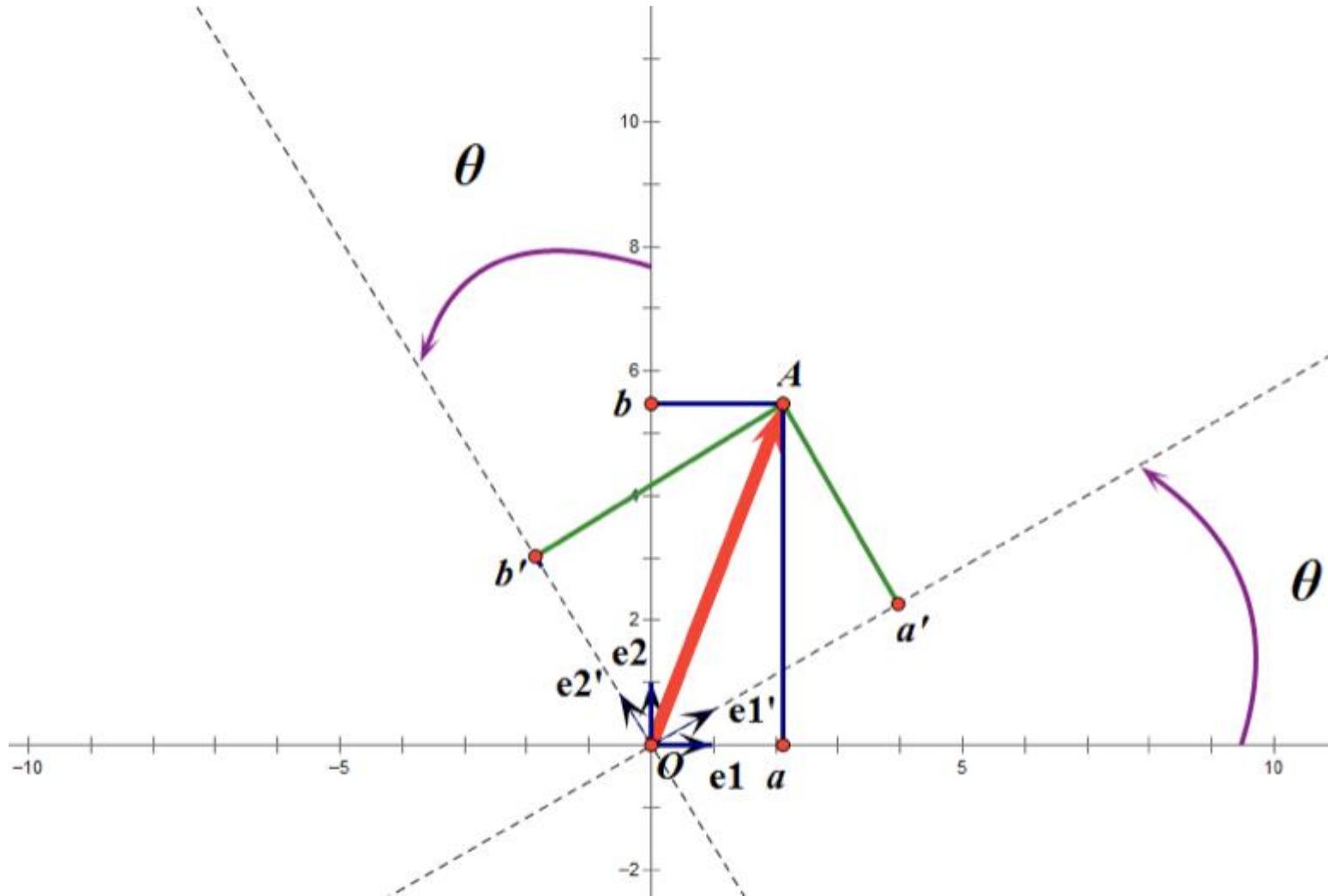
$$\text{rank}(A) = r, N(A) = \text{Span}\{v_{r+1}, \dots, v_n\},$$

$$R(A) = \text{Span}\{u_1, \dots, u_r\}.$$

$$A = \sum_{i=1}^r \sigma_i u_i v_i^T$$

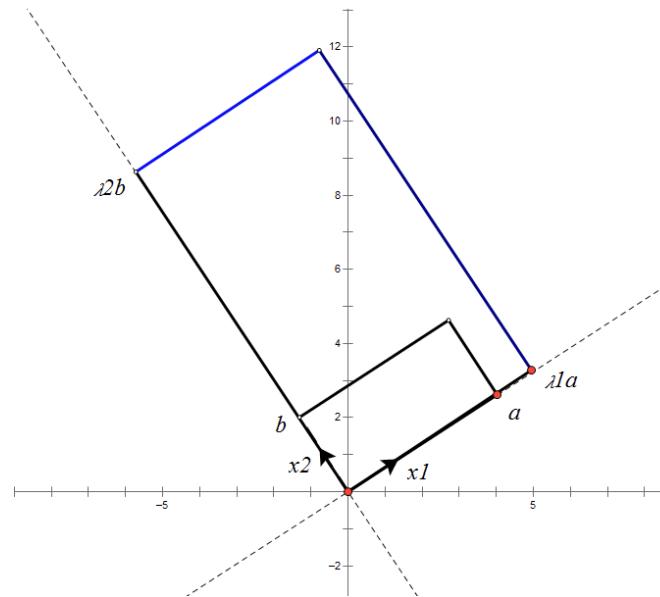
$$\|A\|_F^2 = \sigma_1^2 + \dots + \sigma_r^2$$

奇异值分解(SVD)



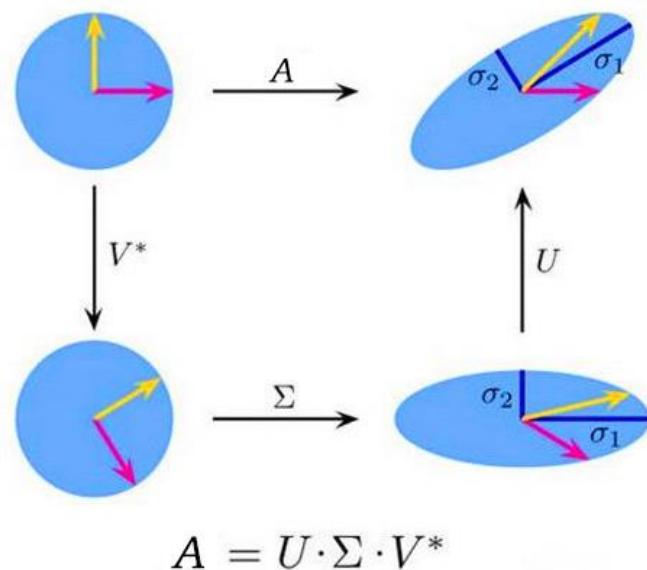
奇异值分解(SVD)

- 旋转是相对的：我们可以说向量的空间位置没有变，标准参考系向左旋转了 θ 角度，而如果我选择了 e_1' 、 e_2' 作为新的标准坐标系，那么在新坐标系中 OA （原标准坐标系的表示）就变成了 OA' ，这样看来就好像坐标系不动。
- 正交矩阵的行（列）向量都是两两正交的单位向量，正交矩阵对应的变换为正交变换，它有两种表现：旋转和反射。正交矩阵将标准正交基映射为标准正交基（从 e_1 、 e_2 到 e_1' 、 e_2' ）。



奇异值分解(SVD)

- 称 σ_i 为矩阵 A 的奇异值。
- 称 U, V 的列向量分别称为 A 的左、右奇异向量。



奇异值分解(SVD)

- $A^T A_{n,n}$ 和 $AA^T_{m,m}$ 都是实对称矩阵，则存在 n 阶正交矩阵 V 和 m 阶正交矩阵 U ，使得

$$U(AA^T)U^T = D_1$$

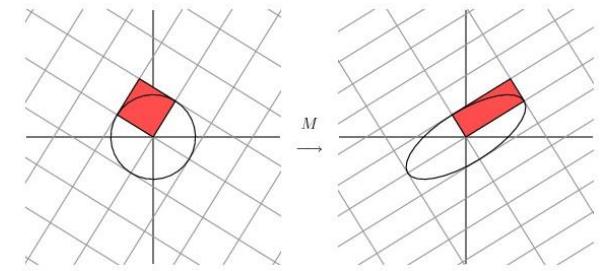
$$V(A^T A)V^T = D_2$$

- $(A^T A)$ 和 (AA^T) 的特征值相同

$$\text{If } (A^T A)\eta = \lambda\eta, \text{ then } (AA^T)A\eta = \lambda A\eta$$

- $A^T A$ 的非 0 特征值的个数等于 A 的秩 r .

$$A_{m \times n} = U_{m \times r} \times \Sigma_{r \times r} \times V^T_{r \times n}$$



Rank, Range, and Null space

- The rank of matrix A can be calculated from SVD by the number of nonzero singular values.
- The range of matrix A is the left singular vectors of U corresponding to the non-zero singular values.
- The null space of matrix A is the right singular vectors of V corresponding to the zeroed singular values.

Rank, Range, and Null space

$$A_{m \times n} = U_{m \times m} W_{n \times n} V_{n \times n}^T$$

Range Rank Null Space

$$\begin{bmatrix} 0 & 0 & 1 \\ 0 & 1 & 0 \\ 0 & 0 & 0 \\ 1 & 0 & 0 \end{bmatrix} \begin{bmatrix} 0 \\ 0 \\ -1 \end{bmatrix} \begin{bmatrix} 4 & 0 & 0 & 0 & 0 \\ 0 & 3 & 0 & 0 & 0 \\ 0 & 0 & \sqrt{5} & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \\ \sqrt{0.2} & 0 & 0 \\ 0 & 0 & 0 \\ -\sqrt{0.8} & 0 & 0 \end{bmatrix} \begin{bmatrix} 0 & 0 \\ 0 & 0 \\ 0 & \sqrt{0.8} \\ 1 & 0 \\ 0 & \sqrt{0.2} \end{bmatrix}^T$$

$$A_{4 \times 5} = \begin{bmatrix} 0 & 0 & 1 \\ 0 & 1 & 0 \\ 0 & 0 & 0 \\ 1 & 0 & 0 \end{bmatrix}_{4 \times 3} \begin{bmatrix} 4 & 0 & 0 \\ 0 & 3 & 0 \\ 0 & 0 & \sqrt{5} \end{bmatrix}_{3 \times 3} \begin{bmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \\ \sqrt{0.2} & 0 & 0 \\ 0 & 0 & 0 \\ -\sqrt{0.8} & 0 & 0 \end{bmatrix}_{5 \times 3}^T$$

Rank(A)=3

Condition Number

- SVD can tell How close a square matrix A is to be singular.
- The ratio of the largest singular value to the smallest singular value can tell us how close a matrix is to be singular:

$$A = U \begin{bmatrix} \sigma_1 & 0 & \cdots & 0 \\ 0 & \sigma_2 & \cdots & 0 \\ \vdots & \vdots & \cdots & \vdots \\ 0 & 0 & 0 & \sigma_k \end{bmatrix} V^T, \quad c = \frac{\sigma_1}{\sigma_k}$$

- A is singular if c is infinite.
- A is ill-conditioned if c is too large (machine dependent).

奇异值分解(SVD)

- 定理2 (Eckart and Young): Let the SVD of A be given by $A=UDV^T$ with $r = \text{rank}(A) < p = \min(m, n)$ and define

$$A_k = \sum_{i=1}^k u_i \sigma_i v_i^T$$

Then

$$\min_{\text{rank}(B)=k} \|A - B\|_F^2 = \|A - A_k\|_F^2 = \sigma_{k+1}^2 + \cdots + \sigma_p^2$$

其中 $\|\cdot\|_F$ 为 Frobenius 范数, 定义为

$$\|A\|_F^2 = \sum_{i=1}^m \sum_{j=1}^n |a_{ij}|^2 = \text{tr}(A^* A) = \sum_{i=1}^{\min(m,n)} \sigma_i^2$$

奇异值分解(SVD)

- In other words, A_k , which is constructed from the k-largest singular triplets of A , is the closest rank-k matrix to A . In fact, A_k is the best approximation to A for any unitarily invariant norm[*].

* L. MIRSKY, Symmetric gage functions and unitarily invariant norms, Q. J. Math, 11(1960), pp. 50-59.

矩阵近似

- 前k个奇异值的贡献率

$$\rho = \frac{\sum_{j=1}^k \sigma_j^2}{\sum_{j=1}^r \sigma_j^2}$$

- 可以根据贡献率截取前k个奇异值近似矩阵

$$A \simeq \sum_{j=1}^k \sigma_j u_j v_j^T$$

SVD举例

$$A = \begin{pmatrix} 1 & 2 & 1 & 1 & 1 \\ 1 & 2 & 1 & 1 & 1 \\ 2 & 3 & 2 & 1 & 0 \end{pmatrix}$$

$$r(A) = 2, \sigma_1 = 5.70863, \sigma_2 = 1.188084, \sigma_3 = 0$$

$$U = (u_1, u_2, u_3)$$

$$U = \begin{pmatrix} 0.4836967 & 0.5157882 & 0.7071068 \\ 0.7294347 & -0.6840504 & 0 \\ 0.4836967 & 0.5157882 & -0.7071068 \end{pmatrix}$$

SVD举例

$$V = (v_1, v_2, v_3, v_4, v_5)$$

$$V = \begin{pmatrix} -0.4250167 & -0.28324973 & 0.00000000 & 0.85972695 & 0.00000000 \\ -0.7222558 & 0.00926002 & 0.50863102 & -0.35400522 & 0.306978282 \\ -0.4250167 & -0.28324973 & -0.80435350 & -0.30343304 & -0.008618095 \\ -0.2972391 & 0.29250975 & 0.08281395 & -0.05057217 & -0.903698656 \\ -0.1694616 & 0.86826924 & -0.29572248 & 0.20228869 & 0.298360187 \end{pmatrix}$$

$$A = \sigma_1 u_1 v_1^T + \sigma_2 u_2 v_2^T, \quad c = \frac{\sigma_1^2}{\sigma_1^2 + \sigma_2^2} = 0.958484$$

$$A \simeq \sigma_1 u_1 v_1^T$$

$$= \begin{pmatrix} 1.173575 & 1.994325 & 1.173575 & 0.8207501 & 0.4679250 \\ 1.769800 & 3.007526 & 1.769800 & 1.2377253 & 0.7056503 \\ 1.173575 & 1.994325 & 1.173575 & 0.8207501 & 0.4679250 \end{pmatrix}.$$

SVD in R

- `svd()`
- `la.svd()`
- A good demo
http://www.ats.ucla.edu/stat/r/pages/svd_demos.htm

SVD vs PCA

- If X is centering each column, then $X^T X$ is proportional to the covariance matrix of variables g_i . So, the right singular vectors $\{v_k\}$ are the same as the principal components.
- If instead each row of X is centered, XX^T is proportional to the covariance matrix of the variables a_j . In this case, the left singular vectors $\{u_k\}$ are the same as the principal components of $\{a_j\}$.

Data Fitting Problem

$$y = ax^2 + bx + c$$

$$\underbrace{\begin{bmatrix} x_1^2 & x_1 & 1 \\ x_2^2 & x_2 & 1 \\ \vdots & \vdots & \vdots \\ x_N^2 & x_N & 1 \end{bmatrix}}_S \underbrace{\begin{bmatrix} a \\ b \\ c \end{bmatrix}}_a = \underbrace{\begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_M \end{bmatrix}}_y$$

$$\mathbf{a} = V \cdot [\text{diag}(1/\sigma_i)] \cdot (U^\top \mathbf{y})$$

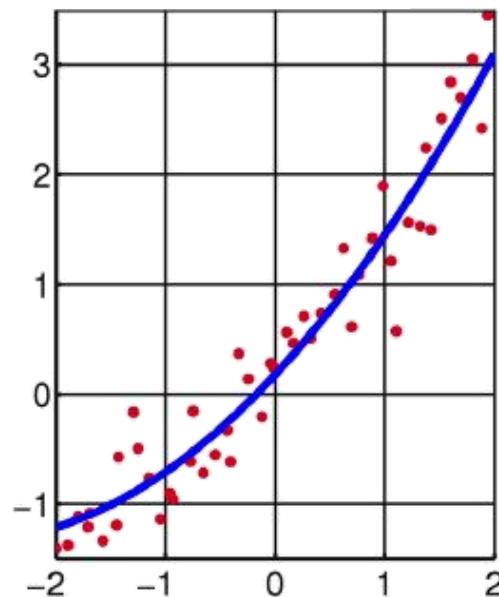


Image Processing

[U,W,V]=svd(A)

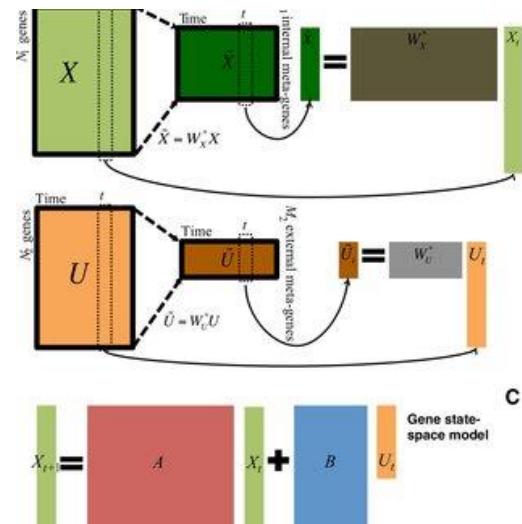
NewImg=U(:,1)*W(1,1)*V(:,1)'



Digital Signal Processing (DSP)

- SVD is used as a method for noise reduction.
- Let a matrix A represent the noisy signal:
 - compute the SVD,
 - and then discard small singular values of A .
- It can be shown that the small singular values mainly represent the noise, and thus the rank- k matrix A_k represents a filtered signal with less noise.

奇异值分解(SVD)



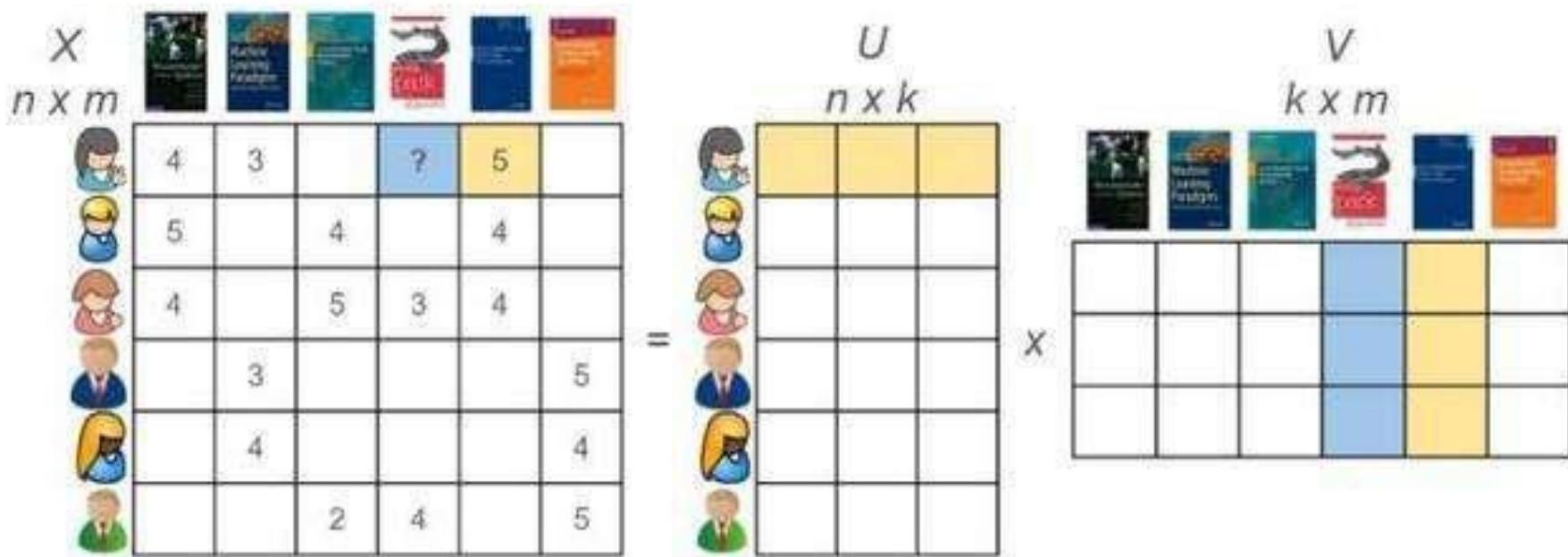
图像矩阵的秩 = 510



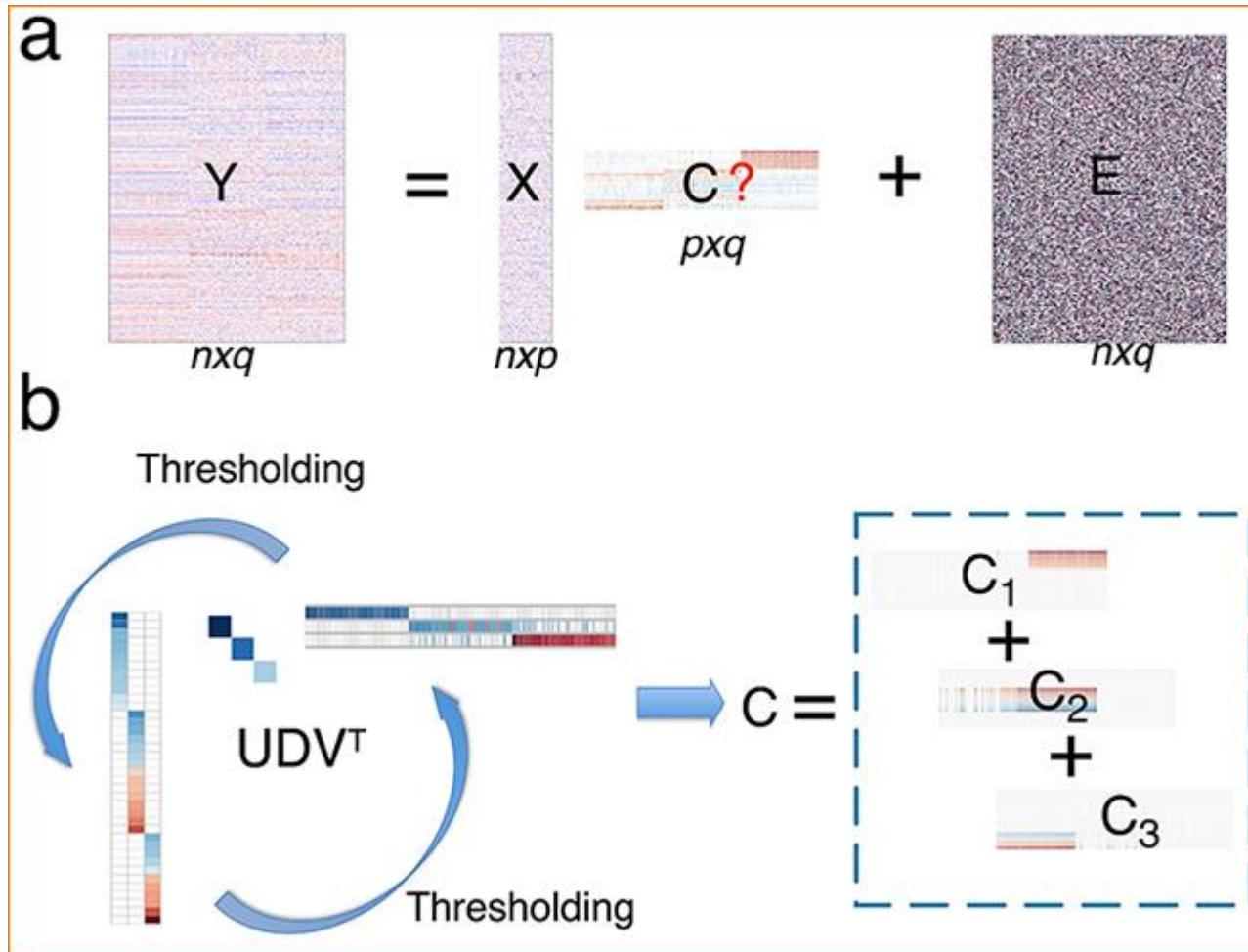
选取的奇异值的个数 = 41

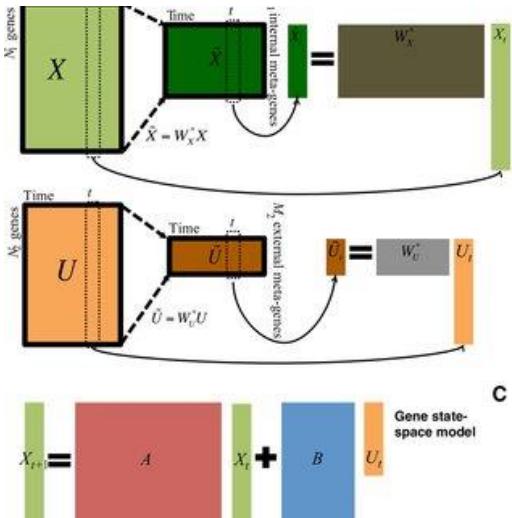


奇异值分解(SVD)

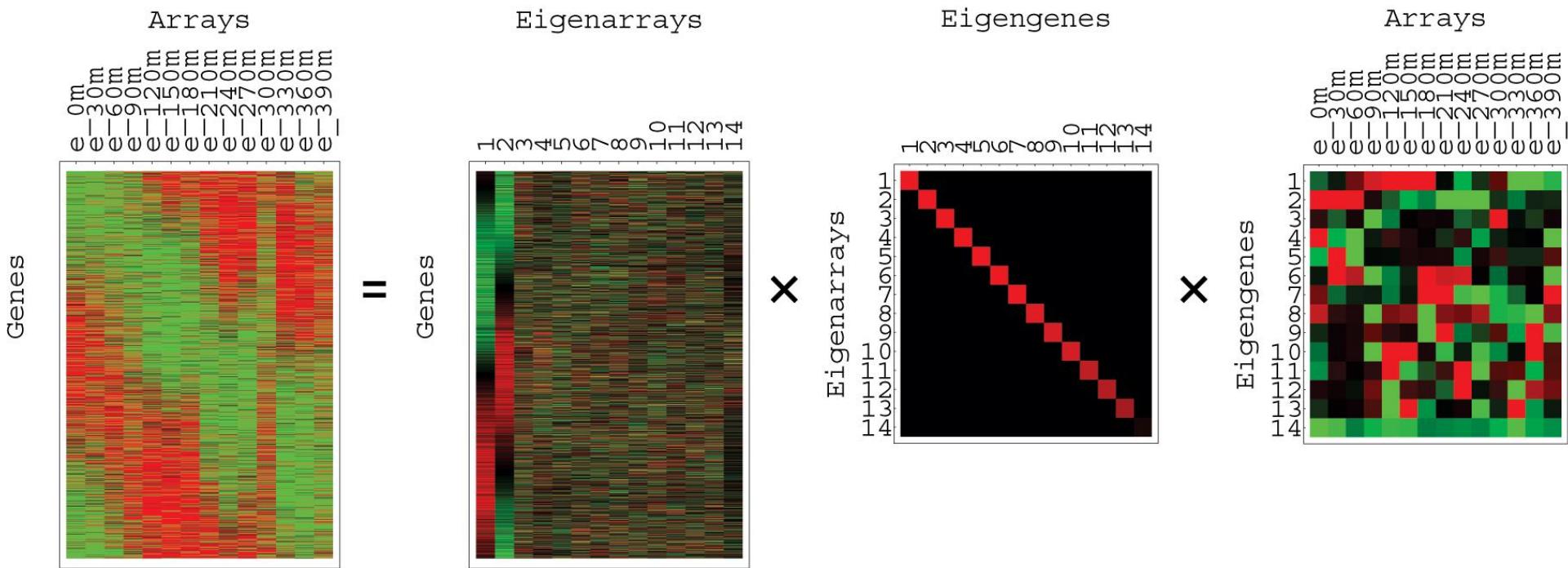
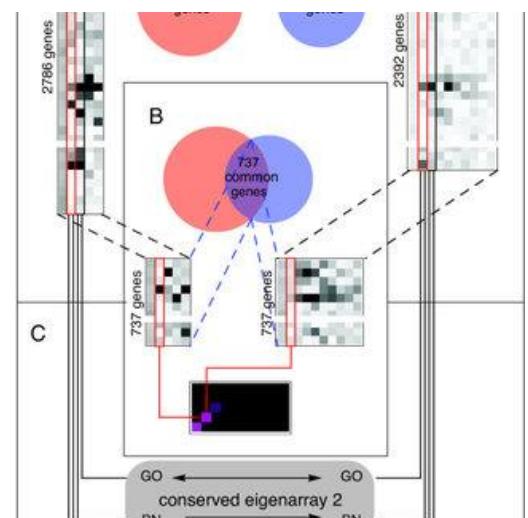


奇异值分解(SVD)





奇异值分解(SVD)



Applications

- Information retrieval
 - LSI: Latent semantic indexing
 - SVD applied to term document matrix
 - compute best rank k approximation
 - eigenvectors correspond to linguistic concepts
- Gene expression data analysis
 - SVD useful preprocessing step
 - grouping genes by transcriptional response,
grouping assays by expression profile

Latent Semantic Indexing (LSI)

by

Singular Value Decomposition

Michael W. Berry, Susan T. Dumais, Gavin W. O'Brien. Using Linear Algebra for Intelligent Information Retrieval. SIAM Review 37(4): 573-595, 1995.

Information Retrieval

- Q : “Light waves.”
- D₁: “Particle and wave models of light.”
- D₂: “Surfing on the waves under star lights.”
- D₃: “Electro-magnetic models for fotons.”

	particle	wave	model	light	surfing	star	electro-magnetic	foton	REL	MATCH
D ₁	x	x [*]	x	x [*]					R	M
D ₂		x [*]		x [*]	x	x				M
D ₃			x				x	x	R	

Motivation for LSI

- To find and fit a useful model of the relationships between terms and documents.
- To find out what terms "really" are implied by a query .
- LSI allow the user to search for concepts rather than specific words.
- LSI can retrieve documents related to a user's query even when the query and the documents do not share any common terms.

How LSI Works?

- Uses multidimensional vector space to place all documents and terms.
- Each dimension in that space corresponds to a concept existing in the collection.
- Thus underlying topics of the document is encoded in a vector.
- Common related terms in a document and query will pull document and query vector close to each other.

LSI Method

- Create a rank- k approximation to A
- $k < r_A$ or $k = r_A$,
- $A_k = U_k S_k V^T$

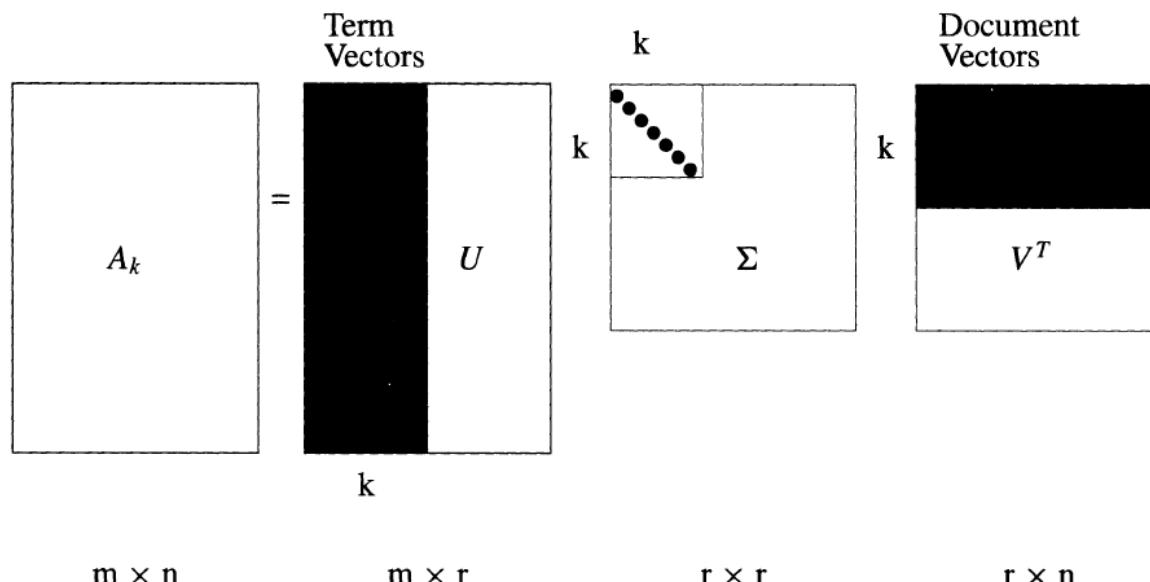


FIG. 1. Mathematical representation of the matrix A_k .

LSI Method

- Using truncated SVD, underlying *latent* structure is represented in reduced-k dimensional space.
- *Noise* in word usage is eliminated

LSI-Procedure

- Obtain term-document matrix.
- Compute the SVD.
- Truncate-SVD into reduced-k LSI space.
 - k-dimensional semantic structure
 - Similarity on reduced-space
 - Term-term
 - Term-document
 - Document-document

Query processing

- Map the query to reduced k-space

$$q' = q^T U_k S^{-1} \cdot k$$

- Retrieve documents or terms within a proximity.
 - Cosine
 - Best m

Updating

- Folding-in

$$\mathbf{d}' = \mathbf{d}^T \mathbf{U}_k \mathbf{S}^{-1} \mathbf{U}_k$$

similar to query projection

- SVD re-computation

Folding-In (Documents)

$$\begin{array}{c} A_k \\ m \times n \end{array} \quad \begin{matrix} p \\ | \\ \hline \end{matrix} = \begin{array}{c} U_k \\ m \times k \end{array} \quad \begin{array}{c} \Sigma_k \\ k \times k \end{array} \quad \begin{array}{c} V_k^T \\ k \times n \end{array} \quad \begin{matrix} p \\ | \\ \hline \end{matrix}$$

$m \times (n+p)$ $m \times k$ $k \times k$ $k \times (n+p)$

FIG. 2. Mathematical representation of folding-in p documents.

Folding-In (Terms)

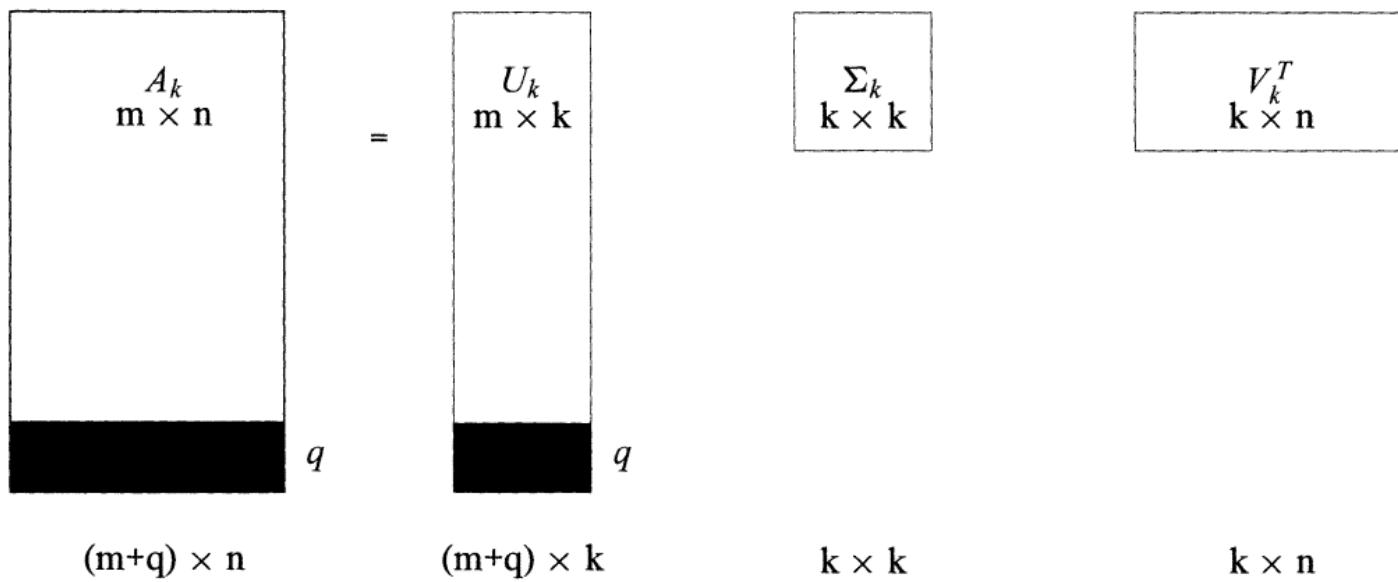


FIG. 3. *Mathematical representation of folding-in q terms.*

Example:SIAM Review Titles

Label	Titles
B1	A Course on <u>Integral Equations</u>
B2	Attractors for Semigroups and Evolution Equations
B3	Automatic Differentiation of <u>Algorithms: Theory, Implementation, and Application</u>
B4	Geometrical Aspects of <u>Partial Differential Equations</u>
B5	Ideals, Varieties, and <u>Algorithms – An Introduction to Computational Algebraic Geometry and Commutative Algebra</u>
B6	<u>Introduction to Hamiltonian Dynamical Systems</u> and the <u>N-Body Problem</u>
B7	Knapsack <u>Problems: Algorithms and Computer Implementations</u>
B8	<u>Methods of Solving Singular Systems of Ordinary Differential Equations</u>
B9	<u>Nonlinear Systems</u>
B10	<u>Ordinary Differential Equations</u>
B11	<u>Oscillation Theory for Neutral Differential Equations with Delay</u>
B12	<u>Oscillation Theory of Delay Differential Equations</u>
B13	Pseudodifferential Operators and <u>Nonlinear Partial Differential Equations</u>
B14	<u>Sinc Methods for Quadrature and Differential Equations</u>
B15	Stability of Stochastic <u>Differential Equations with Respect to Semi-Martingales</u>
B16	The Boundary <u>Integral Approach to Static and Dynamic Contact Problems</u>
B17	The Double Mellin-Barnes Type <u>Integrals and Their Applications to Convolution Theory</u>

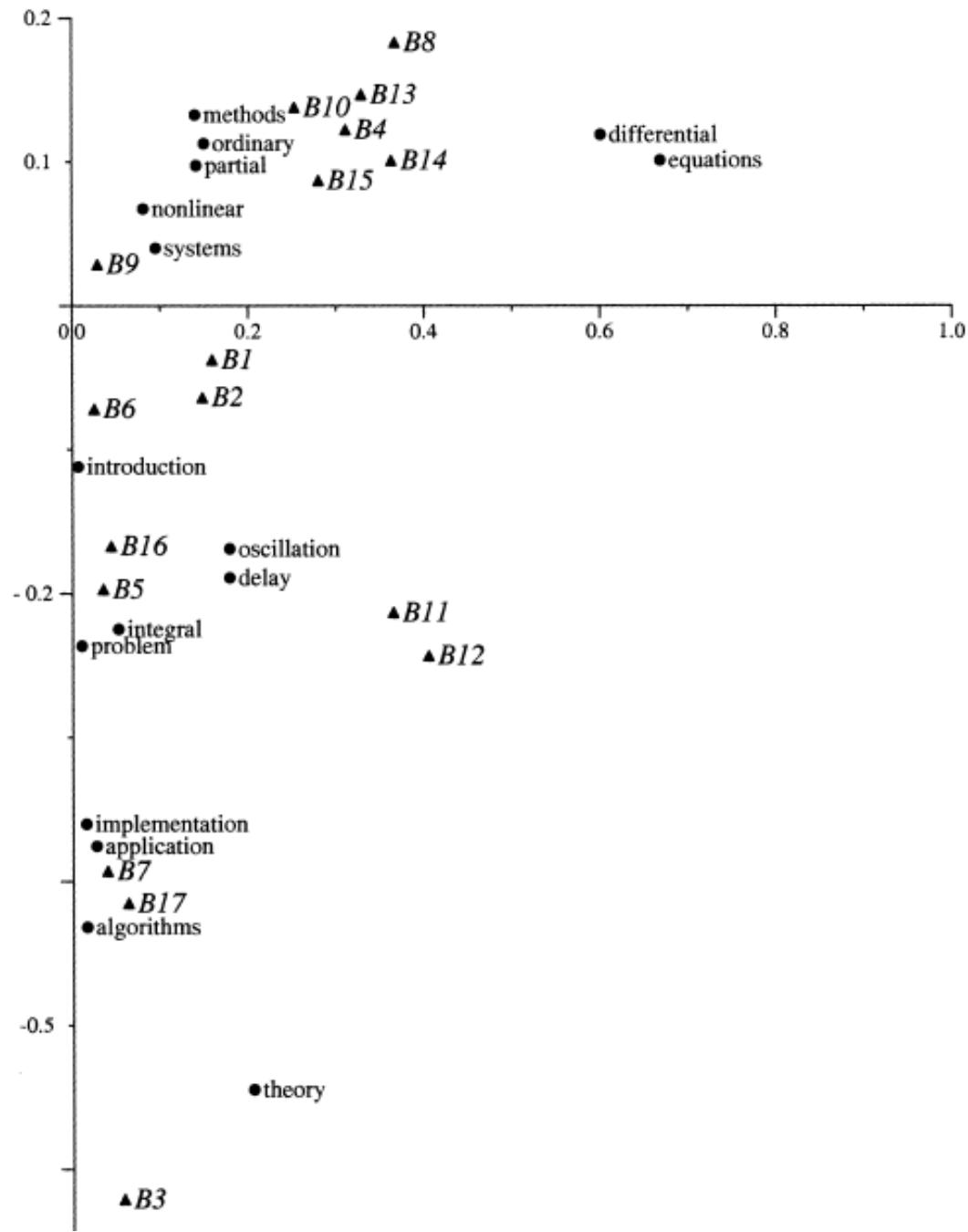
Example

Terms	Documents																
	B1	B2	B3	B4	B5	B6	B7	B8	B9	B10	B11	B12	B13	B14	B15	B16	B17
algorithms	0	0	1	0	1	0	1	0	0	0	0	0	0	0	0	0	0
application	0	0	1	0	0	0	0	0	0	0	0	0	0	0	0	0	1
delay	0	0	0	0	0	0	0	0	0	0	1	1	0	0	0	0	0
differential	0	0	0	1	0	0	0	1	0	1	1	1	1	1	1	0	0
equations	1	1	0	1	0	0	0	1	0	1	1	1	1	1	1	0	0
implementation	0	0	1	0	0	0	1	0	0	0	0	0	0	0	0	0	0
integral	1	0	0	0	0	0	0	0	0	0	0	0	0	0	0	1	1
introduction	0	0	0	0	1	1	0	0	0	0	0	0	0	0	0	0	0
methods	0	0	0	0	0	0	0	1	0	0	0	0	0	0	1	0	0
nonlinear	0	0	0	0	0	0	0	0	1	0	0	0	0	1	0	0	0
ordinary	0	0	0	0	0	0	0	1	0	1	0	0	0	0	0	0	0
oscillation	0	0	0	0	0	0	0	0	0	0	1	1	0	0	0	0	0
partial	0	0	0	1	0	0	0	0	0	0	0	0	1	0	0	0	0
problem	0	0	0	0	0	1	1	0	0	0	0	0	0	0	0	1	0
systems	0	0	0	0	0	1	0	1	1	0	0	0	0	0	0	0	0
theory	0	0	1	0	0	0	0	0	0	0	1	1	0	0	0	0	1

Mapping Method

- Compute the truncated SVD ($k=2$)
- Terms plane
 - x-coordinates : The first column of U_2 multiplied by the first singular value, for the
 - y-coordinates: second column of U_2 multiplied by the second singular value.
- Documents plane:
 - x-coordinates : The first column of V_2 scaled by the first singular value, for the
 - y-coordinates: second column of V_2 scaled by the second singular value.

Mapping Results



Query

- Query: “application and theory”

$$q^T = [0 \ 1 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 1].$$

$$\begin{pmatrix} 0.0511 & -0.3337 \end{pmatrix} = \begin{pmatrix} 0 \\ 1 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 1 \end{pmatrix}^T \begin{pmatrix} 0.0159 & -0.4317 \\ 0.0266 & -0.3756 \\ 0.1785 & -0.1692 \\ 0.6014 & 0.1187 \\ 0.6691 & 0.1209 \\ 0.0148 & -0.3603 \\ 0.0520 & -0.2248 \\ 0.0066 & -0.1120 \\ 0.1503 & 0.1127 \\ 0.0813 & 0.0672 \\ 0.1503 & 0.1127 \\ 0.1785 & -0.1692 \\ 0.1415 & 0.0974 \\ 0.0105 & -0.2363 \\ 0.0952 & 0.0399 \\ 0.2051 & -0.5448 \end{pmatrix} \begin{pmatrix} 4.5314 & 0 \\ 0 & 2.7582 \end{pmatrix}^{-1}$$

FIG. 5. Derived coordinates for the query of application theory.

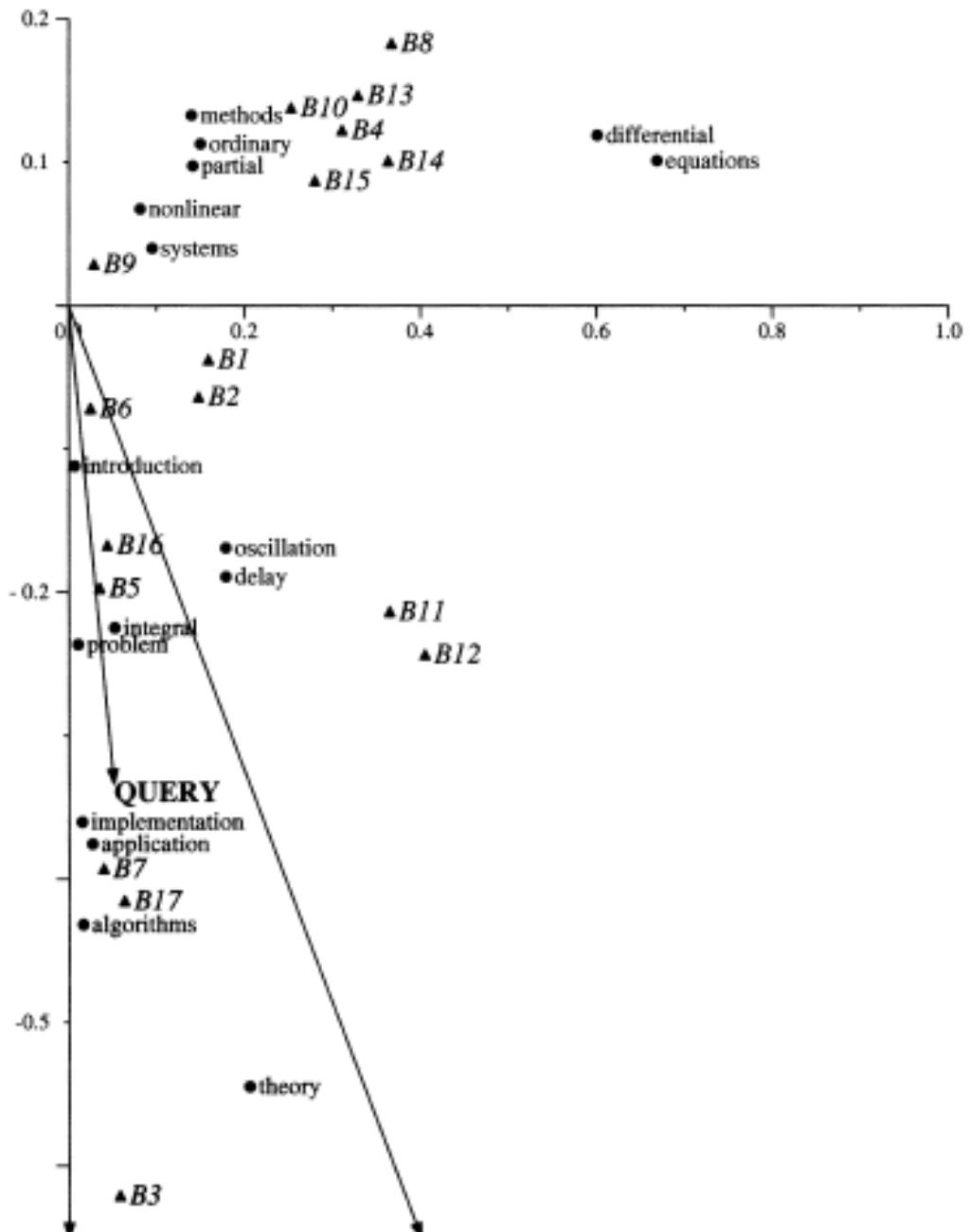


Fig. 6. A two-dimensional plot of terms and documents along with the query application theory.

Comparison with Lexical Matching

- Query of “Application and theory”; k=2
- Using a cosine threshold of 0.9 returns 6 books: B_3 , B_5 , B_6 , B_7 , B_{16} , and B_{17} .
- Using a cosine threshold 0.55 also returns B_{11} and B_{12} (somewhat related)
- But lexical matching only returns 4 books: B_3 , B_{11} , B_{12} , B_{17}

Gene Expression Data Analysis Using SVD

Wall, Michael E., Andreas Rechtsteiner, Luis M. Rocha."Singular value decomposition and principal component analysis". in *A Practical Approach to Microarray Data Analysis*. D.P. Berrar, W. Dubitzky, M. Granzow, eds. pp. 91-109, Kluwer: Norwell, MA (2003). LANL LA-UR-02-4001.

$$X = USV^T$$

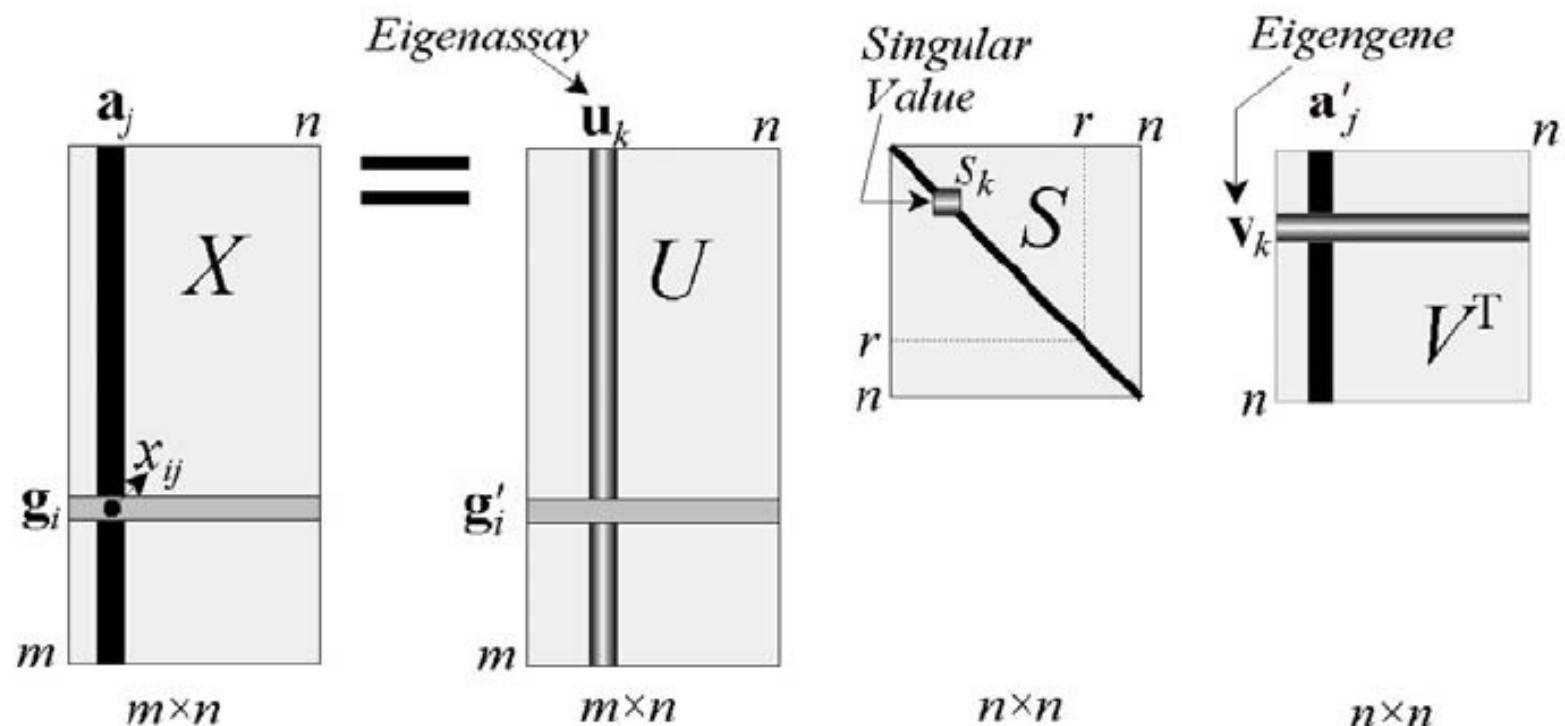


Figure 5.1. Graphical depiction of SVD of a matrix X , annotated with notations adopted in this chapter.

Left Singular Vector and Right Singular Vector

- The right singular vector span the space of the gene transcriptional responses $\{g_i\}$;
- The left singular vectors span the space of the assay expression profiles $\{a_j\}$

Eigen-vectors

- Genes: linear combination of eigen-genes $\{v_k\}$

$$\vec{g}_i = \sum_{k=1}^r u_{ik} s_k \vec{v}_k, \quad i = 1, \dots, m.$$

- Assays: linear combination of eigen-assays $\{u_k\}$

$$\vec{a}_j = \sum_{k=1}^r v_{jk} s_k \vec{u}_k$$

Dimension Reduction

- Using the idea of truncated-SVD, we can represent expression profiles as,

$$\vec{g'}_i = \sum_{k=1}^d u_{ik} s_k \vec{v}_k, \quad i = 1, \dots, m.$$

where $d < r = \text{rank}(X)$.

- In another word, we treat the last several singular values only contributing to the noise.

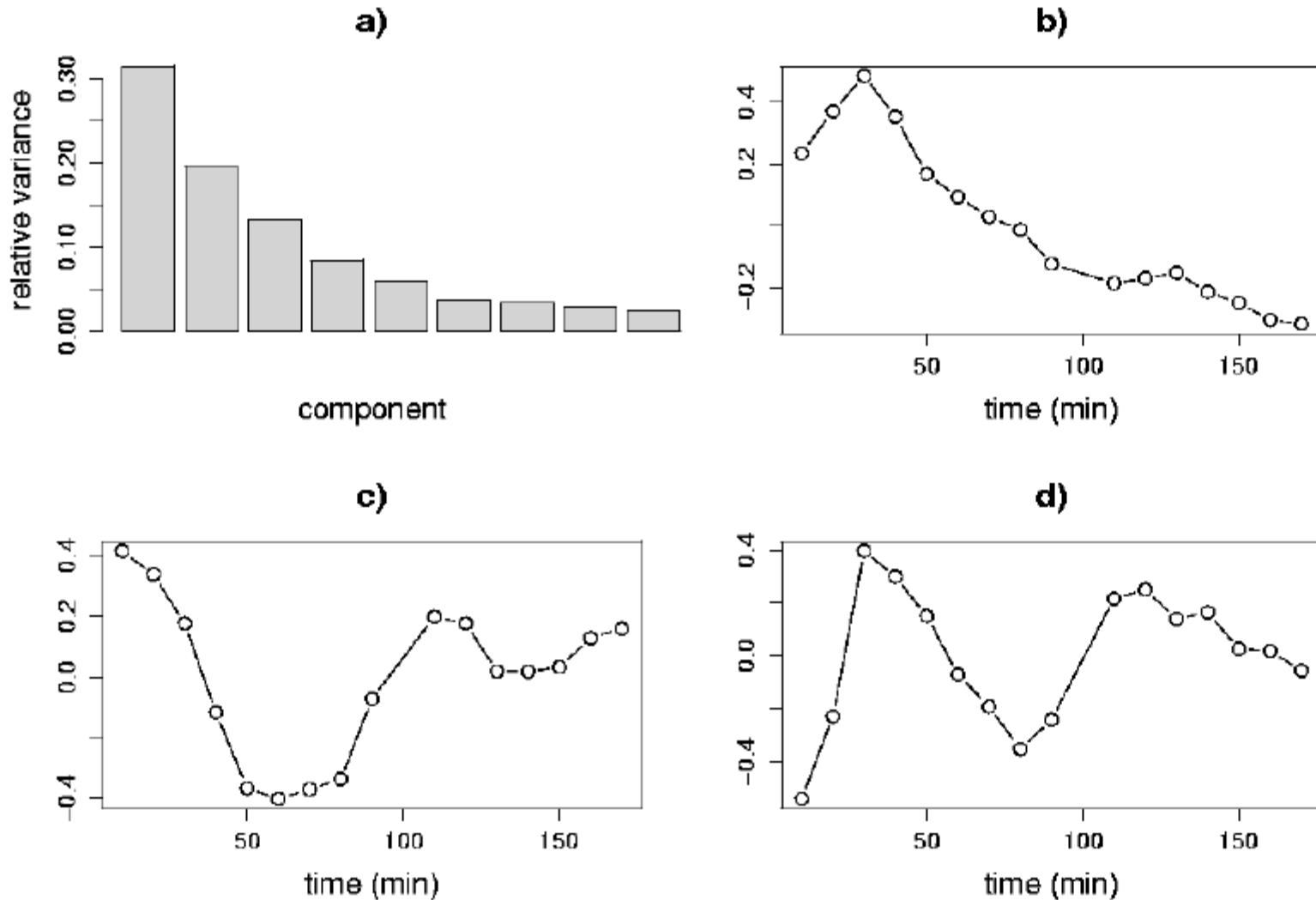


Figure 5.2. Visualization of the SVD of cell cycle data. Plots of relative variance (a); and the first (b), second (c) and third (d) eigengenes are shown. The methods of visualization employed in each panel are described in section 2.1. These data inspired our choice of the sine and exponential patterns for the synthetic data of section 2.1.

Synthetic Time Series Data

- 2000 x 14
- 1600 noisy genes
- 200 genes with noisy sine patterns
 $a \sin\left(\frac{2\pi t}{140}z\right), \quad a \sim U(1.5, 3)$
- 200 genes with exponential patterns

$$b \exp\left(-\frac{t}{100}\right), \quad b \sim U(4, 8)$$

Synthetic Data

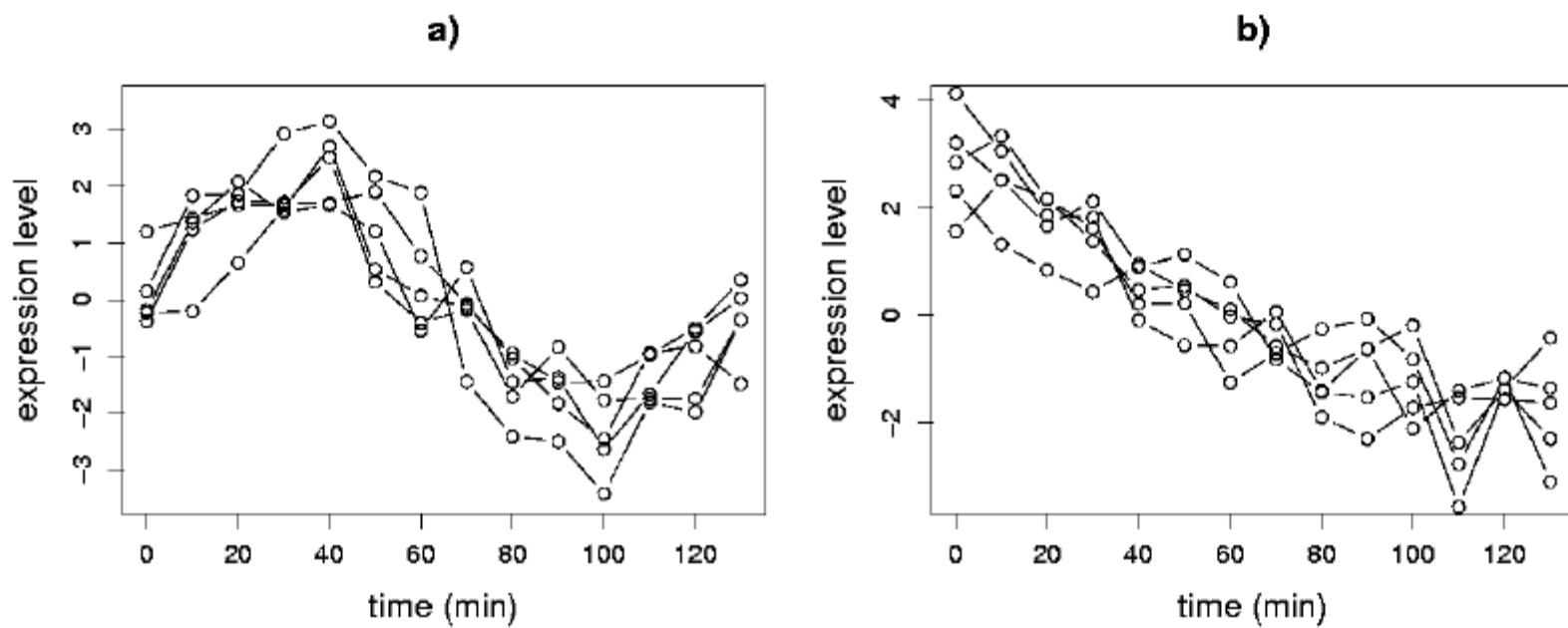


Figure 5.3. Gene transcriptional responses from the synthetic data set. Overlays of a) five noisy sine wave genes and b) five noisy exponential genes.

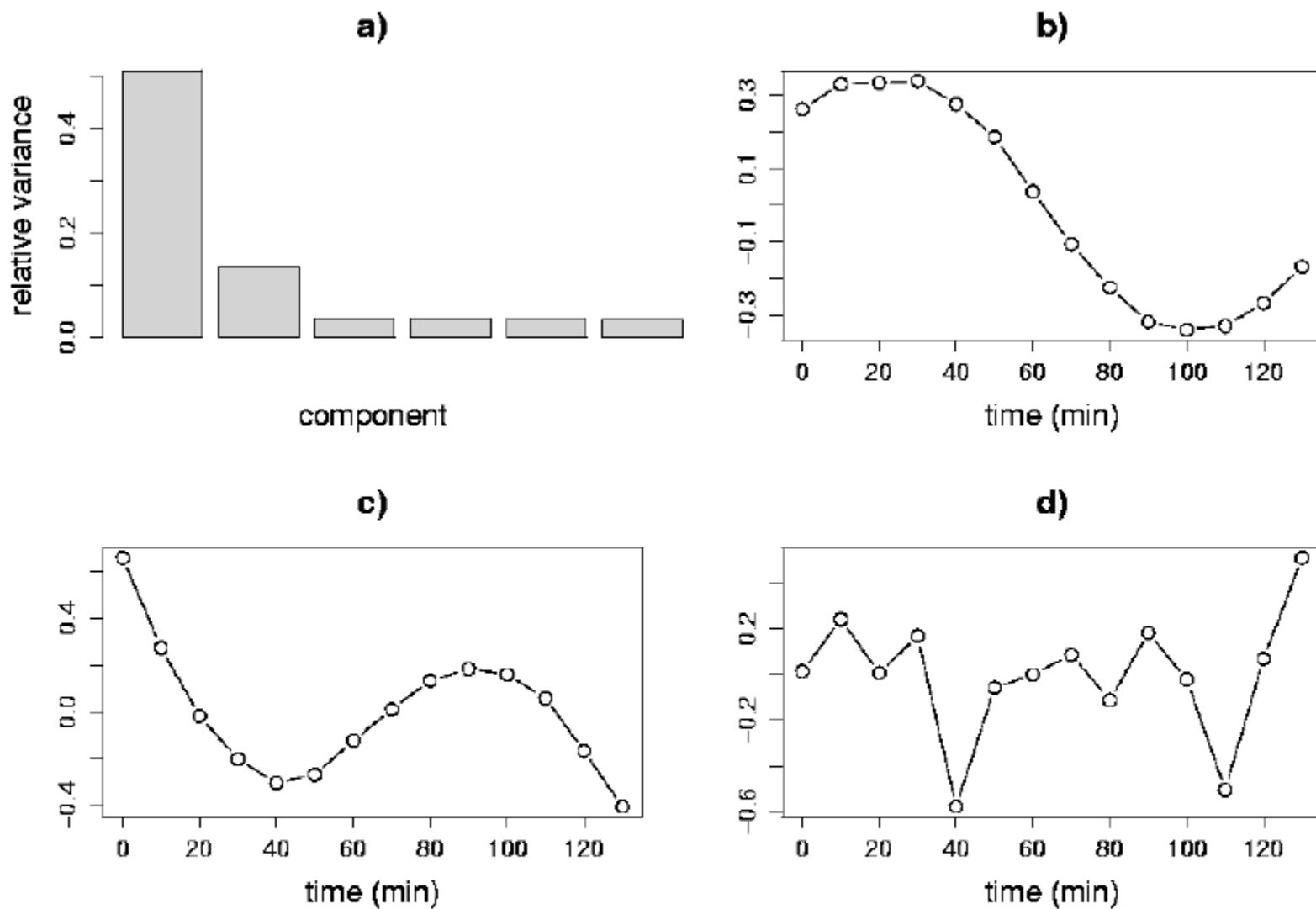


Figure 5.4. Visualization of the SVD of the synthetic data matrix. a) Singular value spectrum in a relative variance plot. The first two singular values account for 64% of the variance. The first (b), second (c), and third (d) eigengenes are plotted vs. time (assays) in the remaining panels. The third eigengene lacks the obvious cyclic structure of the first and second.

Scatter Plots

- Projection of data into SVD subspaces and visualization with scatter plots can reveal structures in the data that may be used for classification.
- Two gene “coordinates” for scatter plots
 - Projection
 - Correlation

Scatter Plots

- Projection of g_i on eigen-gene v_k

$$q_{ik} = \vec{g}_i \cdot \vec{v}_k, XV = US \Rightarrow q_{ik} = (US)_{ik}$$

- Correlation between g_i and eigen-gene v_k

$$r_{ik} = \text{PCC}(\vec{g}_i, \vec{v}_k)$$

$$\text{PCC}(X, Y) = \frac{\sum_i (X_i - \bar{X})(Y_i - \bar{Y})}{\sqrt{\sum_i (X_i - \bar{X})^2 \sum_i (Y_i - \bar{Y})^2}}$$

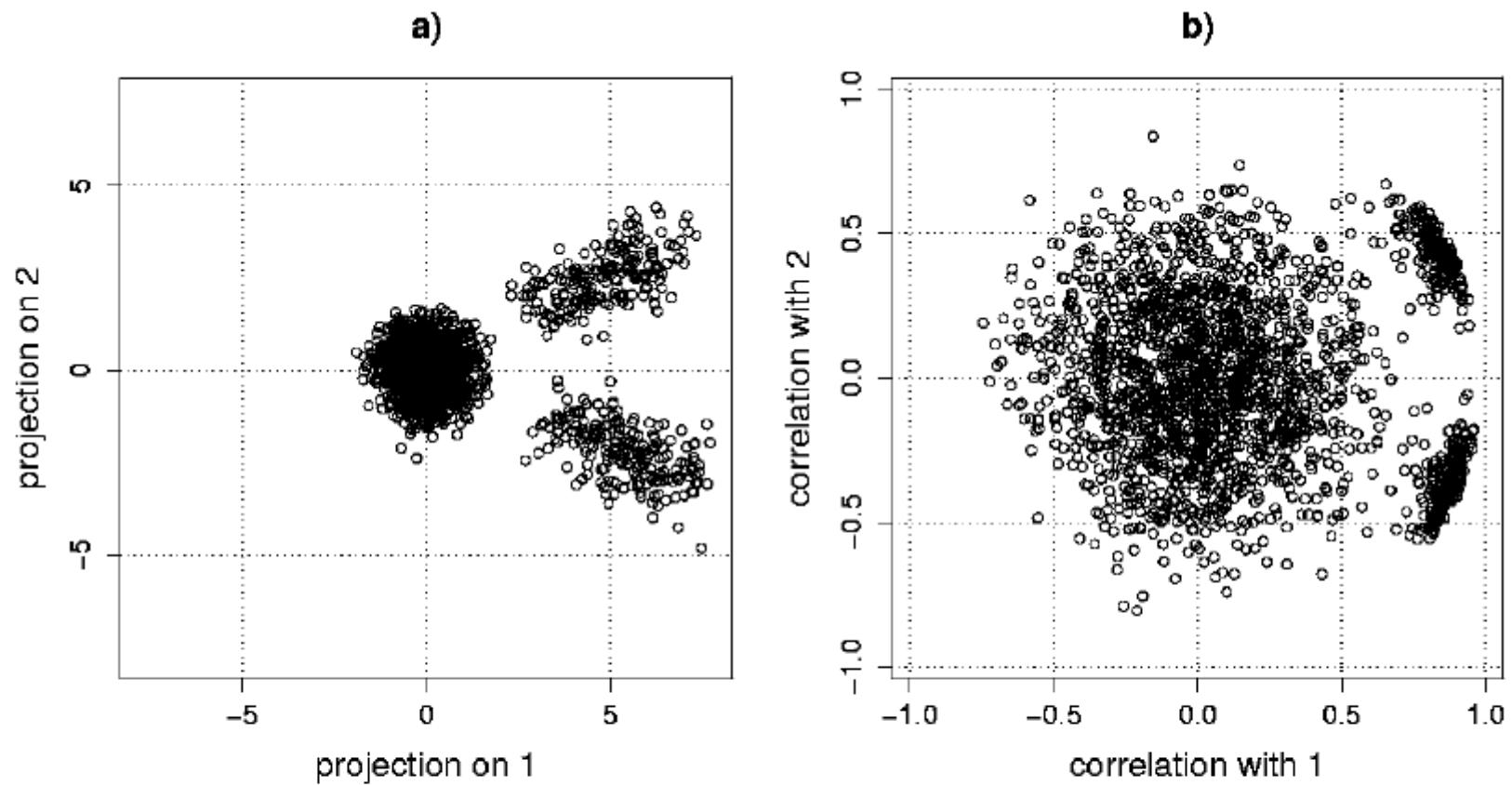


Figure 5.5. SVD scatter plots. Genes from our synthetic example data set are displayed in a) a projection scatter plot; and b) a correlation scatter plot. The bottom right cluster corresponds to sine wave genes, and the top right cluster corresponds to exponential decay genes. The cluster of genes around the origin corresponds to the noise-only genes.

More Applications

- Netflix prize, big data, SVD and R

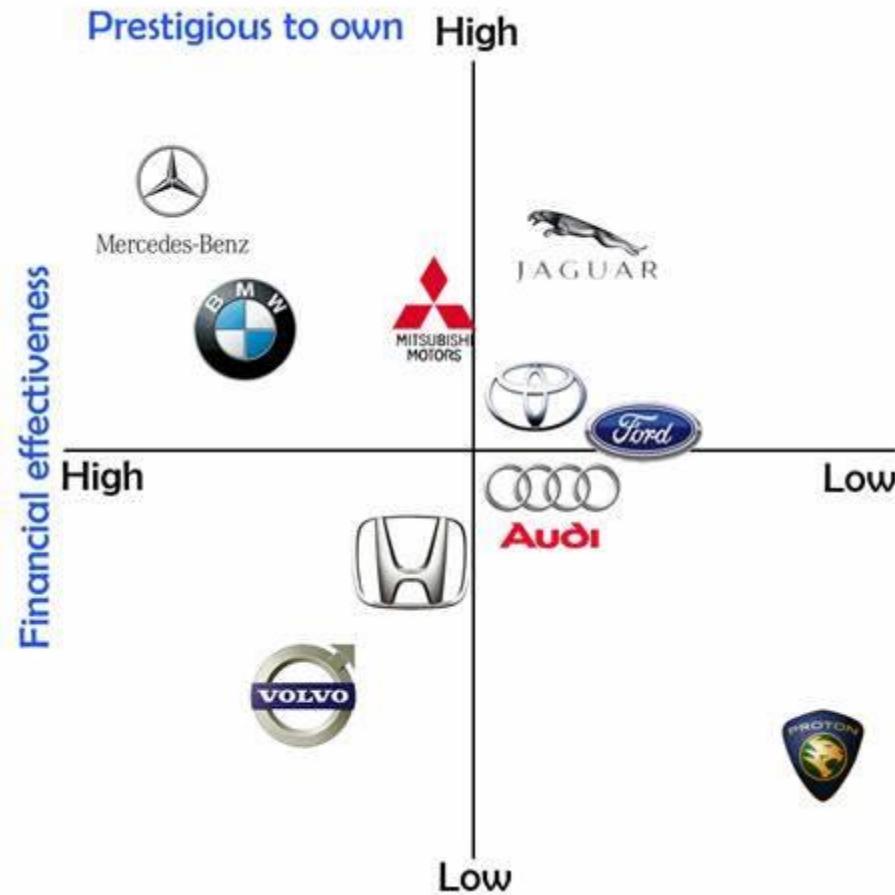
[http://blog.revolutionanalytics.com/2011/05/
the-netflix-prize-big-data-svd-and-r.html](http://blog.revolutionanalytics.com/2011/05/the-netflix-prize-big-data-svd-and-r.html)

Multi-dimensional Scaling (MDS)

部分Slide来自

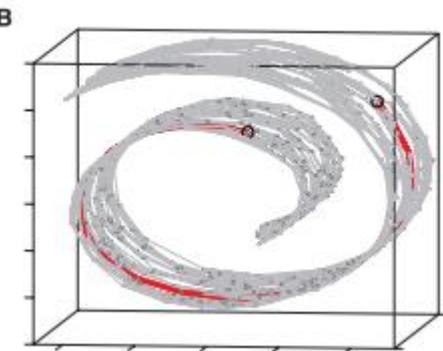
[www.cs.haifa.ac.il/~rita/uml_course/
lectures/PCA_MDS.pdf](http://www.cs.haifa.ac.il/~rita/uml_course/lectures/PCA_MDS.pdf)

MDS (多维尺度分析)



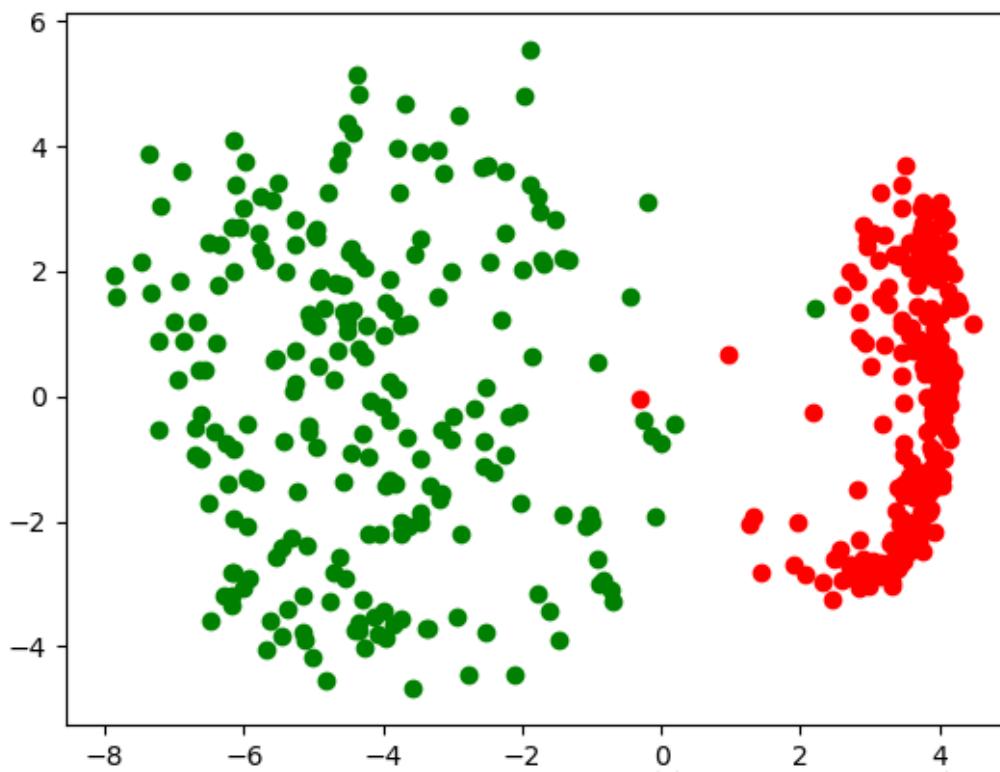
MDS (多维尺度分析)

- MDS attempts to preserve pairwise distances.
- Attempts to construct a configuration of n points in Euclidian space by using the information about the **distances between the n patterns (Not necessary n points in the high dimensional space)** .



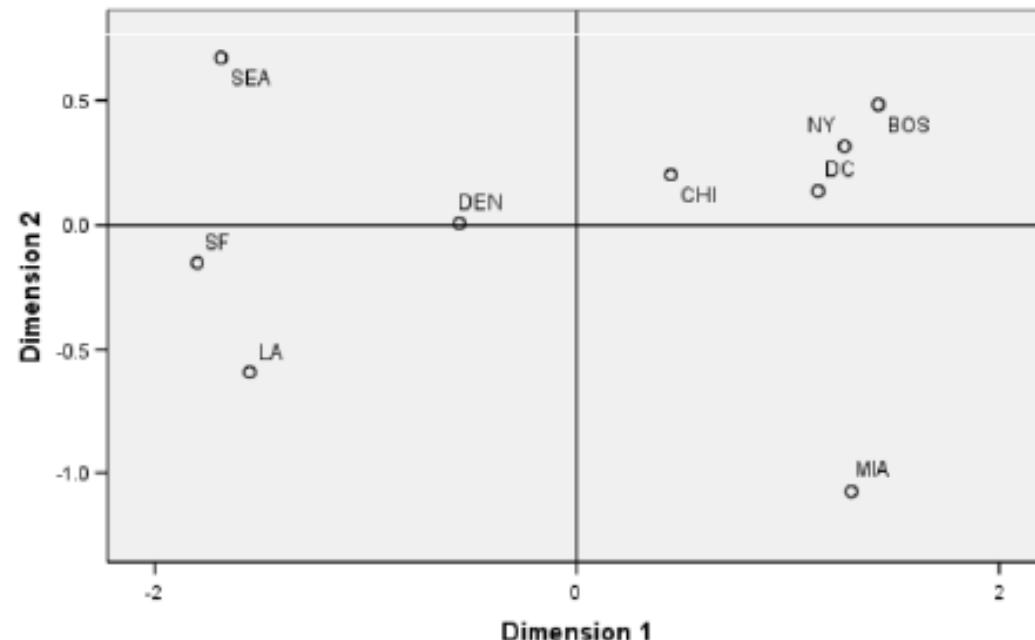
MDS (多维尺度分析)

其主要思想是构造低维空间的内积矩阵，使得该内积矩阵中所表达的任意两点之间的距离与高维空间的相应两点距离相等，然后通过对该内积矩阵进行正交特征值分解，析出两个矩阵相乘(即矩阵与矩阵的转置进行相乘)的形式，获得最终的变换矩阵。



Example: Distances between US Cities

	BOS	CHI	DC	DEN	LA	MIA	NY	SEA	SF
BOS	0	963	429	1,949	2,979	1,504	206	2,976	3,095
CHI	963	0	671	996	2,054	1,329	802	2,013	2,142
DC	429	671	0	1,616	2,631	1,075	233	2,684	2,799
DEN	1,949	996	1,616	0	1,059	2,037	1,771	1,307	1,235
LA	2,979	2,054	2,631	1,059	0	2,687	2,786	1,131	379
MIA	1,504	1,329	1,075	2,037	2,687	0	1,308	3,273	3,053
NY	206	802	233	1,771	2,786	1,308	0	2,815	2,934
SEA	2,976	2,013	2,684	1,307	1,131	3,273	2,815	0	808
SF	3,095	2,142	2,799	1,235	379	3,053	2,934	808	0



Classical MDS (I)

- Given n points $X_1, \dots, X_n \in R^p$, the pairwise dissimilarities

$$X_i = \begin{pmatrix} X_{i1} \\ X_{i2} \\ \dots \\ X_{ip} \end{pmatrix} \quad X = \begin{pmatrix} X_{11} & X_{21} & \dots & X_{n1} \\ X_{12} & X_{22} & \dots & X_{n2} \\ \dots & \dots & \dots & \dots \\ X_{1p} & X_{2p} & \dots & X_{np} \end{pmatrix}_{p \times n}$$

$$\begin{aligned}\delta_{ij}^2 &= (X_i - X_j)^T (X_i - X_j) \\ &= X_i^T X_i + X_j^T X_j - 2X_i^T X_j \\ &= l_i^2 + l_j^2 - 2X_i^T X_j\end{aligned}$$

- Where l_i is the length of i th point.

Classical MDS (II)

- Let

$$D = (\delta_{ij} j^2), K = X^T X$$

$$k = \text{diag}\{l_1^2, l_2^2, \dots, l_n^2\}$$

- Then

$$D = k e e^T + e e^T k - 2K$$

$$e = \begin{pmatrix} 1 \\ 1 \\ \dots \\ 1 \end{pmatrix}_{n \times 1}, \quad e e^T = \begin{pmatrix} 1 & 1 & \dots & 1 \\ 1 & 1 & \dots & 1 \\ \dots & \dots & \dots & \dots \\ 1 & 1 & \dots & 1 \end{pmatrix}_{n \times n}$$

Classical MDS (III)

- Centralization

$$\overline{X} = \frac{1}{n} \sum_{i=1}^n X_i = \frac{1}{n} e^T X$$

$$\tilde{X} = X - e\overline{X} = X - \frac{1}{n} Xee^T$$

- Let

$$H = I_n - \frac{1}{n} ee^T$$

$$\tilde{X} = XH$$

Classical MDS (IV)

- We have

$$H = \frac{1}{n} \begin{pmatrix} n-1 & -1 & \cdots & -1 \\ -1 & n-1 & \cdots & -1 \\ \cdots & \cdots & \cdots & \cdots \\ -1 & -1 & \cdots & n-1 \end{pmatrix}$$

$$\begin{aligned} H^2 &= (I_n - \frac{1}{n}ee^T)(I_n - \frac{1}{n}ee^T) & H &= H^T \\ &= I_n - \frac{2}{n}ee^T + \frac{1}{n^2}e(e^Te)e^T \\ &= H \end{aligned}$$

Classical MDS (V)

- Let

$$\tilde{K} = \tilde{X}^T \tilde{X}$$

$$\tilde{K} = (XH)^T(XH) = H^T X^T X H = HKH$$

- Let

$$B = -\frac{1}{2} HDH$$

Classical MDS (VI)

- Then

$$\begin{aligned}B &= -\frac{1}{2}HDH \\&= -\frac{1}{2}H(kee^T + ee^T k - 2K)H^T \\&= HKH^T - \frac{1}{2}Hk ee^T H^T - \frac{1}{2}He e^T k H^T \\&= HKH^T = \tilde{K}\end{aligned}$$

- Here

$$He = (0, 0, \dots, 0)^T$$

Classical MDS (VII)

- Finally, we have

$$B_{n \times n} = \tilde{K}_{n \times n} = \tilde{X}^T \tilde{X} \geq 0$$

- Then we can have a SVD of B ($p < n$)

$$B = V \Lambda V^T$$

$$\Lambda = \text{diag}\{\lambda_1, \lambda_2, \dots, \lambda_p, 0, \dots, 0\}$$

$$\lambda_1 \geq \lambda_2 \geq \dots \lambda_p > 0$$

where V is a matrix whose columns are the eigenvectors of B

Classical MDS (VIII)

- With p-eignvectors, we can reconstruct B

$$(\hat{X}_p)_{p \times n} = \text{diag}\{\sqrt{\lambda_1}, \dots, \sqrt{\lambda_p}\}(v_1, \dots, v_p)^T$$

$$B = (\hat{X}_p)^T \hat{X}_p$$

- With k-eignvectors, we can reduce the dimension to k-dimensional space

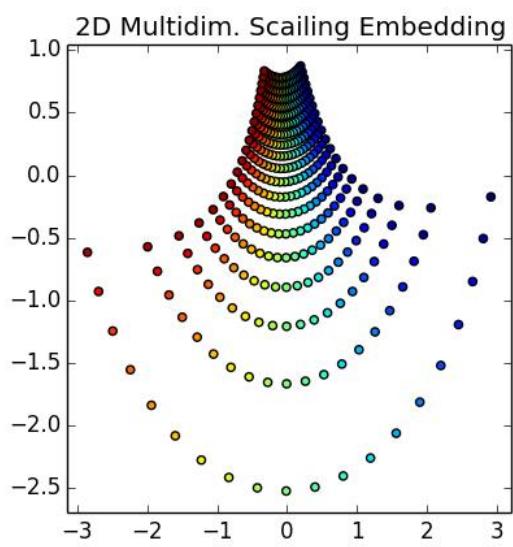
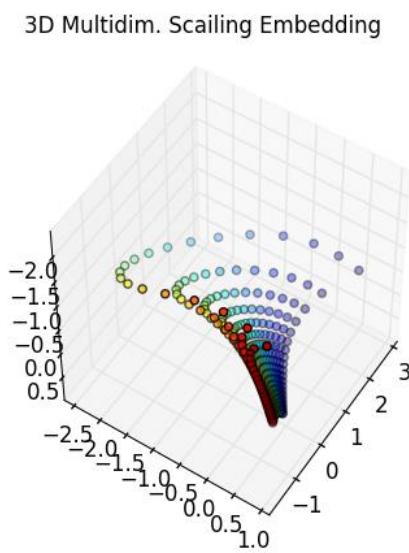
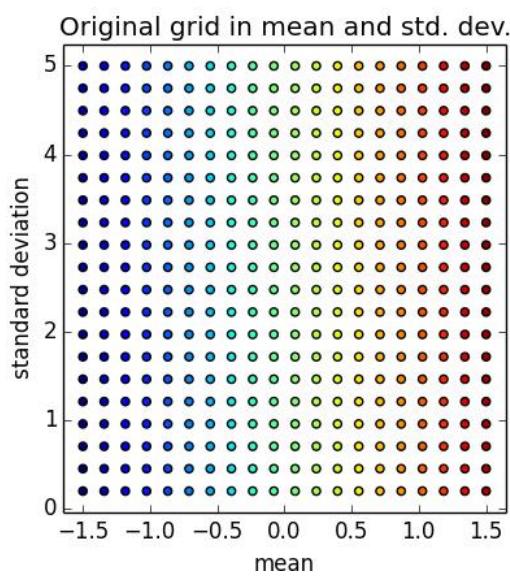
$$Y_{k \times n} = \text{diag}\{\sqrt{\lambda_1}, \dots, \sqrt{\lambda_k}\}(v_1, \dots, v_k)^T$$

Classical MDS (IX)

- Optimality property

Theorem 1. Let X denote a configuration of points in \mathbb{R}^p , with interpoint distances $\delta_{ij}^2 = (\mathbf{x}_i - \mathbf{x}_j)^T(\mathbf{x}_i - \mathbf{x}_j)$. Let L be a $p \times p$ rotation matrix and set $L = (L_1, L_2)$, where L_1 is $p \times k$ for $k < p$. Let $\hat{X} = XL_1$, the projection of X onto a k -dimensional subspace of \mathbb{R}^p , and let $\hat{d}_{ij}^2 = (\hat{\mathbf{x}}_i - \hat{\mathbf{x}}_j)^T(\hat{\mathbf{x}}_i - \hat{\mathbf{x}}_j)$. Amongst all projections $\hat{X} = XL_1$, the quantity $\phi = \sum_{i,j} (\delta_{ij}^2 - \hat{d}_{ij}^2)$ is minimized when X is projected onto its principal coordinates in k dimensions. For all i, j we have $\hat{d}_{ij} \leq \delta_{ij}$. The value of ϕ for the principal coordinate projection is $\phi = 2n(\lambda_{k+1} + \dots + \lambda_p)$.

MDS (多维尺度分析)



Relation to PCA

	PCA	MDS
Spectral Decomposition	Covariance matrix ($D \times D$)	Gram matrix ($n \times n$)
Eigenvalues	Matrices share nonzero eigenvalues up to constant factor	
Results		Same
Computation	$O((n+d)D^2)$	$O((D+d)n^2)$

Non-metric MDS

- Transform pairwise distances: $\delta_{ij} \rightarrow g(\delta_{ij})$
 - Transformation: nonlinear, but monotonic.
 - Preserves rank order of distances.
- Find vectors y_i such that $\|y_i - y_j\| \approx g(\delta_{ij})$

$$Cost = \min_y \sum_{ij} (g(\delta_{ij}) - \|y_i - y_j\|)^2$$

Non-metric MDS

- Possible objective function:

$$Cost = \sum_{i,j} \left(\frac{\|\mathbf{x}_i - \mathbf{x}_j\| - \|\mathbf{y}_i - \mathbf{y}_j\|}{\|\mathbf{x}_i - \mathbf{x}_j\|} \right)^2$$

Non-metric MDS

- Strengths
 - Relaxes distance constraints.
 - Yields nonlinear embeddings.
- Weaknesses
 - Highly nonlinear, iterative optimization with local minima.
 - Unclear how to choose distance transformation.

MDS Implementation

- MDS in R
 - isoMDS(MASS)
 - Kruskal's Non-metric Multidimensional Scaling
 - cmdscale(stats)
 - Classical (Metric) Multidimensional Scaling
 - sammon(MASS)
 - Sammon's Non-Linear Mapping
- Various software and resources about MDS

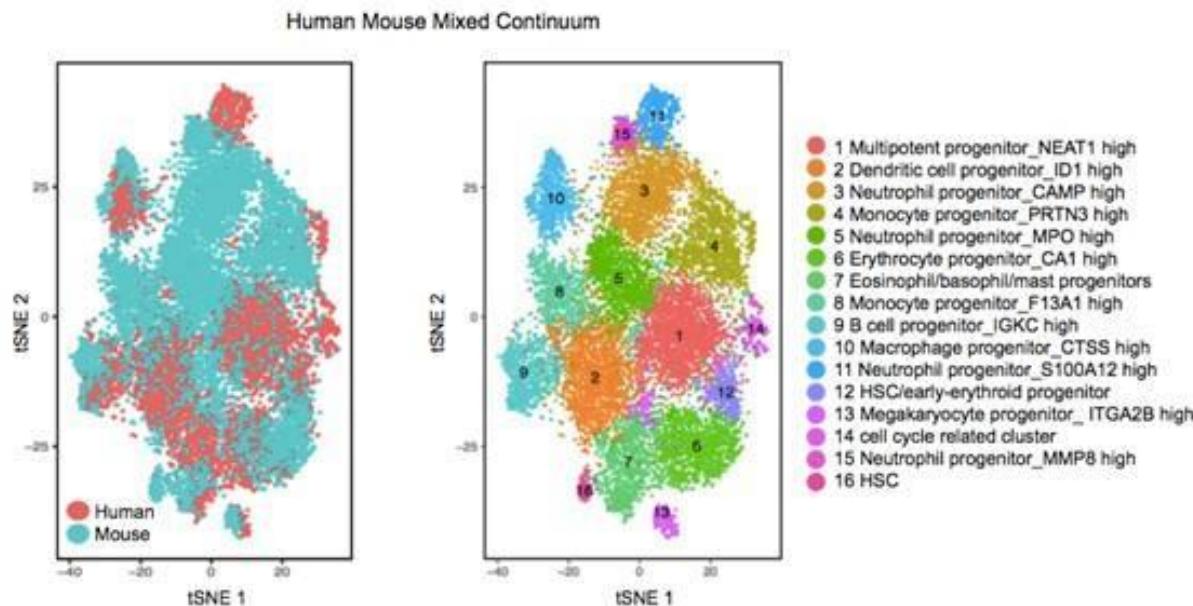
<http://www.granular.com/MDS/>

Other Methods

- Nonlinear dimension reduction
 - Kernel PCA
 - Locally linear embedding (LLE)
 - Isomap
 - t-SNE
 - LASSO
- A good tutorial: Ali Ghodsi. Dimensionality Reduction, A Short Tutorial.
http://www.math.uwaterloo.ca/~aghodsib/courses/s09stat946/readings/tutorial_stat890.pdf
- More materials can be found at his webpage
<http://www.math.uwaterloo.ca/~aghodsib/courses/s09stat946/>

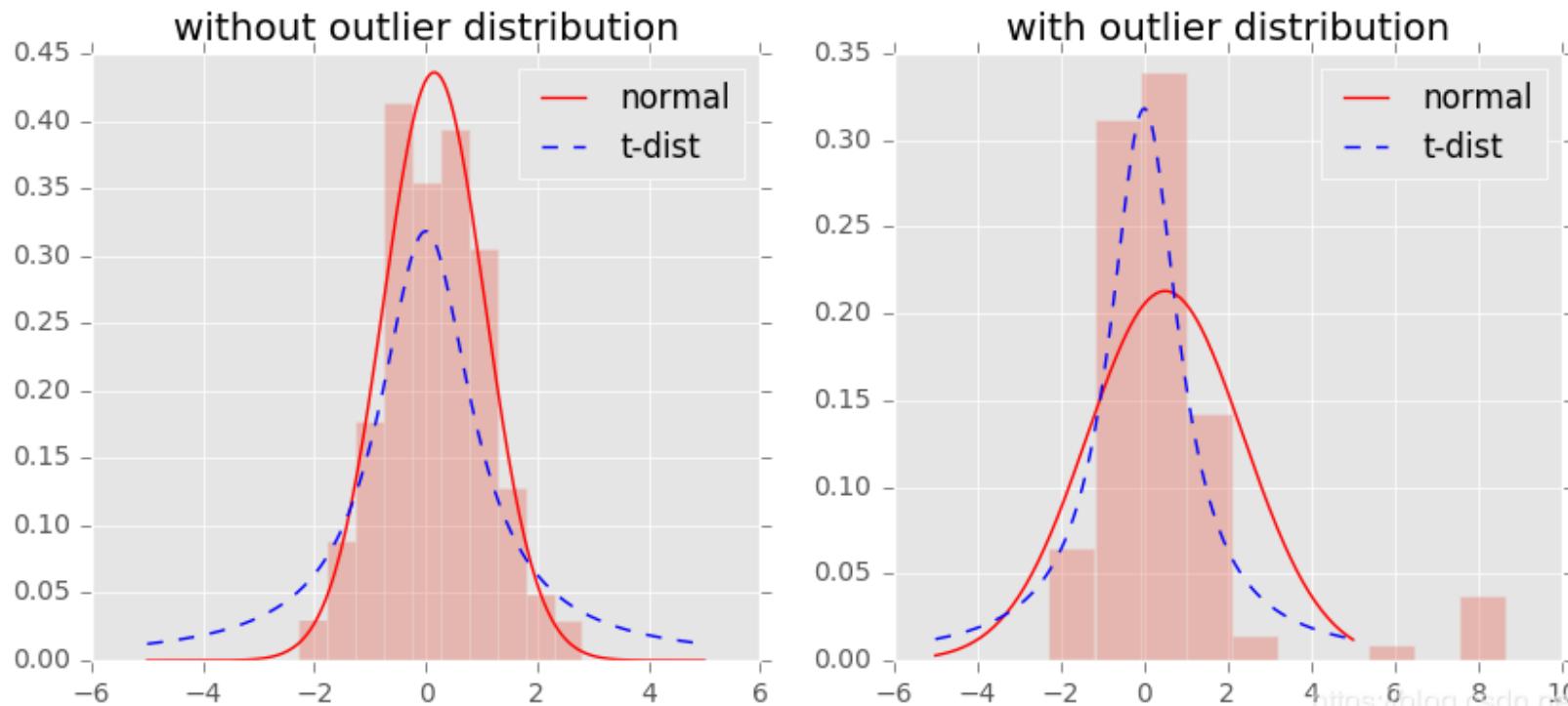
t-SNE(t-distributed stochastic neighbor embedding)

- SNE是通过仿射(affinitie)变换将数据点映射到概率分布上，主要包括两个步骤：
 - SNE构建一个高维对象之间的概率分布，使得相似的对象有更高的概率被选择，而不相似的对象有较低的概率被选择。
 - SNE在低维空间里在构建这些点的概率分布，使得这两个概率分布之间尽可能的相似。



t-SNE(t-distributed stochastic neighbor embedding)

- 对称SNE实际上在高维度下另外一种减轻“拥挤问题”的方法：在高维空间下，在高维空间下我们使用高斯分布将距离转换为概率分布，在低维空间下，我们使用更加偏重长尾分布的方式来将距离转换为概率分布，使得高维度下中低等的距离在映射后能够有一个较大的距离。

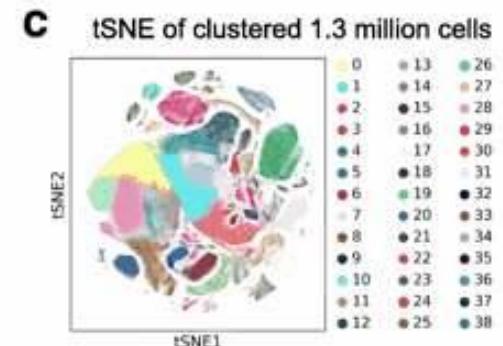
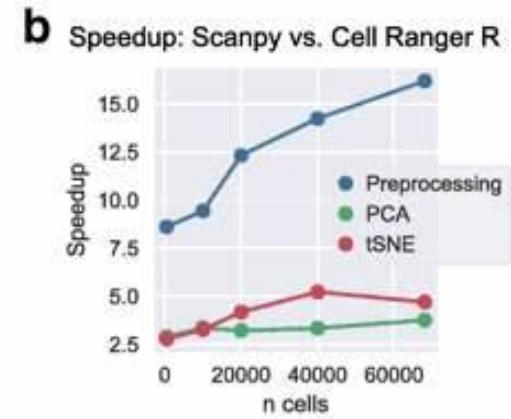
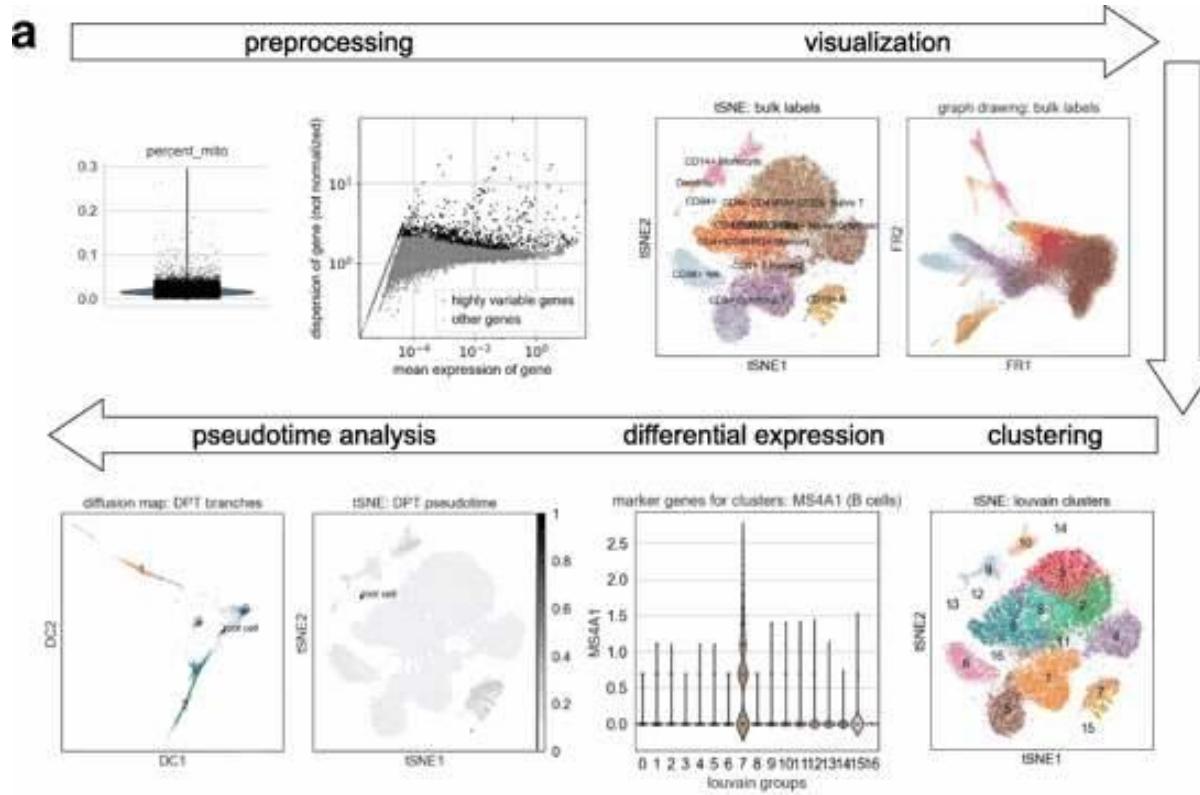


t-SNE(t-distributed stochastic neighbor embedding)

9

t-SNE(t-distributed stochastic neighbor embedding)

Single-cell applications



t-SNE(t-distributed stochastic neighbor embedding)

- 主要不足有四个：
 - 主要用于可视化，很难用于其他目的。比如测试集合降维，因为他没有显式的预估部分，不能在测试集合直接降维；比如降维到10维，因为t分布偏重长尾，1个自由度的t分布很难保存好局部特征，可能需要设置成更高的自由度
 - t-SNE倾向于保存局部特征，对于本征维数(intrinsic dimensionality)本身就很高的数据集，是不可能完整的映射到2-3维的空间
 - t-SNE没有唯一最优解，且没有预估部分。如果想要做预估，可以考虑降维之后，再构建一个回归方程之类的模型去做。但是要注意，t-sne中距离本身是没有意义，都是概率分布问题
 - 训练太慢。

LASSO

- Lasso是一种数据降维方法
- 不仅适用于线性情况，也适用于非线性情况
- Lasso是基于惩罚方法对样本数据进行变量选择
 - 通过对原本的系数进行压缩，将原本很小的系数直接压缩至0
 - 将这部分系数所对应的变量视为非显著性变量
 - 将不显著的变量直接舍弃

Lasso Model

(least absolute shrinkage and selection operator)

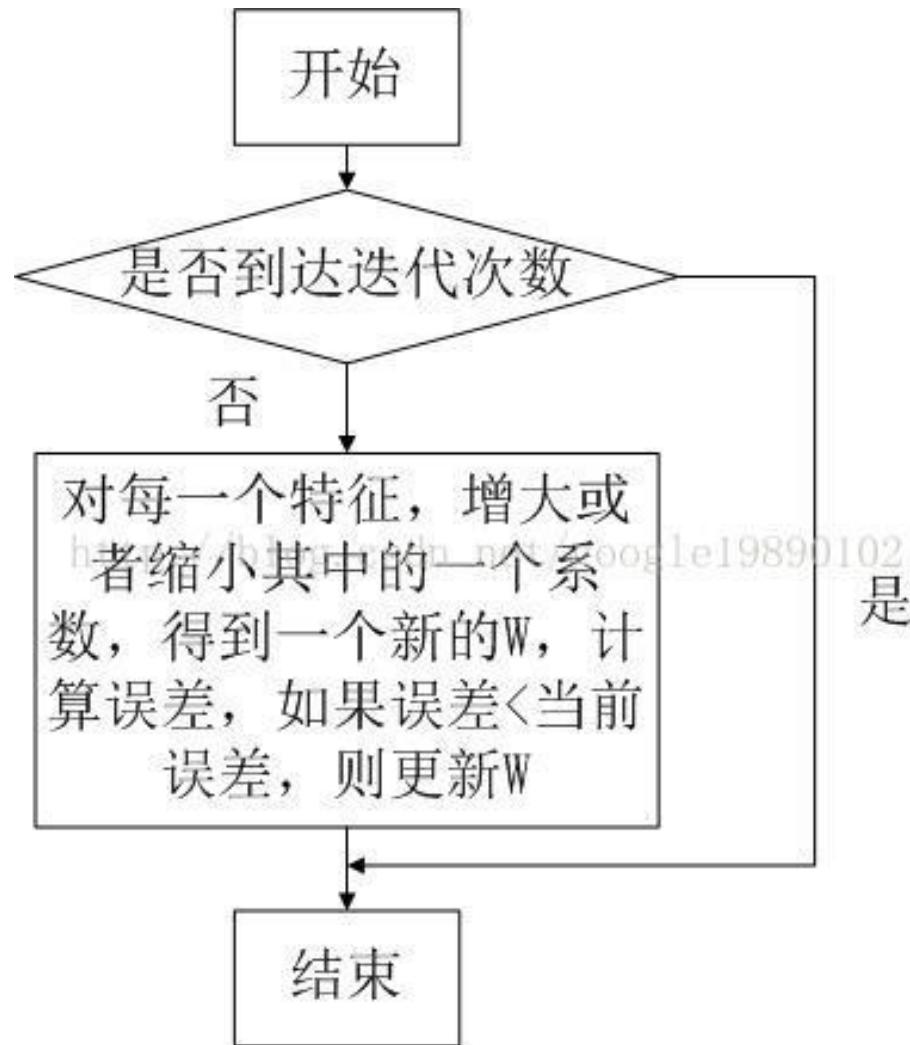
- Lasso: Least Absolute Shrinkage and Selection Operator
- Minimize
$$\min_{\beta} \sum_{i=1}^n \frac{1}{2} (y_i - \sum_{j=1}^p \beta_j x_{ij})^2 + \lambda \sum_{j=1}^p |\beta_j|$$
- Equivalent to minimizing sum of squares with constraint (Lagragian function)

$$\sum_{j=1}^p |\beta_j| \leq s$$

Lasso Explanation

- The bound "s" is a tuning parameter. When "s" is large enough, the constraint has no effect and the solution is just the usual multiple linear least squares regression of y on $x_1, x_2, \dots x_p$.
- However when for smaller values of s ($s >= 0$) the solutions are shrunken versions of the least squares estimates. Often, some of the coefficients b_j are zero. Choosing "s" is like choosing the number of predictors to use in a regression model, and cross-validation is a good tool for estimating the best value for "s".

Algorithms for Lasso



Algorithms for Lasso

- Standard convex optimizer
- Least angle regression (LAR) - Efron et al 2004-computes
- Entire path of solutions. State-of-the-Art until 2008
- Pathwise coordinate descent---New

Ridge Regression

- Minimize

$$\min_{\beta} \frac{1}{2} \sum_{i=1}^n (y_i - \sum_{j=1}^p \beta_j x_{ij})^2 + \lambda \sum_{j=1}^p |\beta_j|^2$$

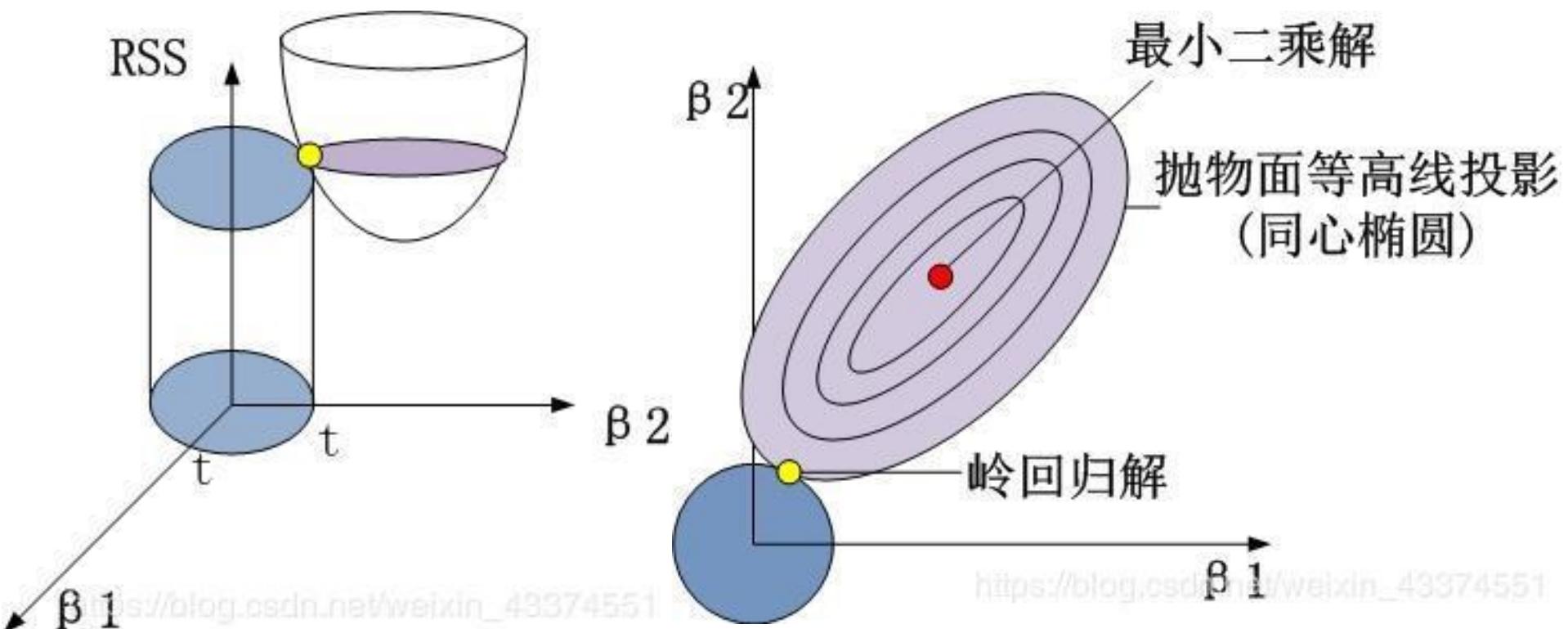
- Equivalent to minimizing sum of squares with constraint

$$\sum_{j=1}^p |\beta_j|^2 \leq s$$

- Close-form solution

$$\beta^{\text{ridge}} = (X^T X + \lambda I)^{-1} X^T Y$$

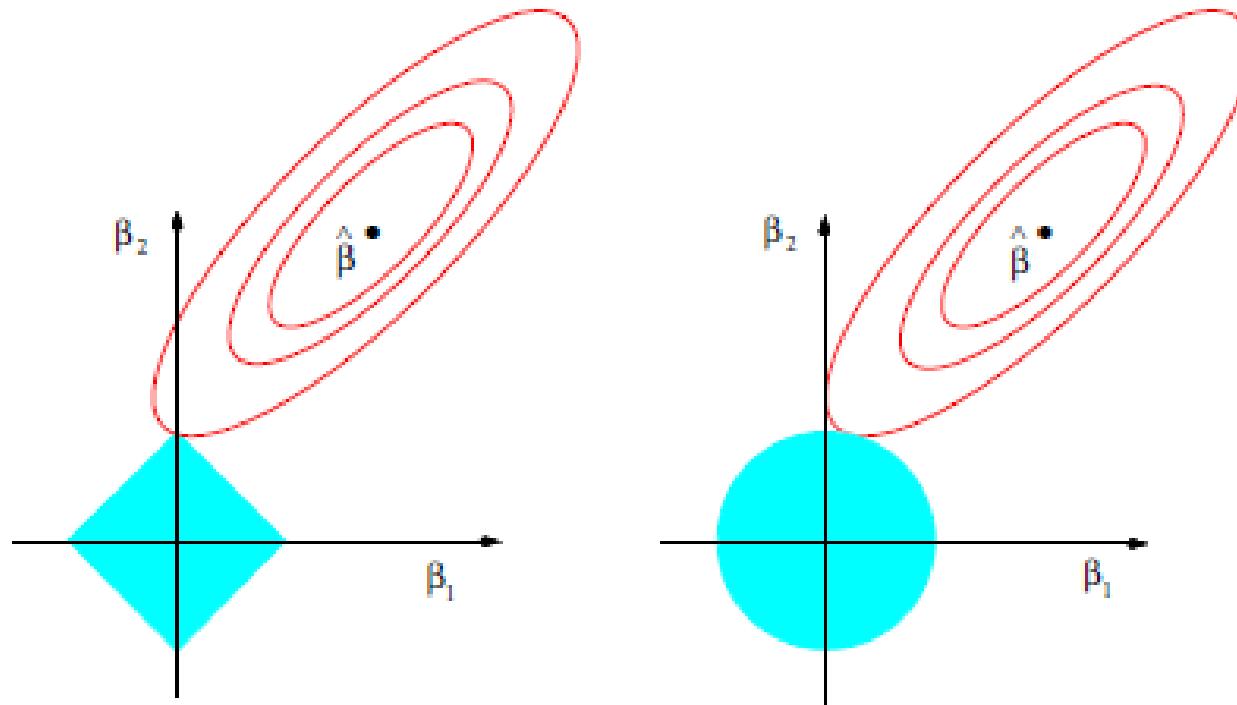
Ridge Regression



Ridge方法

https://blog.csdn.net/weixin_43374551

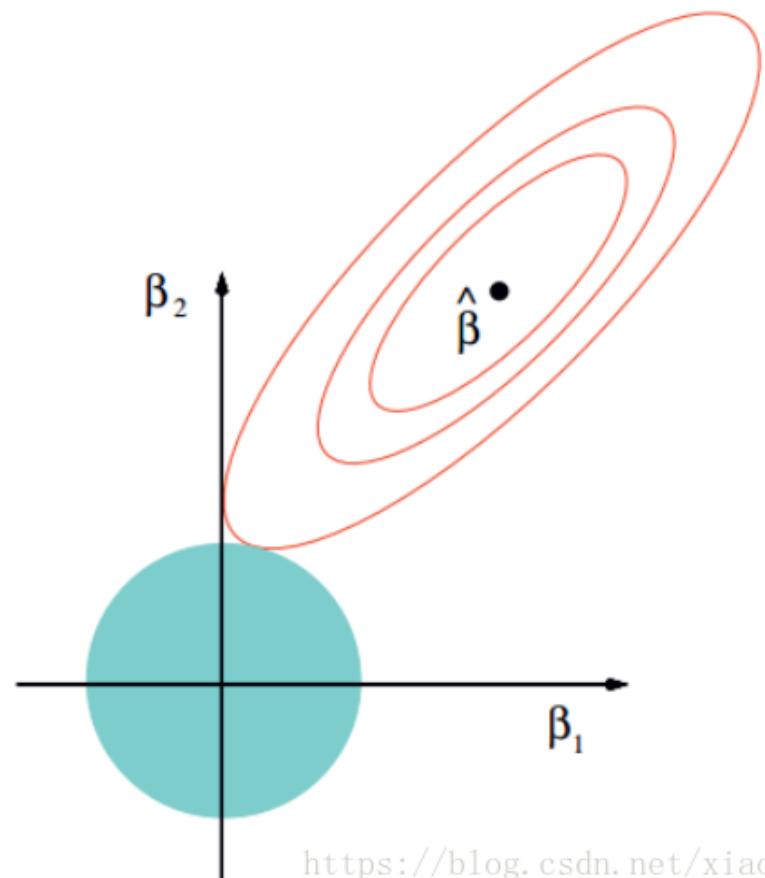
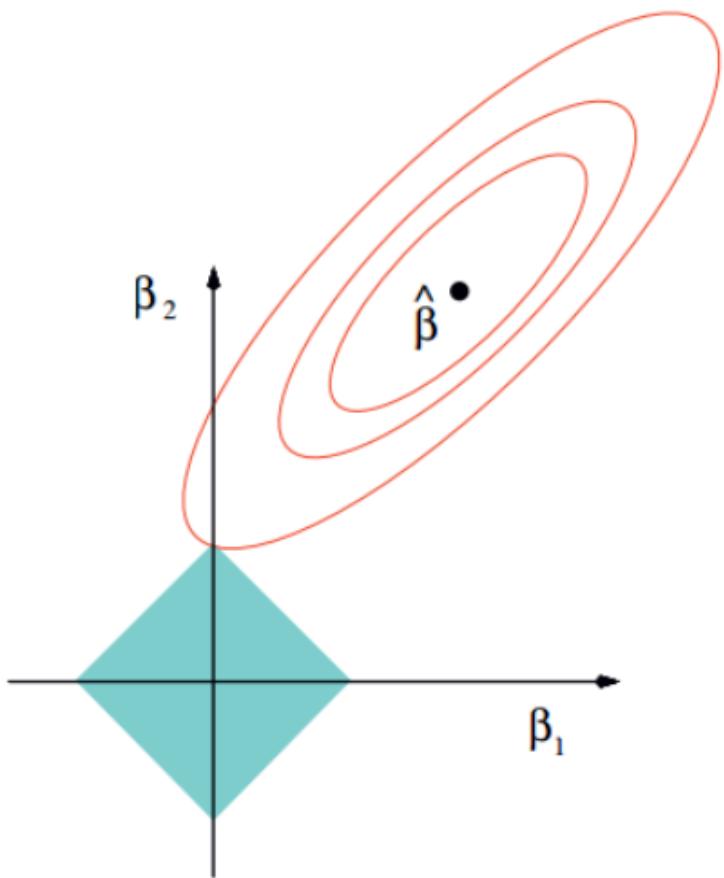
Lasso and Ridge Regression



左图：只要不是特殊情况下与正方形的边相切，一定是与某个顶点优先相交，那必然存在横纵坐标轴中的一个系数为0，起到对变量的筛选的作用。

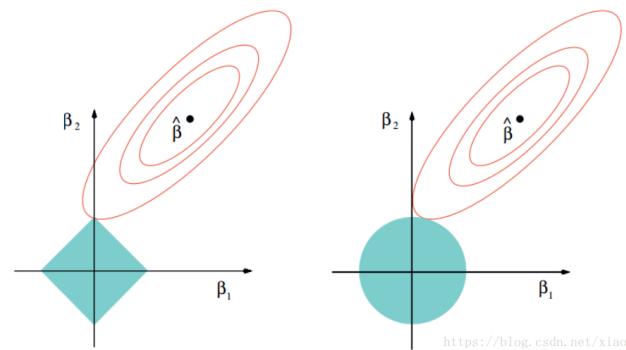
右图：这个圆的限制下，点可以是圆上的任意一点，所以 $q=2$ 的时候也叫做岭回归，岭回归是起不到压缩变量的作用的，在这个图里也是可以看出来的。

LASSO

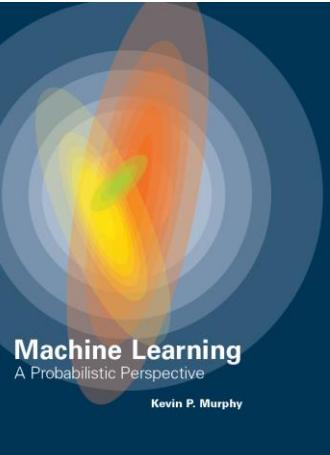


https://blog.csdn.net/xiaozhu_1024

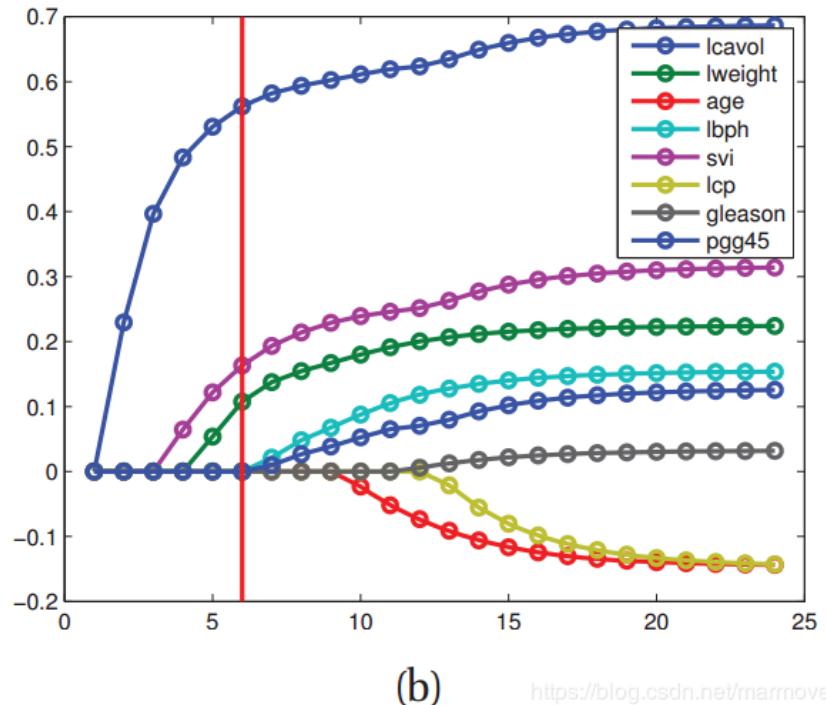
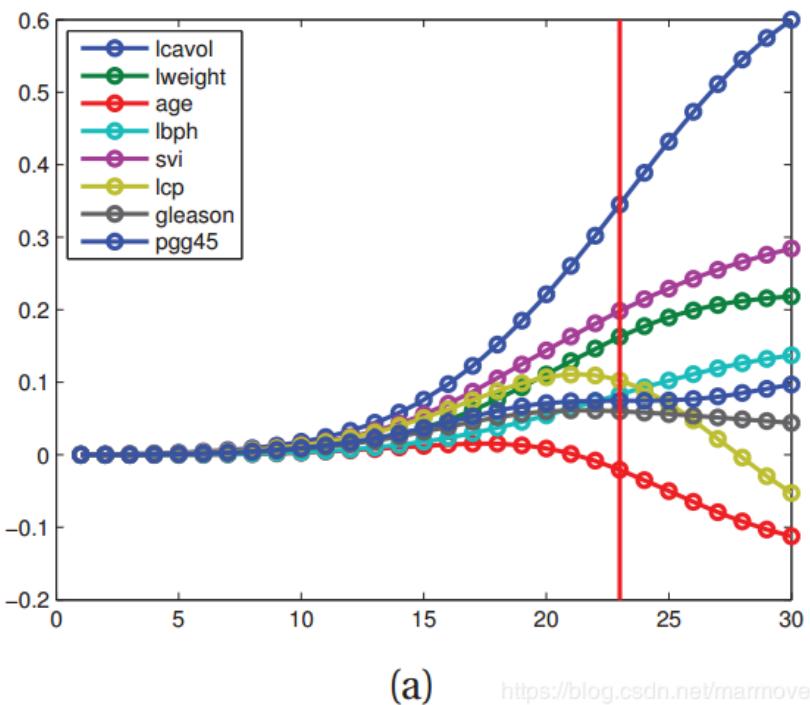
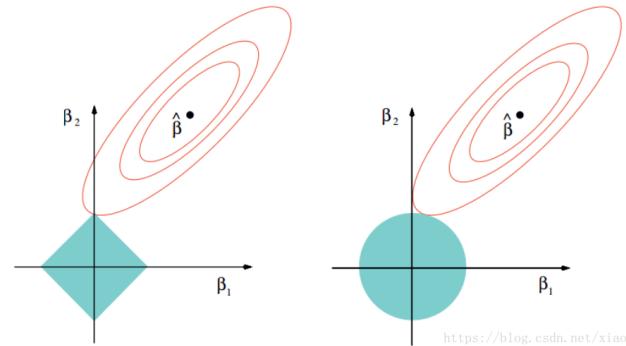
LASSO



- 以二维数据空间为例，说明lasso和Ridge两种方法的差异，左图对应于Lasso方法，右图对应于Ridge方法。
- 两个图是对于两种方法的等高线与约束域。
- 红色的椭圆代表的是随着 λ 的变化所得到的残差平方和， $\hat{\beta}$ 为椭圆的中心点，为对应普通线性模型的最小二乘估计。
- 左右两个图的区别在于约束域，即对应的蓝色区域。
- 等高线和约束域的切点就是目标函数的最优解，Ridge方法对应的约束域是圆，其切点只会存在于圆周上，不会与坐标轴相切，则在任一维度上的取值都不为0，因此没有稀疏；对于Lasso方法，其约束域是正方形，会存在与坐标轴的切点，使得部分维度特征权重为0，因此很容易产生稀疏的结果。
- 所以，Lasso方法可以达到变量选择的效果，将不显著的变量系数压缩至0，而Ridge方法虽然也对原本的系数进行了一定程度的压缩，但是任一系数都不会压缩至0，最终模型保留了所有的变量。



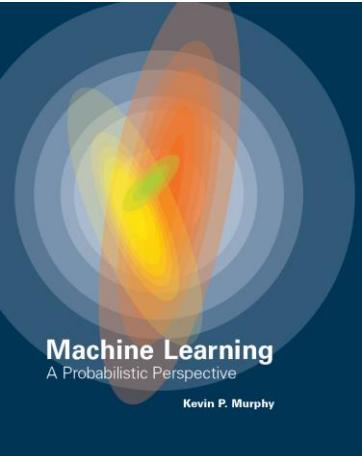
LASSO



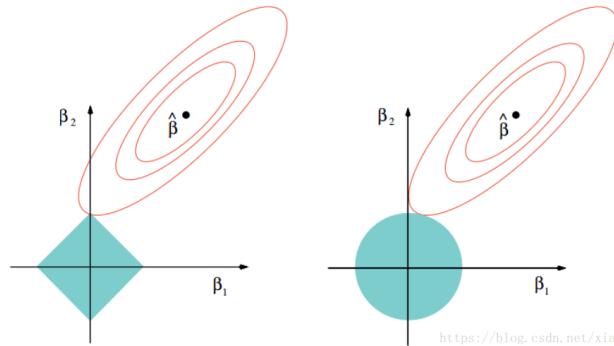
Ridge方法

[https://blog.csdn.net/marmove
/article/details/85260241](https://blog.csdn.net/marmove/article/details/85260241)

LASSO方
法

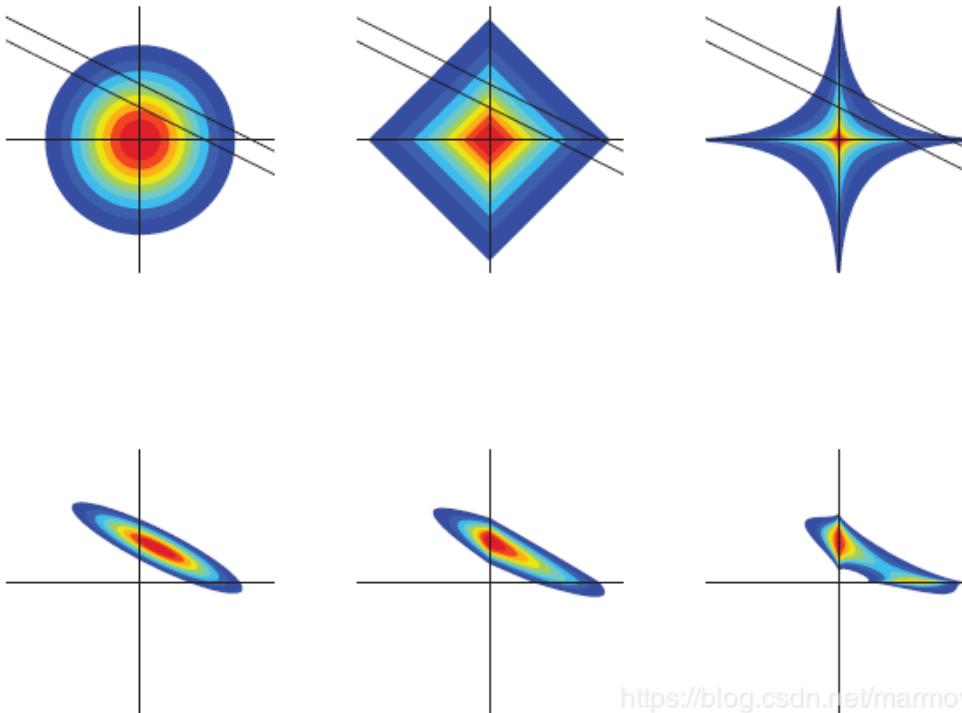


LASSO



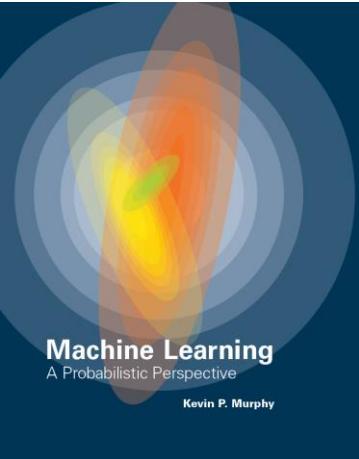
https://blog.csdn.net/xiaozhu_1024

更加灵活的先验：Bridge回归方法

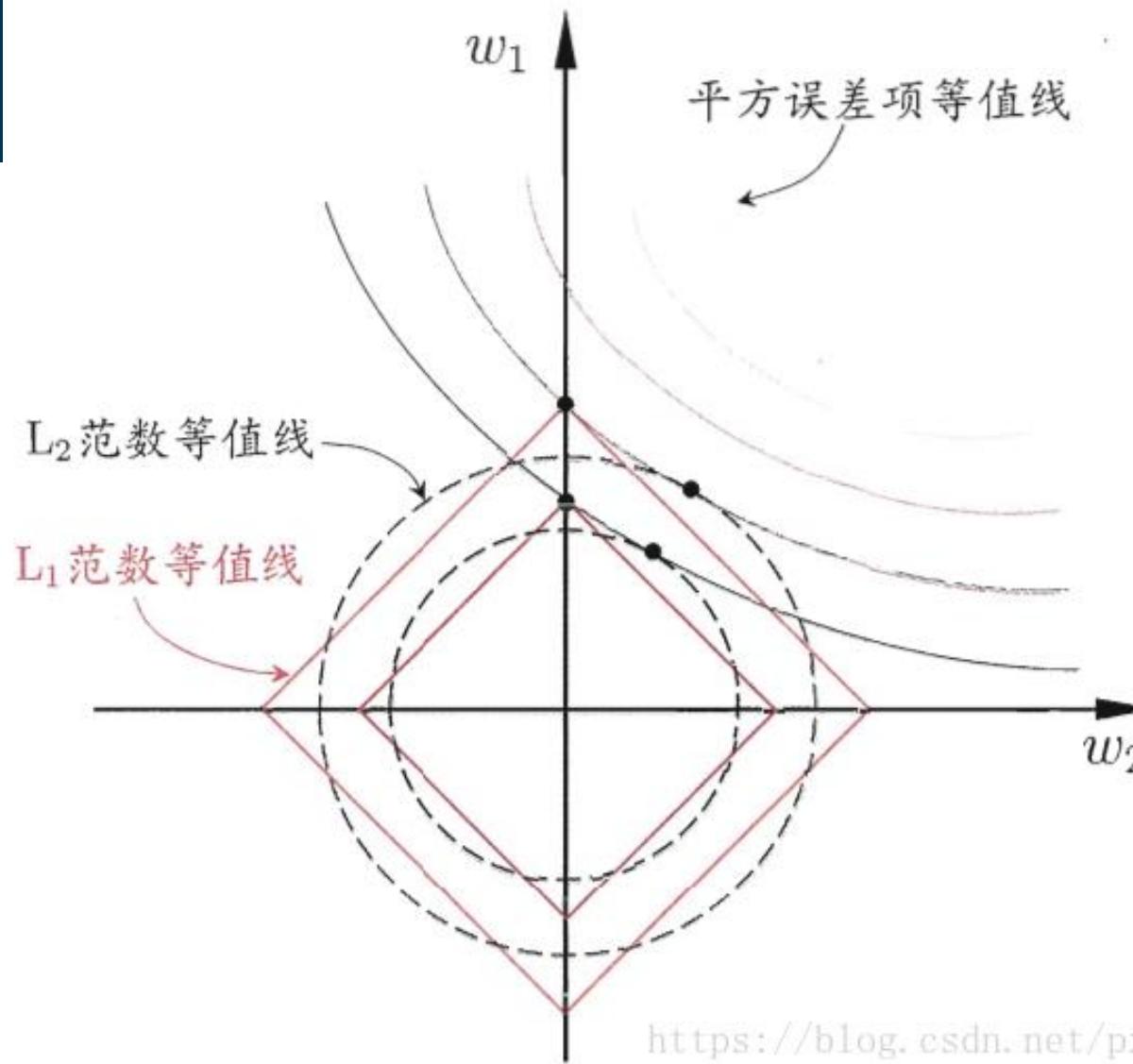


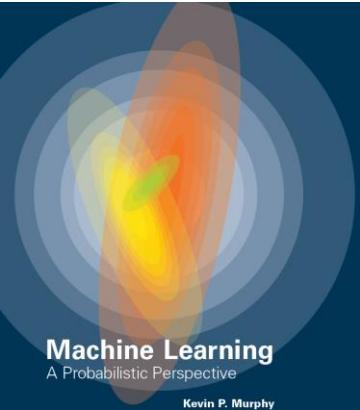
<https://blog.csdn.net/marmove>

[https://blog.csdn.net/marmove
/article/details/85260241](https://blog.csdn.net/marmove/article/details/85260241)



LASSO



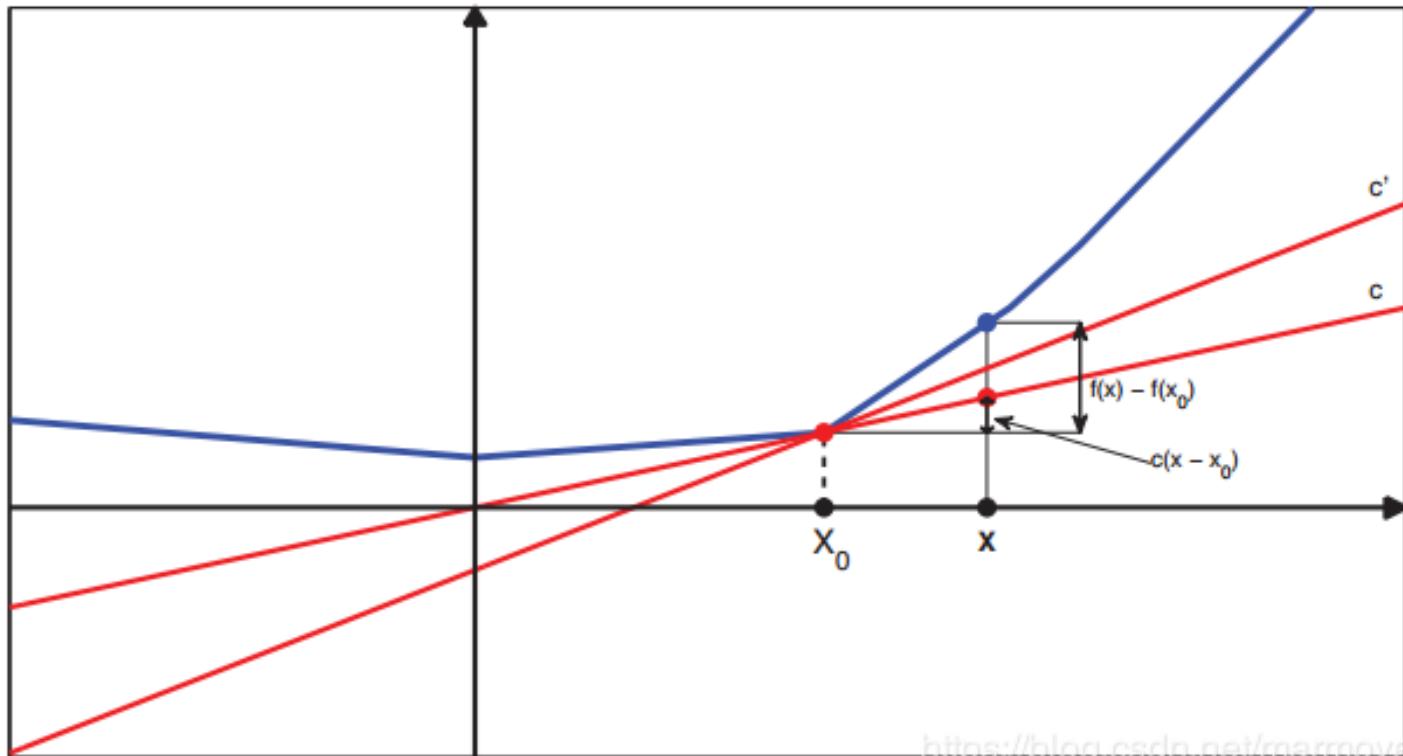


LASSO

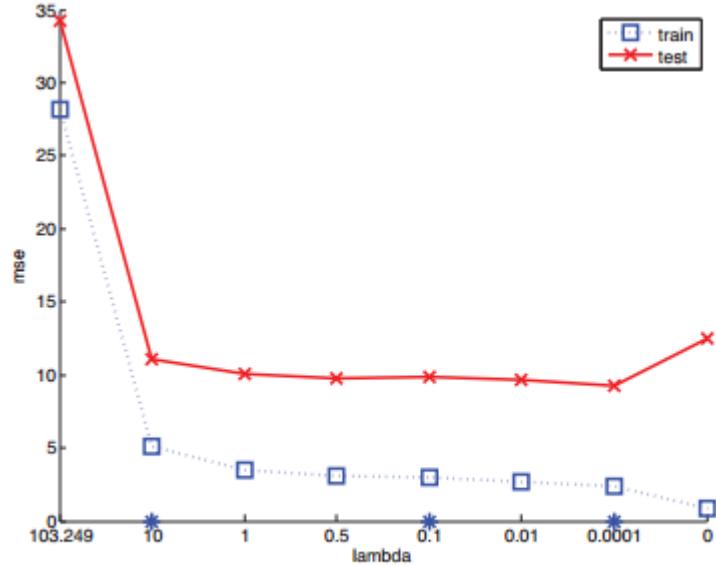
Machine Learning

A Probabilistic Perspective

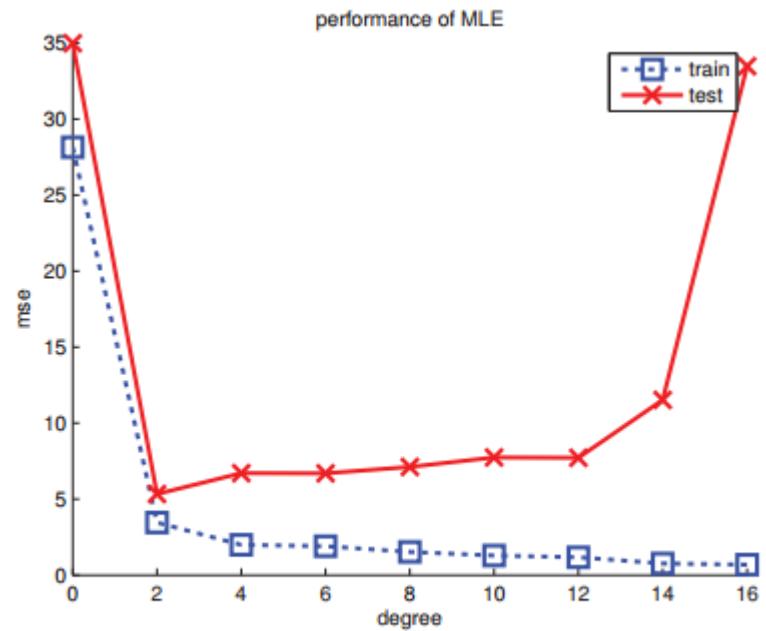
Kevin P. Murphy



LASSO



(a)

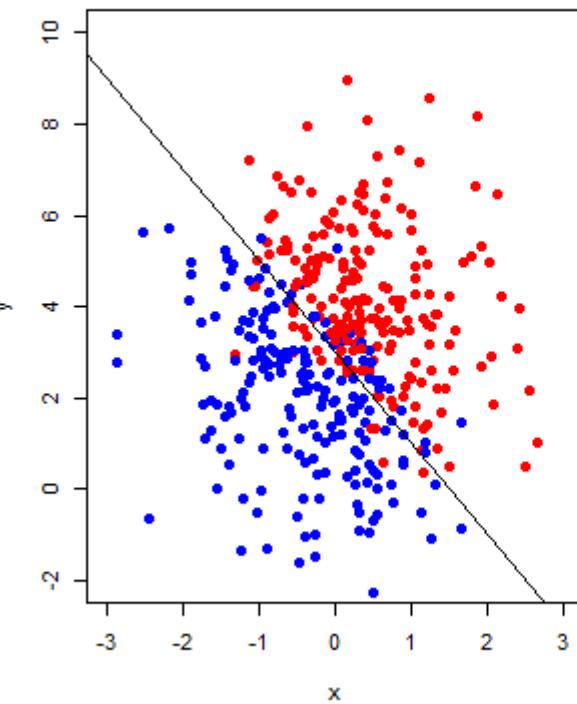


(b) <https://blog.csdn.net/marmove>

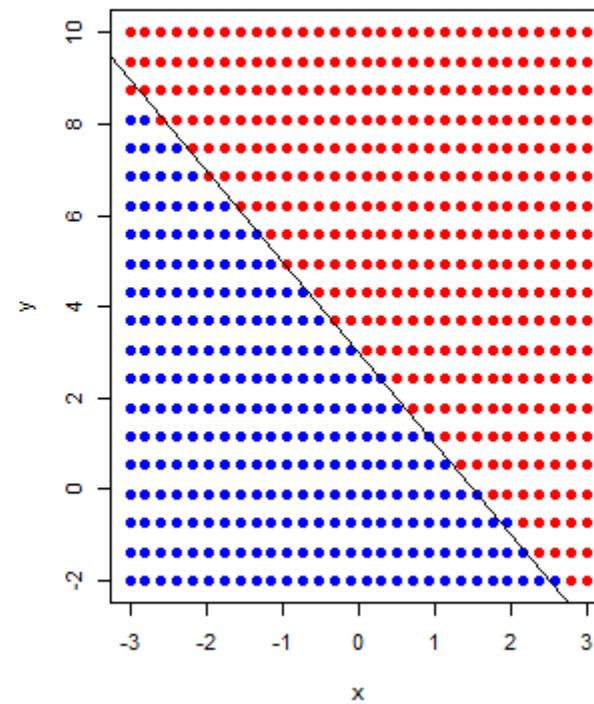
- 左边是lasso， MSE随着变化的情况
- 右边则是subset selection随着K的变化
- 这两个算法的性能是比较接近的

Lasso Model

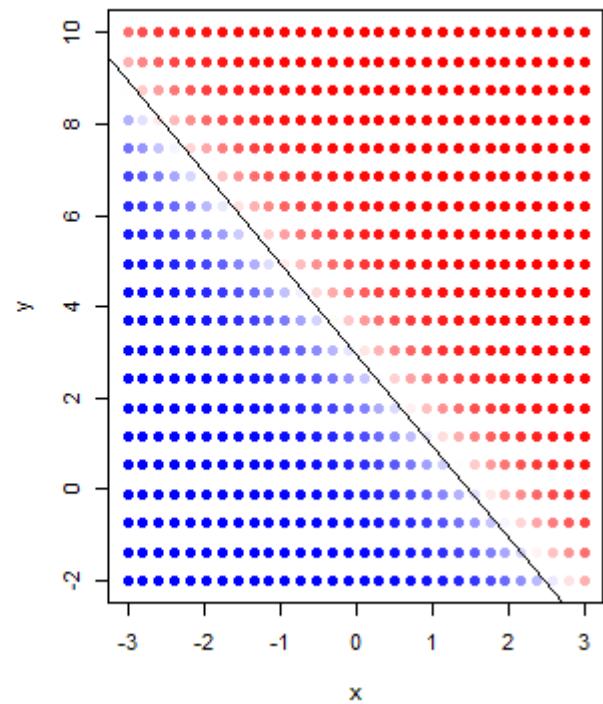
Training Data



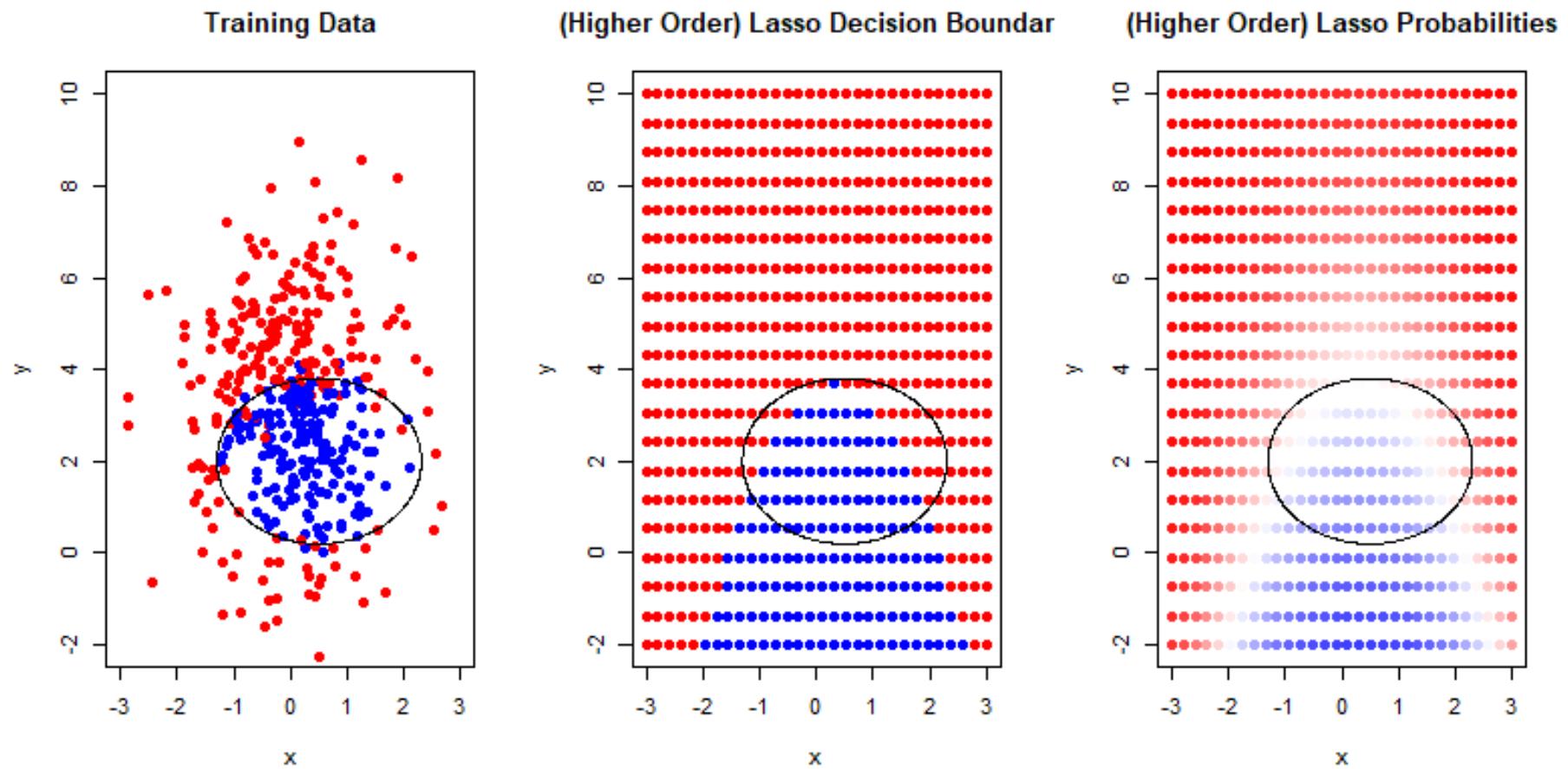
Lasso Decision Boundary



Lasso Probabilities

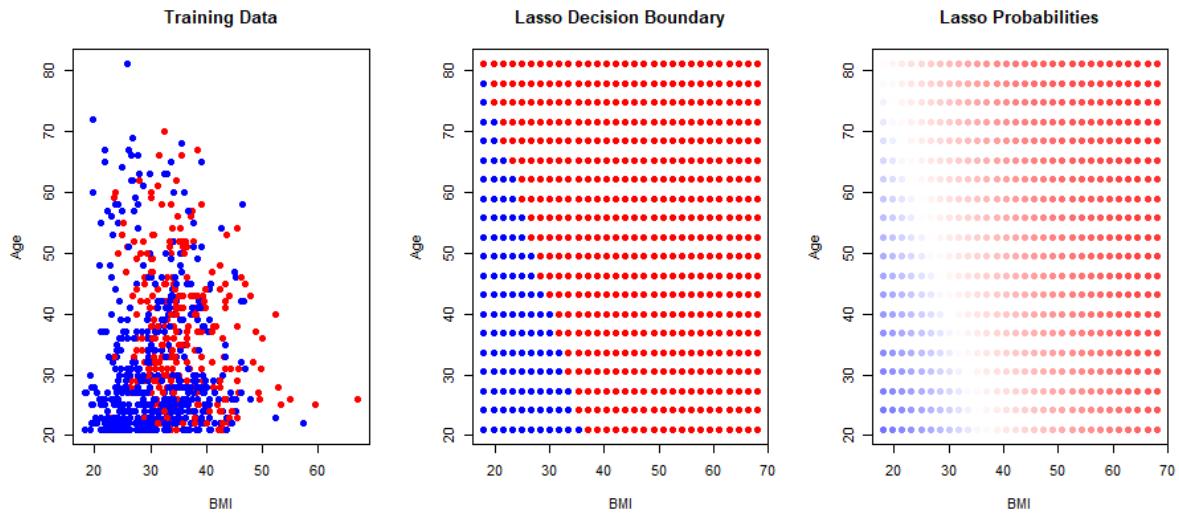


Lasso Model

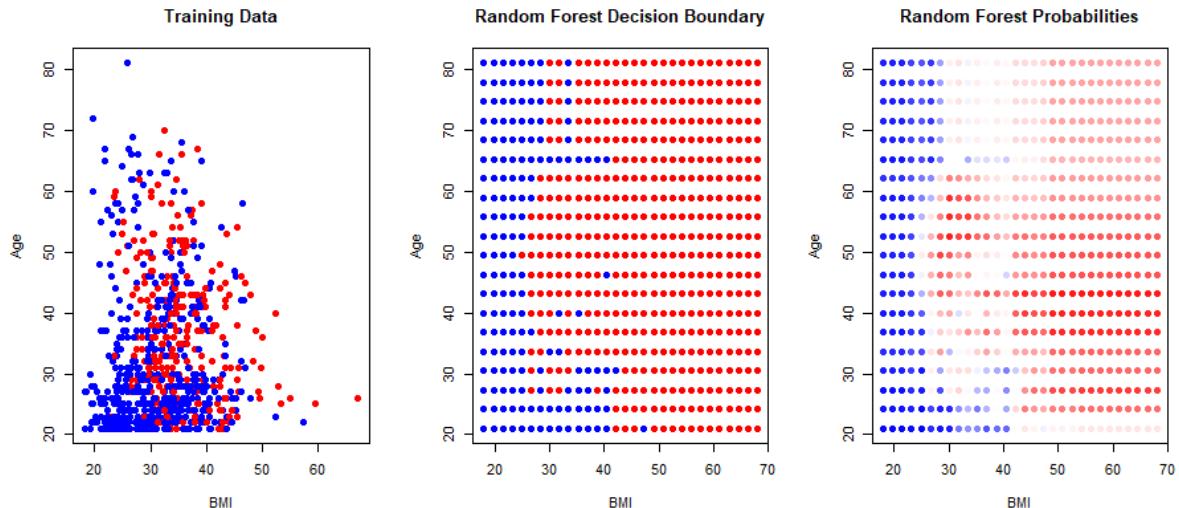


Lasso vs. Random forest

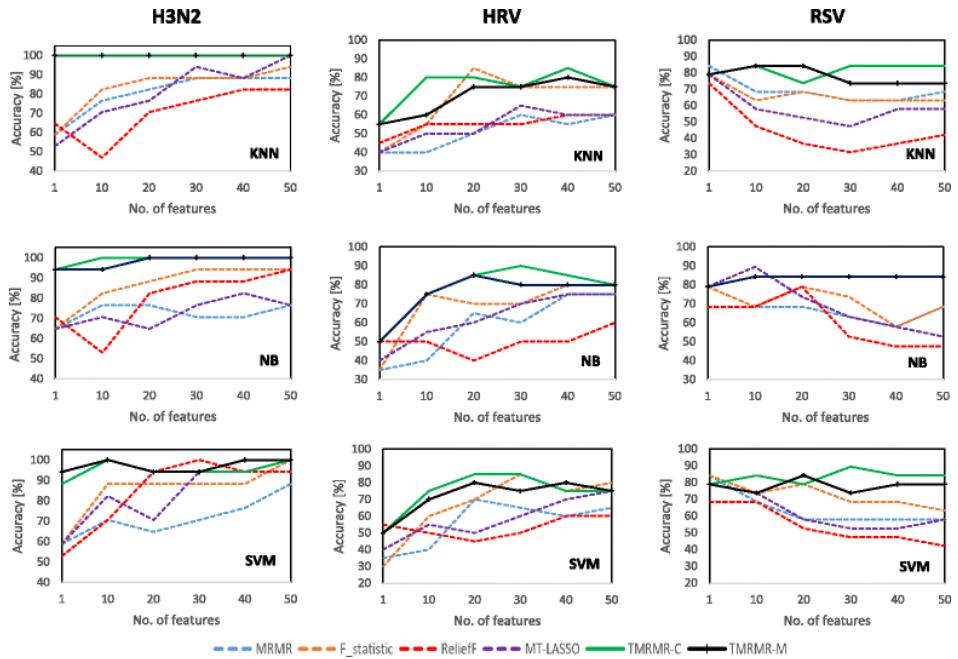
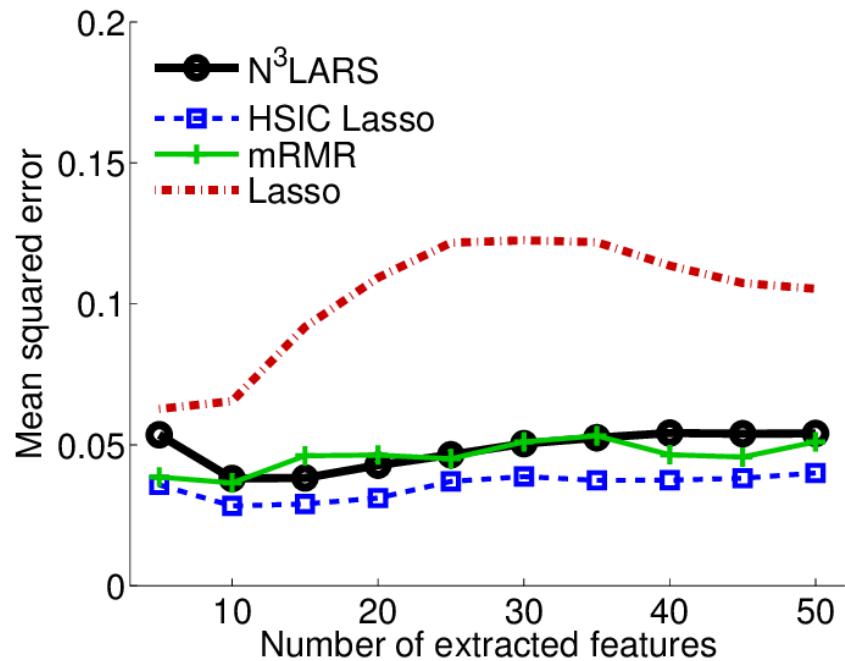
Lasso



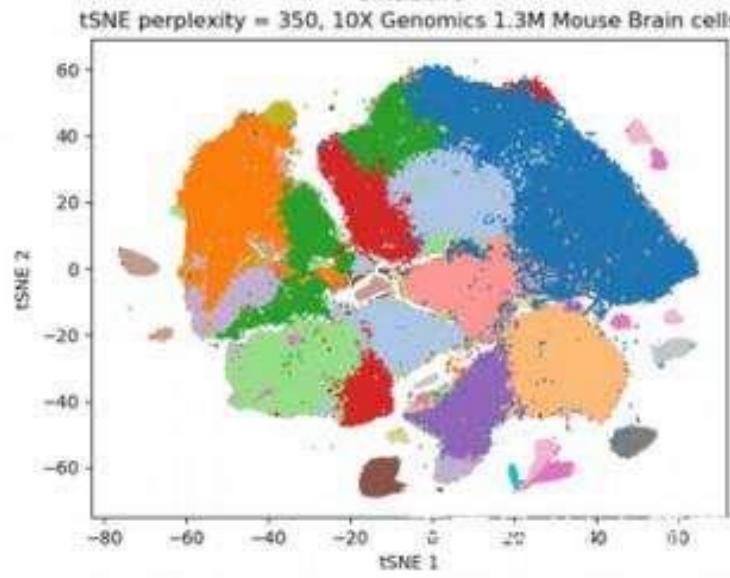
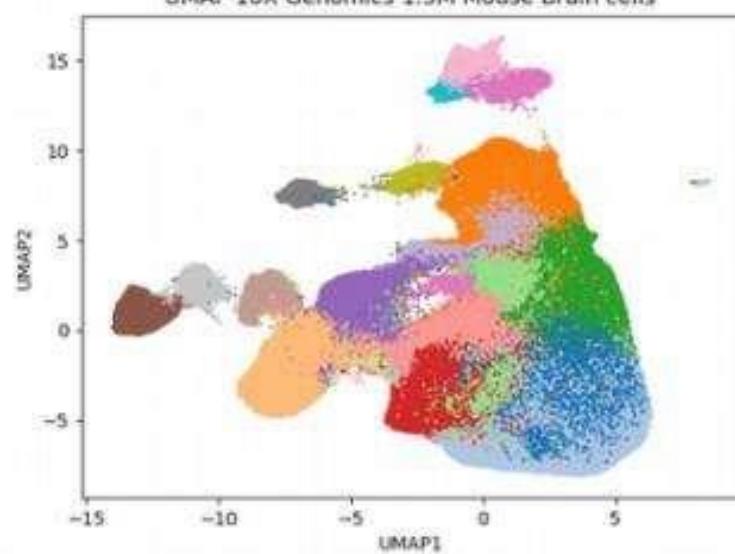
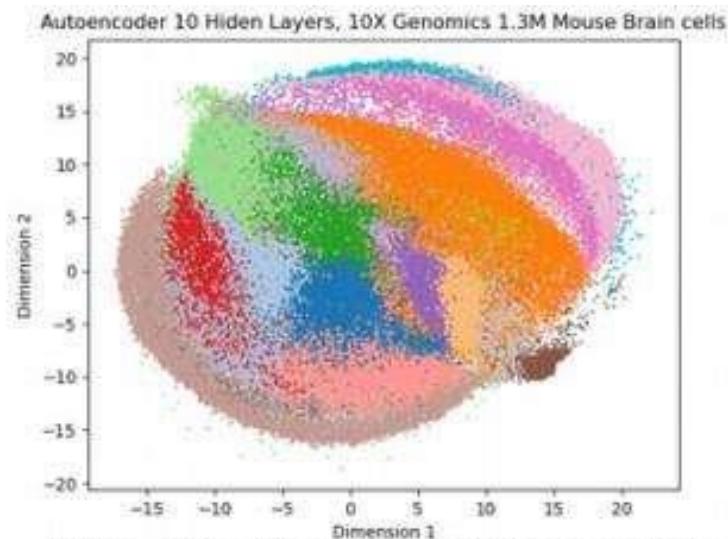
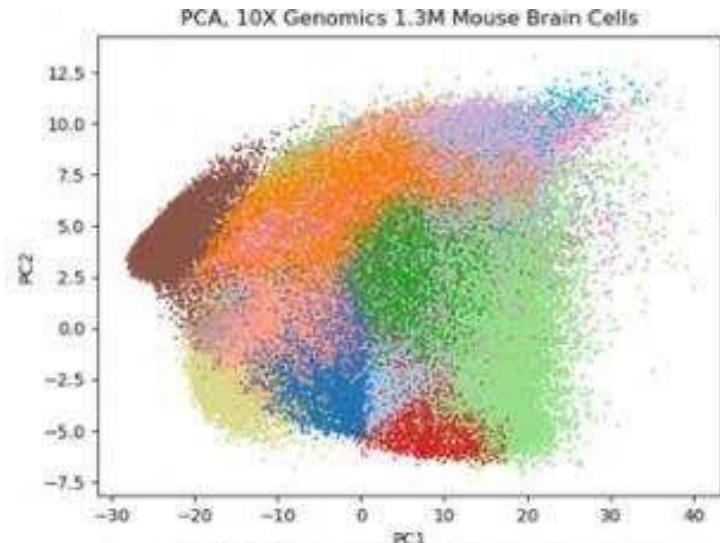
Random forest



Other variable selection methods

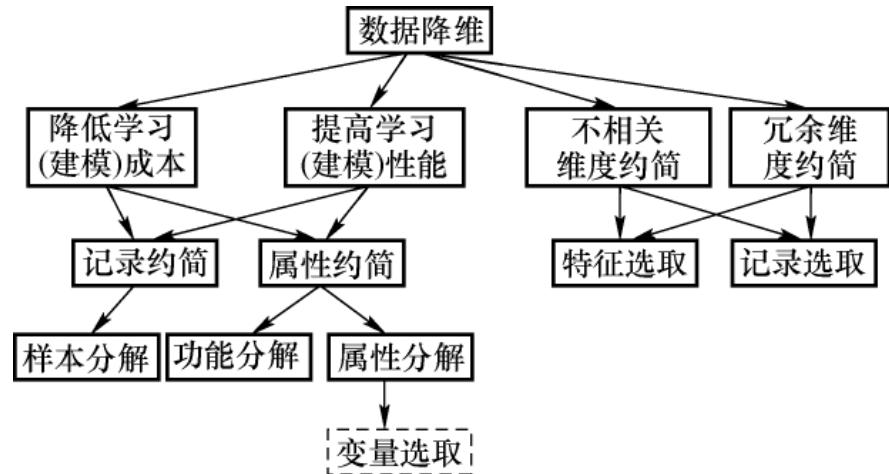


Other dimension reduction methods

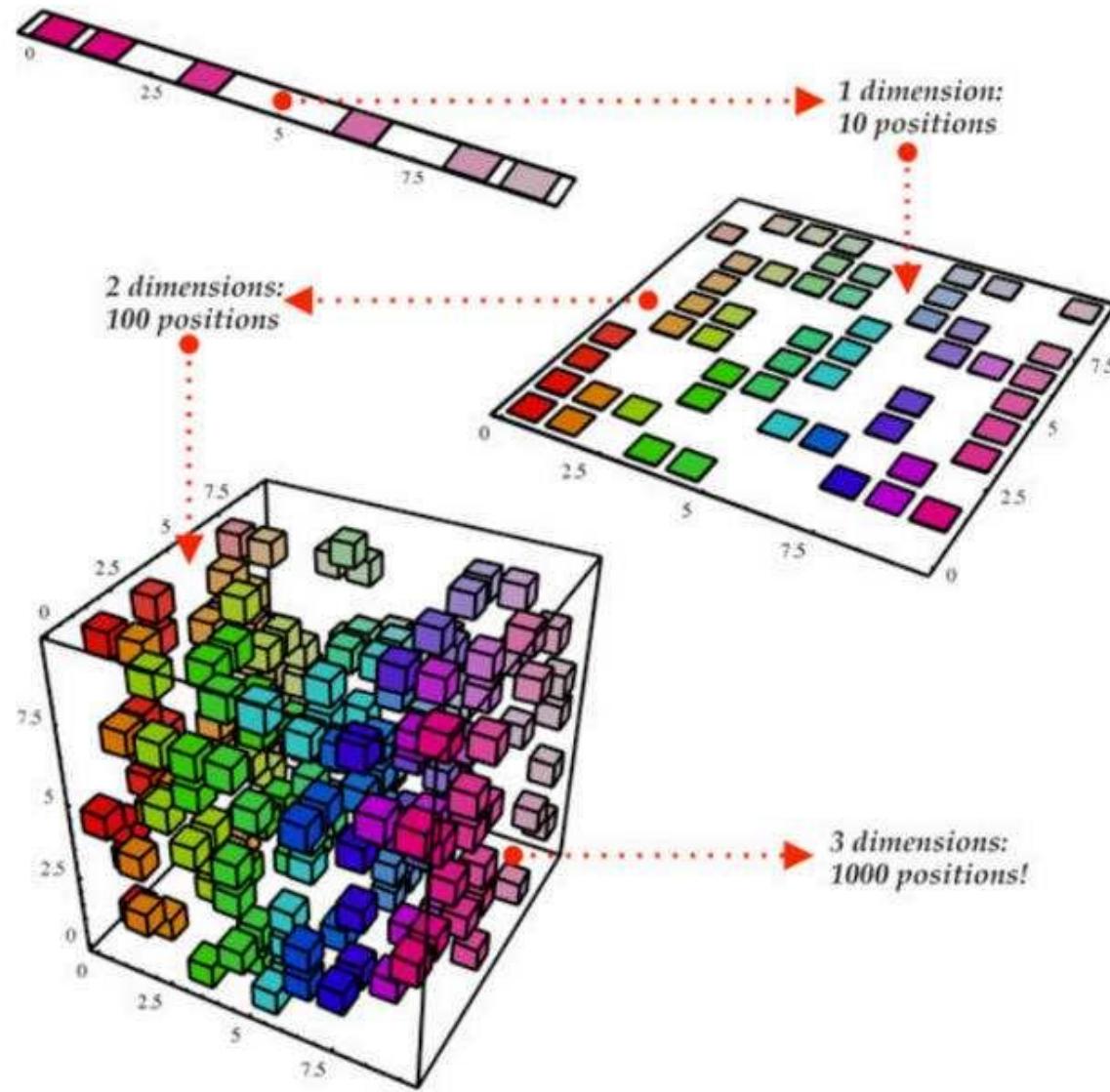


Recap (知识点总结)

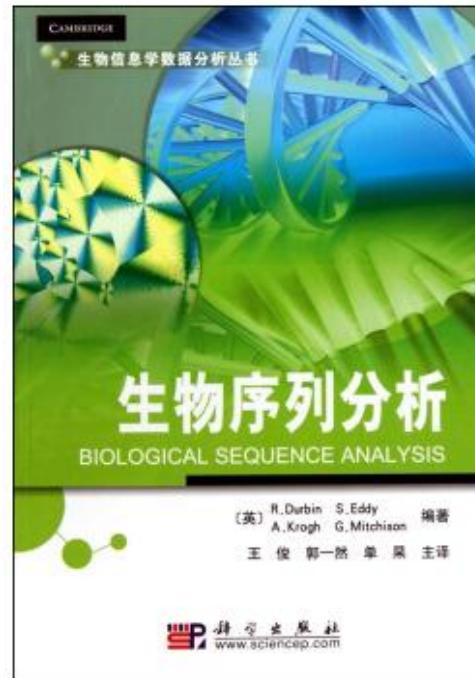
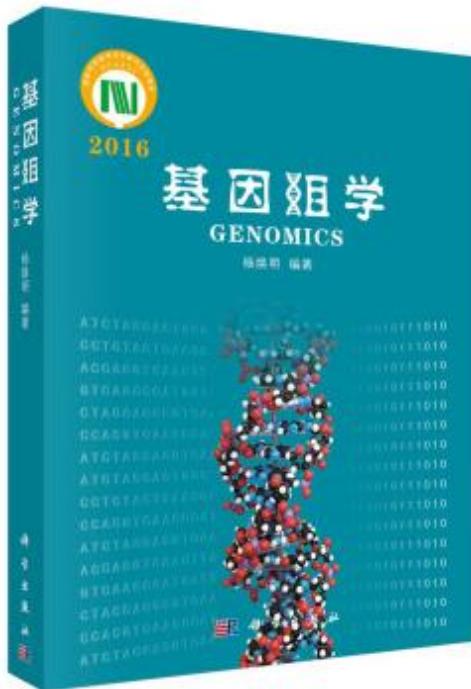
- 降维的必要性：
 - 维度爆炸和分类可视化
 - 特征的冗余
- 降维的方法：
 - 特征选取：mRMR
 - 线性降维：PCA, LDA, CCA
 - 非线性降维：MDS, tSNE
 - 其它特征选取和降维方法：LASSO

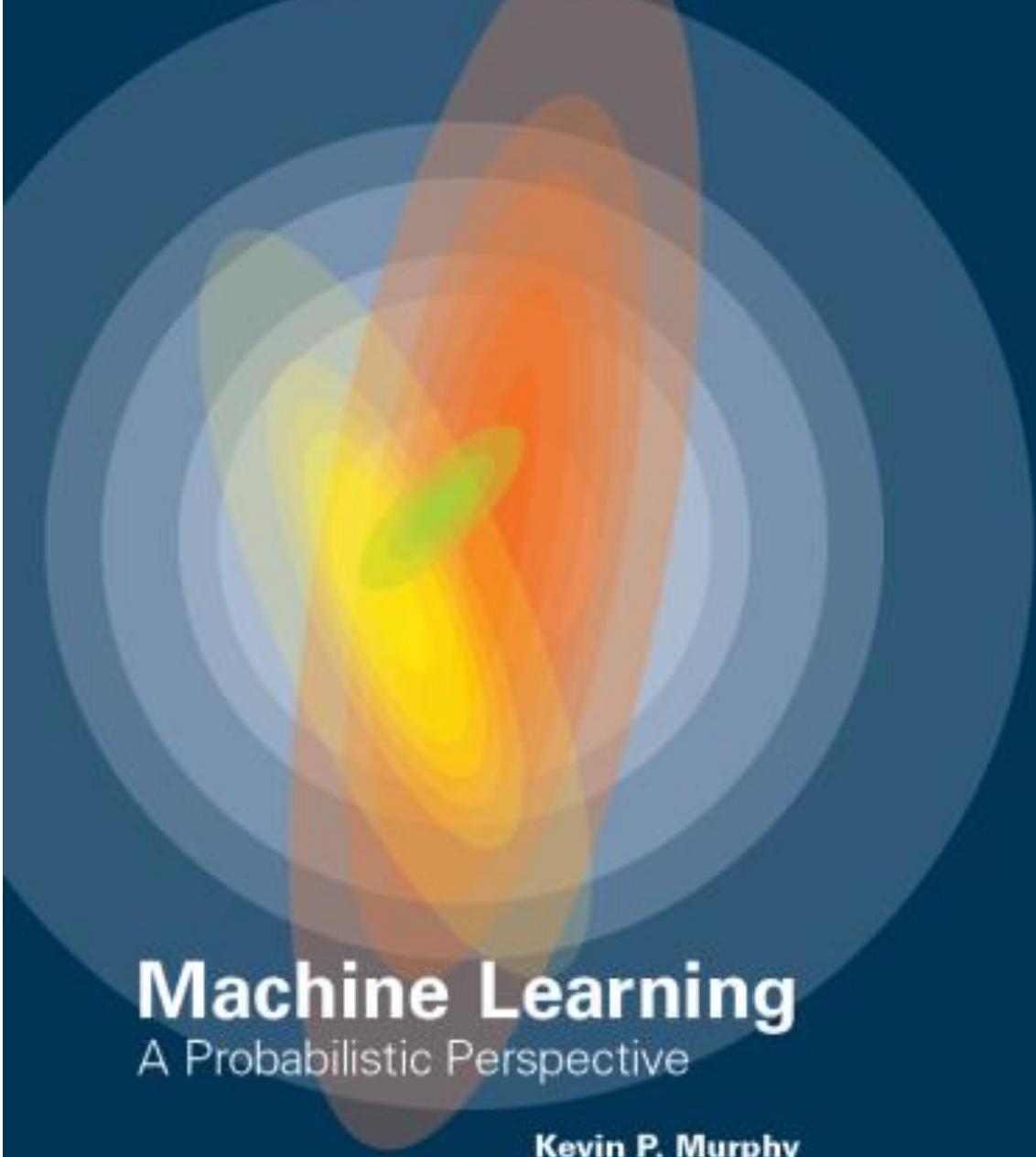


Recap (知识点总结)



References





Machine Learning

A Probabilistic Perspective

Kevin P. Murphy

DEEP LEARNING

Ian Goodfellow, Yoshua Bengio,
and Aaron Courville



